Impact of lockdown on key workers: findings from the COVID-19 survey in four UK national longitudinal studies

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ABSTRACT

Background Key workers played a pivotal role during the national lockdown in the UK’s response to the COVID-19 pandemic. Although protective measures have been taken, the impact of the pandemic on key workers is yet to be fully elucidated.

Methods Participants were from four longitudinal age-homogeneous British cohorts (born in 2001, 1990, 1970 and 1958). A web-based survey provided outcome data during the first UK national lockdown (May 2020) on COVID-19 infection status, changes in financial situation, trust in government, conflict with people around, household composition, psychological distress, alcohol consumption, smoking and sleep duration.

Generalised linear models with logit link assessed the association between being a key worker and the above outcomes. Adjustment was made for cohort design, non-response, sex, ethnicity, adult socioeconomic position (SEP), childhood SEP, the presence of a chronic illness and receipt of a shielding letter. Meta-analyses were performed across the cohorts.

Findings 13,736 participants were included. During lockdown, being a key worker was associated with increased chances of being infected with COVID-19 (OR 1.43, 95% CI 1.22 to 1.68) and experiencing conflict with people around (OR 1.19, 95% CI 1.03 to 1.37). However, key workers were less likely to be worse off financially (OR 0.32, 95% CI 0.24 to 0.65), to consume more alcohol (OR 0.88, 95% CI 0.79 to 0.98) or to smoke more (OR 0.60, 95% CI 0.44 to 0.80) during lockdown. Interestingly, being a key worker was not associated with psychological distress (OR 0.95, 95% CI 0.85 to 1.05).

Interpretation Being a key worker during the first UK COVID-19 lockdown was a double-edged sword, with both benefits and downsides. The UK government had the basic duty to protect its key workers from SARS-CoV-2 infection, but it may have failed to do so, and there is an urgent need to rectify this in light of the ongoing third wave.

INTRODUCTION

The WHO declared the SARS-CoV-2/COVID-19 outbreak a global pandemic on 11 March 2020. In an attempt to contain and limit the spread of the virus, the UK government imposed a national lockdown on 23 March 2020 across England, Scotland and Wales. The restrictions were gradually relaxed starting from June 2020.

Key worker occupations refer to those that are vital for the COVID-19 pandemic response. They account for 33% of the total national workforce according to the UK Office for National Statistics. Although the most common group of key worker occupations are health and social care workers, the majority of key workers actually work in non-health and social care sectors such as in education and childcare, key public services, transport and utilities.

During the UK national lockdown, key workers played a pivotal role in the COVID-19 response. Studies emerged showing that during the pandemic, key workers were suffering from negative health such as more COVID-19 infections and mental distress. Their labour has been praised by the media and there have been multiple editorial letters calling for their better support and protection. Various societal efforts such as the ‘clap for carers’ campaign, food vouchers, free food deliveries, free transportation and parking, and priority access at supermarkets were established during lockdown in an attempt to alleviate some of the burdens faced by key workers. However, the impact of the COVID-19 lockdown on key workers is yet to be fully elucidated.

To answer this question, we used electronic survey data captured at the peak of the UK national lockdown (May 2020) from participants of four UK longitudinal studies spanning multiple generations (18–62 years old). In this exploratory study, we sought to investigate whether being a key worker during lockdown was associated with a set of health and socioeconomic outcomes selected based on previous key worker findings, lockdown trends already observed for the general population (Bann et al, 2020) and topics of public interest as reported in the media.

METHODS

Study design

Participants were from three UK birth cohorts: National Child Development Study (NCDS) born in 1958, 1970 British Cohort Study (BCS70) born in 1970 and Millennium Cohort Study (MCS) born in 2000–2002; and one national longitudinal study with follow-up from adolescence: Next Steps (NS) born in 1989–1990.
During the UK national lockdown, all participants who were alive and not lost to follow-up were sent (between 2 May 2020 and 1 June 2020) an online questionnaire which measured demographic, behavioural and health variables. Survey questions used to derive the exposures, covariates and outcomes of this study are presented in online supplemental table S1. The full survey questions can be accessed here: (https://cls.ucl.ac.uk/wp-content/uploads/2020/12/COVID-19-Online-Survey-Questionnaire-Wave-1-April-2020-Version-2.pdf). Participants provided informed consent.

**Outcomes**
COVID-19 infection was recorded as 0=no and 1=yes, based on a positive antigen or antibody test or strong personal suspicion due to symptoms. Change in financial situation, trust in government and conflict during lockdown (compared with before the COVID-19 outbreak) were standalone self-rated survey question with three answer options: less/worse, same or more/better. We recoded the change in financial situation and trust in government as 0=same or more/better and 1=less/worse. The change in conflict with the people around was recoded as 0=less and 1=more or better. Respondents were also asked if there has been a change in household composition with 0=no and 1=yes available as options. We calculated change scores for psychological distress, alcohol, smoking and sleep. Mental health state was measured both before the outbreak and during the national lockdown using the General Health Questionnaire (GHQ-12) for NS, using a shortened 9-item Malaise inventory for NCDS and BCS70, and using the Kessler K6 score for MCS. Psychological distress was defined as an increase in the corresponding mental state score during compared with the pre-pandemic state. Alcohol consumption was captured as two variables: how often (0–≥4 times per week) and how many (number of drinks per day). A composite variable was computed for the number of drinks per week. An increase in the number of drinks per week during lockdown compared with the pre-pandemic state was scored as 1, and 0 otherwise. For smoking, we calculated the difference between the number of cigarettes smoked before and during lockdown, thus recoding as 0=same or fewer, and 1=more cigarettes smoked. Lastly, participants were asked how many hours of sleep they had per night on average during lockdown and pre-pandemic. Participants who slept less hours were scored as 1, while the remaining were scored as 0.

**Exposures**
Key worker status was self-assigned based on whether the participant believed their work has been classified as critical to the COVID-19 response. Participants actively working during lockdown, but not as key workers, were referred to as other workers. Participants in unpaid employment or volunteers were referred to as volunteers. Participants who were not in paid or unpaid employment but were receiving teaching during lockdown were referred to as in teaching.

**Covariates**
Sex was recorded as 0=male and 1=female, while ethnicity was recoded as 0=White and 1=non-White. NCDS and BCS70 have a very small proportion of non-white participants, thus, ethnicity differences were assessed only for the NS and MCS cohorts. Adult socioeconomic position (SEP) was defined according to the highest educational attainment: degree/higher, advanced-level examination/diploma, ordinary-level examination/General Certificate of Secondary Education or none. For MCS participants who were still in education, their parents’ highest education attainment was used instead. Childhood social class has been recorded according to the UK Office of Population Censuses and Surveys Registrar General’s social class as: professional, managerial and technical, skilled non-manual, skilled manual, partly skilled and unskilled. The presence (yes/no) and nature of any long-standing chronic illnesses were also noted. Individuals who were considered to be at high risk of negative health outcomes in the event of COVID-19 infection should have received a shielding letter, so receipt of a shielding letter was also recorded (yes/no).

**Statistical analysis**
Statistical analysis was performed in R (V3.6.3). Histograms enabled visual assessment of data. Categorical variables were expressed as counts and per cent for each available category. Weights to account for the stratified designs of NS and MCS cohorts have been previously developed. Logistic regression models predicting the response during the COVID-19 data sweep based on demographic, socioeconomic, household and individual predictors of non-response at previous data collection points were used to calculate non-response weights. In the logistic regression models, missing covariate values were generated using multiple imputation. For the COVID-19 survey respondents, the probability of response was predicted, and non-response weights were derived as the inverse of the probability of response and were further calibrated so they sum to the number of respondents in each cohort. The stratified survey design and non-response weights were combined to generate an individualised combined weight for each study respondent (full details available here in the Centre for Longitudinal Study COVID-19 Survey User Guide: https://cls.ucl.ac.uk/wp-content/uploads/2020/12/UCL-Cohorts-COVID-19-Survey-user-guide.pdf).

Generalised linear regression models with logit link were developed to assess whether key worker status was associated with COVID-19 infection; changes in financial situation; trust in government; conflict; household composition; psychological distress; alcohol consumption; smoking and sleep duration during lockdown. For each outcome, the regression models were adjusted for the combined weight, sex, ethnicity, adult SEP, household SEP, the presence of a chronic illness and the receipt of a shielding letter. Each analysis was initially conducted per cohort. Meta-analyses were then performed across the cohorts for each outcome. Cochran’s Q p value and Higgins I² were employed to assess study heterogeneity, while Egger’s test was used to evaluate funnel plot asymmetry to assess whether smaller cohorts led to higher effect sizes (ie, ‘small-study effects’). The results were then corrected for multiple testing at a false discovery rate (FDR) of 0.05. For each outcome significant in meta-analysis, we provide a breakdown per working status category per cohort. Comparisons were made using the X² test.

**Sensitivity analysis**
As mentioned above, we have calculated change scores for psychological distress, alcohol, smoking and sleep. We have pursued further adjustment for baseline states (ie, pre-pandemic mental state score, number of drinks per week, cigarettes smoked per day and sleep hours per night, respectively).

As we have scored the COVID-19 variable as 1, either for strong personal suspicion of infection or a positive antigen or antibody test, self-reporting bias was a concern. Therefore, we ran a sensitivity analysis where only a positive antigen or antibody was scored as 1. As having a COVID-19 test was highly

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selective in May 2020 with health and social care workers being prioritised, we have run an additional sensitivity analysis where we included only participants who had been tested.

RESULTS

Overall, 13 953 out of 38 727 participants (36%) responded to the survey as follows: 5178 out of 8943 (58%) for NCDS, 4223 out of 10 458 for BCS70 (40%), 1907 out of 9380 (20%) for NS and 2645 out of 9946 (27%) for MCS. Responding to the survey we calculated change scores, results when adjustment for baseline state (over and above the covariates in the fully adjusted model) can be found in online supplemental table S2. For outcomes presented in table 1, we included only participants who had been tested.

Significant associations persisted when COVID-19 infection status was scored only based on a positive antigen or antibody test. The meta-analysis revealed a pooled OR 2.74 (95% CI 1.46 to 5.13, p<0.002) (online supplemental table S4). Across all samples, 333 individuals were tested (64 were still waiting their results and 8 had an inconclusive test). In the tested population only, significant associations did not persist. In general, key workers were more likely to have COVID-19 infection compared with other workers and those not working/retired (online supplemental table S5).

Changes in financial situation during lockdown

Across all cohorts, key workers were less likely to be financially worse-off during lockdown with a pooled OR 0.32 (95% CI 0.24 to 0.65, p<0.0001) on the background of insignificant intercohort heterogeneity (I²=0.00%, p=0.631). The association persisted despite correction for multiple testing. Smaller cohorts were not associated with larger effect sizes (Egger’s test p value=0.090, table 3). In general, key workers were less likely to be financially worse-off compared with any other working status category (online supplemental table S5).

Changes in trust in government during lockdown

Key worker status was not associated with changes in trust in government in any of the cohorts (table 2).

Changes in conflict during lockdown

In the BCS70, being a key worker was associated with a change in conflict with the people around (OR 1.38, 95% CI 1.11 to 1.72, p=0.004). The meta-analysis revealed that key workers were more likely to experience same or more conflict during

<table>
<thead>
<tr>
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</thead>
<tbody>
<tr>
<td>Sample size</td>
<td>5178</td>
<td>4223</td>
<td>1907</td>
<td>2645</td>
</tr>
<tr>
<td>Included participants</td>
<td>5119</td>
<td>4132</td>
<td>1876</td>
<td>2609</td>
</tr>
<tr>
<td>Age (years)</td>
<td>62</td>
<td>50</td>
<td>30–31</td>
<td>19–20</td>
</tr>
<tr>
<td>Males (%)</td>
<td>2432 (47.51)</td>
<td>1708 (41.48)</td>
<td>633 (34.76)</td>
<td>770 (29.51)</td>
</tr>
<tr>
<td>Non-white ethnicity (%)</td>
<td>N/A</td>
<td>N/A</td>
<td>361 (19.27)</td>
<td>367 (14.17)</td>
</tr>
<tr>
<td>Adult SEP GCSE—none (%)</td>
<td>1837 (36.61)</td>
<td>1416 (36.43)</td>
<td>434 (24.82)</td>
<td>986 (40.12)</td>
</tr>
<tr>
<td>Childhood SEP I–III (%)</td>
<td>1897 (43.61)</td>
<td>1727 (48.74)</td>
<td>1227 (69.36)</td>
<td>1755 (79.70)</td>
</tr>
<tr>
<td>Chronic Illness (%)</td>
<td>3099 (61.24)</td>
<td>1955 (48.08)</td>
<td>715 (39.20)</td>
<td>830 (33.33)</td>
</tr>
<tr>
<td>Shielding letter (%)</td>
<td>334 (6.57)</td>
<td>196 (4.77)</td>
<td>56 (3.00)</td>
<td>60 (2.30)</td>
</tr>
<tr>
<td>Key workers (%)</td>
<td>938 (18.32)</td>
<td>1396 (33.79)</td>
<td>583 (31.08)</td>
<td>196 (7.51)</td>
</tr>
</tbody>
</table>

Table 1 Characteristics of participants by cohort

The percentages have been derived after subtracting the number of missing data.

GCSE, General Certificate of Secondary Education; N/A, not available; SEP, socioeconomic position.

Cohort study birth year


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Cohort study birth year

lockdown (OR 1.19, 95% CI 1.03 to 1.37, p=0.016) even at an FDR of 0.05. There was low heterogeneity between cohorts ($I^2=29.86\%$, p=0.233) (table 3).

Changes in household composition during lockdown
Key worker status was not associated with changes in household composition in any of the cohorts (table 2).

Changes in psychological distress during lockdown
In the NS cohort, key workers were less likely to have a higher GHQ-12 score (OR 0.77, 95% CI 0.61 to 0.98, p=0.034). However, the meta-analyses revealed that being a key worker was not associated with psychological distress (OR 0.95, 95% CI 0.85 to 1.05, p=0.320, table 2).

Changes in alcohol consumption during lockdown
Only in the BCS70 cohort, key worker status was associated with drinking a lower number of alcoholic drinks per week compared with pre-lockdown (OR 0.77, 95% CI 0.66 to 0.90, p=0.001). The association remained significant even after adjusting for the number of drinks per week prior to lockdown (online supplemental table S3). The meta-analysis confirmed that the association was observed across all cohorts as the pooled OR was 0.88 (95% CI 0.79 to 0.98, p=0.022, table 3) even at an FDR of 0.05. The heterogeneity was high ($I^2=62.46\%$, p=0.046). Key workers were less likely to consume more alcohol compared with volunteers, those in teaching and non-working/retired people (online supplemental table S5).

Changes in smoking during lockdown
In the BCS70 cohort, key workers were less likely to smoke more cigarettes (OR 0.43, 95% CI 0.27 to 0.66, p=0.002, table 2). Even after adjusting for the number of cigarettes smoked per day before lockdown, the association remained (online supplemental table S3).

The association persisted in the meta-analysis as well (pooled OR 0.60, 95% CI 0.44 to 0.80, p=0.005, table 3), on a background of a relatively low heterogeneity ($I^2=24.09\%$, p=0.257) and despite correction for multiple testing. Key workers were less likely to smoke more cigarettes compared with those not working/retired.

Changes in sleep during lockdown
Only participants from the MCS cohort reported that they slept less since the UK national lockdown (OR 1.64, 95% CI 1.11 to 2.38, p=0.011). The association persisted even after adjusting for the number of hours slept before lockdown (online supplemental table S3).

However, the meta-analysis failed to confirm such association (pooled OR 1.06, 95% CI 0.94 to 1.19, p=0.350, table 3). Heterogeneity was high ($I^2=61.75\%$, p=0.049).

DISCUSSION
Findings statement
Our data from the four UK national longitudinal studies during the COVID-19 lockdown (May 2020) show that key worker status was associated with both negative (more COVID-19...
<table>
<thead>
<tr>
<th>Cohort</th>
<th>Effect size</th>
<th>COVID-19 infection*</th>
<th>Change in financial situation†</th>
<th>Change in trust in government</th>
<th>Change in conflict§</th>
<th>Change in household composition¶</th>
<th>Change in psychological distress**</th>
<th>Change in alcohol consumption††</th>
<th>Change in smoking‡‡</th>
<th>Change in sleep duration§§</th>
</tr>
</thead>
<tbody>
<tr>
<td>1958</td>
<td>OR (95% CI)</td>
<td>1.95 (1.45 to 2.61)</td>
<td>0.34 (0.27 to 0.42)</td>
<td>0.94 (0.76 to 1.14)</td>
<td>1.16 (0.91 to 1.51)</td>
<td>0.92 (0.71 to 1.18)</td>
<td>1.00 (0.82 to 1.21)</td>
<td>0.91 (0.73 to 1.11)</td>
<td>0.78 (0.40 to 1.40)</td>
<td>1.16 (0.93 to 1.45)</td>
</tr>
<tr>
<td></td>
<td>P value</td>
<td>&lt;0.0001</td>
<td>-</td>
<td>0.516</td>
<td>0.244</td>
<td>0.522</td>
<td>0.990</td>
<td>0.344</td>
<td>0.428</td>
<td>0.184</td>
</tr>
<tr>
<td>1970</td>
<td>OR (95% CI)</td>
<td>1.18 (0.92 to 1.51)</td>
<td>0.31 (0.26 to 0.37)</td>
<td>1.12 (0.94 to 1.34)</td>
<td>1.38 (1.11 to 1.72)</td>
<td>1.17 (0.94 to 1.44)</td>
<td>1.04 (0.87 to 1.24)</td>
<td>0.77 (0.66 to 0.90)</td>
<td>0.43 (0.27 to 0.66)</td>
<td>0.95 (0.79 to 1.14)</td>
</tr>
<tr>
<td></td>
<td>P value</td>
<td>0.184</td>
<td>-</td>
<td>0.191</td>
<td>0.04</td>
<td>0.160</td>
<td>0.669</td>
<td>0.001</td>
<td>0.0002</td>
<td>0.577</td>
</tr>
<tr>
<td>1989–1990</td>
<td>OR (95% CI)</td>
<td>1.36 (0.98 to 1.90)</td>
<td>0.29 (0.22 to 0.38)</td>
<td>0.95 (0.75 to 1.21)</td>
<td>1.00 (0.69 to 1.40)</td>
<td>0.87 (0.63 to 1.19)</td>
<td>0.77 (0.61 to 0.98)</td>
<td>1.01 (0.80 to 1.27)</td>
<td>0.70 (0.39 to 1.21)</td>
<td>0.94 (0.72 to 1.22)</td>
</tr>
<tr>
<td></td>
<td>P value</td>
<td>0.067</td>
<td>&lt;0.0001</td>
<td>0.191</td>
<td>0.04</td>
<td>0.160</td>
<td>0.669</td>
<td>0.001</td>
<td>0.0002</td>
<td>0.651</td>
</tr>
<tr>
<td>2000–2002</td>
<td>OR (95% CI)</td>
<td>1.47 (0.78 to 2.76)</td>
<td>0.39 (0.24 to 0.65)</td>
<td>0.75 (0.49 to 1.11)</td>
<td>0.89 (0.56 to 1.37)</td>
<td>0.72 (0.46 to 1.07)</td>
<td>0.84 (0.59 to 1.20)</td>
<td>1.35 (0.87 to 2.04)</td>
<td>0.93 (0.32 to 2.20)</td>
<td>1.64 (1.11 to 2.38)</td>
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<tr>
<td></td>
<td>P value</td>
<td>0.234</td>
<td>0.0002</td>
<td>0.158</td>
<td>0.614</td>
<td>0.118</td>
<td>0.340</td>
<td>0.172</td>
<td>0.895</td>
<td>0.011</td>
</tr>
</tbody>
</table>

All analyses used generalised linear models with logit link. Significant p values are highlighted in bold.

Adjustment was made for the individualised combined weight (accounting for non-response and stratified cohort design where appropriate), sex, ethnicity, highest educational attainment (adult SEP), childhood SEP, presence of chronic illness and shielding letter.

*COVID-19 infection was coded as 0=no, 1=Yes.
†Change in financial situation was coded as 0=same or better, 1=worst.
‡Change in trust in government was coded as 0=same or more, 1=less.
§Change in conflict during lockdown was coded as 0=less, 1=more.
¶Change in household composition during lockdown was coded as 0=no, 1=Yes.
**Change in psychological distress during lockdown was coded as 1=an increase in the mental health score, 0=same or lower score.
††Change in alcohol consumption was coded as 1=more drinks per week during lockdown compared with pre-lockdown or 0=same or lower.
‡‡Change in smoking was coded as 0=less cigarettes smoked during lockdown compared with before lockdown, 1=more cigarettes.
§§Change in sleep duration was coded as 0=less or more hours slept during lockdown and 1=less hours slept during lockdown.
SEP, socioeconomic position.
infection and conflict with people around) and positive outcomes (financial stability, less drinks per week and less cigarettes per day). There was no association between being a key worker and changes in household composition, trust in government, psychological distress or sleeping less during lockdown.

**Interpretation**

Being a key worker during the UK national lockdown was a challenging experience. Key workers were three times less likely to report psychological distress or sleeping less during lockdown. Key workers were three times less likely to report psychological distress or sleeping less during lockdown. Better access to the vaccine programme for non-healthcare key workers as well. Key workers tended to experience more conflict with the people around them, some of which might have been augmented by relatives’ fear of getting infected, or key workers’ worries over bringing the virus home to their loved ones.

**Implications for key stakeholders**

Pandemics can last multiple years and new viral mutations can arise.16 With the worrying emergence of multiple COVID-19 strains, some of which could potentially be vaccine resistant,17 concerns about the prospect of a fourth wave and a fourth lockdown are more than justified. UK policymakers know about the higher risk of COVID-19 infections in key workers, and they have a duty to implement better protective measures to those used in the first wave, which may not have been enough. Potential solutions include the free and more abundant supply of personal protective equipment, the free provision of COVID-19 rapid lateral flow tests and the enforcement of more secure work-related controls to ensure that the employers are complying with the COVID-19 legislation and are protecting the well-being of their key workers. Although we did not directly explore the impact of lockdown on the household members of key workers, the excess conflict is a worrying finding and suggests that the government should do more to support the families of key workers as a whole, including through support groups and tangible benefits where appropriate.

**Strengths and weaknesses of the study**

The main strengths of the study are the large number of participants included, age-matching within the cohorts by design and cohort data spanning multiple generations from 19 to 74 years old. The longitudinal nature of the cohorts enabled the derivation of non-response weights which have been included in all analyses to address the data missingness issue.

Limitations include low response rates (especially in younger cohorts such as NS and MCS). Data binarisation enabled us to generate singular ORs across the cohorts to be used in meta-analysis, but continuous non-linear effects would be missed. Except for the psychological distress, which was completely based on validated scores, the majority of survey questions were self-reported, which might have had a potential for social desirability bias. In addition, questions on alcohol consumption and smoking were self-reported, and we could not compare with the pre-lockdown values.

**Table 3**

<table>
<thead>
<tr>
<th>Outcome</th>
<th>n</th>
<th>I²</th>
<th>Q</th>
<th>P value</th>
<th>OR (95% CI)</th>
<th>P value</th>
<th>Egger’s test p value</th>
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<tbody>
<tr>
<td>COVID-19 infection*</td>
<td>11,076</td>
<td>54.72%</td>
<td>6.63</td>
<td>0.085</td>
<td>1.43 (1.22 to 1.68)</td>
<td>&lt;0.0001</td>
<td>0.801</td>
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<td>Change in financial situation†</td>
<td>10,649</td>
<td>0.00%</td>
<td>1.73</td>
<td>0.631</td>
<td>0.32 (0.24 to 0.65)</td>
<td>&lt;0.0001</td>
<td>0.534</td>
</tr>
<tr>
<td>Change in trust in government‡</td>
<td>10,236</td>
<td>29.57%</td>
<td>4.26</td>
<td>0.235</td>
<td>1.00 (0.89 to 1.11)</td>
<td>0.932</td>
<td>0.126</td>
</tr>
<tr>
<td>Change in conflict§</td>
<td>10,234</td>
<td>29.86%</td>
<td>4.28</td>
<td>0.233</td>
<td>1.19 (1.03 to 1.37)</td>
<td>0.016</td>
<td>0.137</td>
</tr>
<tr>
<td>Change in household composition¶</td>
<td>10,736</td>
<td>44.96%</td>
<td>5.45</td>
<td>0.142</td>
<td>0.98 (0.85 to 1.12)</td>
<td>0.743</td>
<td>0.073</td>
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<tr>
<td>Change in psychological distress**</td>
<td>9,877</td>
<td>34.71%</td>
<td>4.60</td>
<td>0.204</td>
<td>0.95 (0.85 to 1.05)</td>
<td>0.320</td>
<td>0.293</td>
</tr>
<tr>
<td>Change in alcohol consumption††</td>
<td>10,269</td>
<td>62.46%</td>
<td>7.99</td>
<td>0.046</td>
<td>0.88 (0.79 to 0.98)</td>
<td>0.022</td>
<td>0.052</td>
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<tr>
<td>Change in smoking¶</td>
<td>11,076</td>
<td>24.09%</td>
<td>3.95</td>
<td>0.257</td>
<td>0.60 (0.44 to 0.80)</td>
<td>0.005</td>
<td>0.144</td>
</tr>
<tr>
<td>Change in sleep duration§§</td>
<td>10,371</td>
<td>61.75%</td>
<td>7.84</td>
<td>0.049</td>
<td>1.06 (0.94 to 1.19)</td>
<td>0.350</td>
<td>0.245</td>
</tr>
</tbody>
</table>

All analyses used a random-effects model meta-analysis. Significant p values which persisted at a false discovery rate of 0.05 are highlighted in bold.

*COVID-19 infection was coded as 0=no, 1=yes.
†Change in financial situation was coded as 0=same or better, 1=worse.
‡Change in trust in government was coded as 0=same or more, 1=less.
§Change in conflict during lockdown was coded as 0=less, 1=more.
¶Change in household composition during lockdown was coded as 0=no, 1=yes.
**Change in psychological distress during lockdown was coded as 1=more in the mental health score, 0=same or lower score.
††Change in alcohol consumption was coded as 1=more drinks per week during lockdown compared with before lockdown, 0=less.
‡‡Change in smoking was coded as 0=less or less cigarettes smoked during lockdown compared with before lockdown, 1=more cigarettes.
§§Change in sleep duration was coded as 0=same or more hours slept during lockdown and 1=less hours slept during lockdown.
employed to derive our outcome variables have not been previously externally validated. COVID-19 infection status included personal suspicion of disease as mass testing was not yet nationally available at the time. This could have led to reporting bias; however, the sensitivity meta-analysis showed that indeed laboratory-confirmed COVID-19 diagnosis was also associated with key worker status. As the change in financial situation, trust in government, conflict and household composition were self-rated, they are prone to perception bias. The change in household composition and change in conflict variables do not take into account individuals who might be living in isolation, although we have adjusted for the receipt of a shielding letter. We did not record key workers according to the Standard Occupational Classification and the Standard Industrial Classification of the Economic Activities, did not capture the type of key workers and thus did not distinguish between frontline workers from others potentially underestimating the burden of lockdown in healthcare workers. As the prevalence of outcomes varied from others potentially underestimating the burden of lockdown workers and thus did not distinguish between frontline workers and key workers according to the Standard Occupational Classification, it is subjected to bias. Meta-analyses with high heterogeneity may largely reflect age group differences as the potential overall between the cohort samples is all. However, we were unable to formally test whether age cohort as a continuous moderator is a source of heterogeneity. The associations between our outcomes and age are probably non-linear, but non-linear meta-regression (eg, using cubic polynomial or splines) would be inappropriate with four data points. On a background of high heterogeneity, the potential of having spurious pooled OR is a possibility. Lastly, performing extensive analyses to separate pandemic effects from known confounders such as seasonal variation was beyond the scope of our study.

Future directions
Further studies are required to assess the multidimensional impact of the COVID-19 pandemic on key workers and their family members. Public health measures implemented to protect key workers should be under continuous review and subject to constant national audits to ensure they are effective and readily accessible.

What is already known on this subject

► Key workers played a monumental role in the UK’s response to the COVID-19 pandemic especially during lockdown. The UK government has taken multiple measures to protect the key workers. However, the impact of the lockdown on key workers is still mostly unknown.

What this study adds

► Our findings show that key workers were more likely to report COVID-19 symptoms or have a COVID-19-positive antibody/antigen test. In addition, they were more likely to experience conflict. However, they were less likely to consume more alcohol or smoke more. Lastly, key worker status was not associated with psychological distress, changes in household composition or trust in the government during lockdown.

CONCLUSION
Being a key worker at the height of the UK national lockdown was a double-edged sword. On one hand, key workers had financial stability and made better lifestyle choices compared with all non-key workers. On the other hand, they were more likely to contract COVID-19 and experience conflict. Despite multiple media reports claiming the contrary, being a key worker was not associated with a change in psychological distress. The UK government had the basic duty to protect its key workers from SARS-CoV-2 infection, but it may have failed to do so, and there is an urgent need to rectify this in light of the ongoing third wave.

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Acknowledgements We are grateful to all the members of our studies for their contribution to this COVID-19 survey and for their ongoing participation in our studies. We thank the survey, data and administrative teams at the Centre for Longitudinal Studies and Unit for Lifelong Health and Ageing, UCL, for enabling the rapid COVID-19 data collection to take place.

Contributors C-CT analysed the data and wrote the manuscript. C-CT, JCM, AH, NC, GC, DB, PP and GC were involved in the study design and implementation. AW, AH, NC, GC, DB and PP actively participated in data acquisition. GC contributed to the data analysis, interpretation of the results and manuscript drafting. All authors were involved in critically revising the manuscript and approved the final version. GC is the guarantor of this work.

Funding The study was funded by the Economic and Social Research Council under the Centre for Longitudinal Studies, Resource Centre 2015–20 (grant number ES/M001660/1), and by the Medical Research Council (grant MC_UU_00019/1). GC is supported by the British Heart Foundation (MyoFit46 Special Programme Grant SP/20/2/34841), the National Institute for Health Research Rare Diseases Translational Research Collaboration (NIHR RD-TRC) and by the NIHR UCL Hospitals Biomedical Research Centre. JCM is directly and indirectly supported by the UCL Hospitals NIHR BRC and Biomedical Research Unit at Barts Hospital, respectively. DB is supported by the Economic and Social Research Council (grant number ES/ M001660/1) and by The Academy of Medical Sciences/Wellcome Trust (“Springboard Health of the Public in 2040” award: HP001/1025). AH receives support from the British Heart Foundation, the Economic and Social Research Council (ESRC), the Horizon 2020 Framework Programme of the European Union, the National Institute on Ageing, the NIHR University College London Hospitals Biomedical Research Centre, the UK Medical Research Council and works in a unit that receives support from the UK Medical Research Council. GC has the support of the European Research Council under the European Union’s Horizon 2020 research and innovation programme (grant agreement no. 819752 DEVORHBIOSHIP–ERC-2018-COG).

Disclaimer None of the funders was involved in the study design, the collection, the analysis, the interpretation of the data and in the decision to submit the article for publication.

Competing interests None declared.

Patient consent for publication Not required.

Ethics approval Ethical approval was obtained from the relevant committees and from the University College London Institute of Education Research Ethics Committee (REC1334).

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Data are available upon reasonable request. Data from the cohorts are available from the UK data archive: https://www.data-archive.ac.uk.

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