

SEX DIFFERENCES IN USE OF INTERVENTIONAL CARDIOLOGY PERSIST AFTER RISK ADJUSTMENT

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Abstract:

Background: Studies from several countries have documented gender disparities in the management of coronary artery disease. We ask whether such gender disparities are seen in Italy and, if so, whether they can be explained by factors such as age and severity of illness.

Methods: 77,974 Piedmontese patients, admitted between 1999 and 2002, with a primary diagnosis of myocardial infarction (ICD 410), angina (ICD 413), chronic ischaemia (ICD 414) and chest pain (ICD 786.5) were studied. The number of males and females undergoing surgical treatment was extracted and the male-female odds ratios calculated. Several risk factors and a risk adjustment technique (APR-DRG) were used to control for possible confounders. Backward stepwise multiple logistic regression was used to adjust for significant covariates.

Results: Crude analysis demonstrated that gender is a discriminating factor in the probability of surgery (OR=2.11, CI 2.04-2.19), with similar findings among those with each main diagnosis. The odds ratios decreased after adjustment for age, co-morbidity, and disease severity but remained significant.

Conclusion: Males and females admitted to hospitals in a region of northern Italy with a diagnosis of cardiovascular disease are treated differently and this cannot be explained by age or severity of disease.

Introduction

Coronary artery disease is the single disorder most likely to kill females in developed countries.^[1-2] There is now evidence from several studies, especially from the USA and United Kingdom^[3-8] but also from Japan^[9] and several continental European countries^[10-13] that, compared to males with cardiovascular disease, females are less likely to receive interventional cardiac procedures. Yet this finding is not universal and some researchers have found no significant gender difference in utilization of either medical therapies^[14-16] or surgical procedures.^[8, 15, 17] Where differences have been found, they have been documented at different stages in the process of accessing investigation and subsequent treatment.^[18]

The contemporary relevance of these findings to Italy can, however, be questioned. First, given how the social meaning of gender is influenced by culture, it cannot be assumed that the phenomenon of gender disparity will be found everywhere. Second, some studies were undertaken when coronary revascularisation was still a relatively new procedure and when cardiovascular disease was still seen as essentially a male disease.^[19] Indeed, it has been argued that the *modus operandi*, both for investigation and for treatment of this disease, was developed primarily for males.^[7, 20] It is plausible that the subsequent expansion of provision could have redressed the observed gender imbalance, as has been seen to some extent in Finland.^[13]

The aim of this study is to investigate whether, in Italy at the beginning of the 21st century, (i) there were gender disparities in the likelihood of being surgically treated for coronary artery disease and, if this is the case, (ii) to ascertain the extent to which age, diagnosis, presence of other risk factors, severity of illness and risk of mortality may account for these differences.

Methods

Data on all discharges from every public and private hospital in the Italian region of Piedmont between 01/01/1999 and 31/12/2002 were obtained. Piedmont is one of the 20 administrative regions of Italy and during those years the population of 4,255,456 was 48% male. In the study period, the hospital network (public and private) comprised about 90 facilities.

Discharge data are collected routinely on all patients admitted to any Italian hospital, with details completed by trained staff (usually physicians or nurses). Each form includes: (i) personal data: name, age, gender, nationality, place of residence; (ii) diagnosis codes (up to 6); (iii) procedures undertaken (up to 6), (iv) length of stay, (v) department of admission and discharge.

Data on diagnosis and procedures are coded using the International Classification of Disease, Ninth Revision, Clinical Modification (ICD-9-CM). Additionally, for each patient admitted in 2000 and in 2002 severity of illness and risk of mortality were calculated using the 3M APR-DRG classification system, version 12. This proprietary software package assigns a value from 0 to 3 to each patient, corresponding to increasing risk. This score is generated using data on age and the presence of co-morbidities, so allowing a meaningful comparison of patients.

Among the 3,320,064 discharge forms collected between 1999 and 2002, we selected, for each year, all patients who were admitted with a primary diagnosis of myocardial infarction (ICD-9-CM code 410.XX), angina (code 413.X), chronic ischaemia (code 414.XX) and chest pain (code 786.5X).

We were interested in the probability of undergoing a first revascularisation; the factors influencing the decision to offer subsequent procedures may be different. Hence, we excluded all records subsequent to one when they underwent a revascularisation.

The choice of diagnoses was guided primarily by the desire to be consistent with previous papers.^[21-22] However, we have also responded to criticism of some earlier papers by including chest pain (which, as with the other diagnoses, is analysed separately). This is because it has been argued that there is a systematic gender bias, in part explained by the tendency for presentation to be “atypical”, that leads to under-diagnosis of chest pain as being of cardiac origin in women.^[8] We do, however, recognize that this is a heterogeneous population, many of whom will not have cardiovascular disease.

Intervention was defined as the presence of codes for either percutaneous coronary intervention (ICD-9-CM code 36.0X) or coronary artery bypass (code 36.1X through 36.2) in one of the six procedure fields of the discharge form.

An initial analysis tabulated the frequencies of each primary diagnosis, age group (≤ 40 , 41-50, 51-60, 61-70, 71-80 and ≥ 81) and secondary diagnosis. As noted above, for 2000 and 2002, the 3M APR-DRG severity of illness (reflecting expected resource use) and risk of mortality were included. The next step was to calculate, for each variable, the percentage, both for males and females, who had been surgically treated.

Where there was an association between gender and revascularization, bivariate analysis was used to estimate the corresponding male to female odds ratios, with 95% confidence intervals, for being treated.

To reduce the influence of potential confounders adjustment for age group, secondary diagnosis of congestive heart failure (code 428.X) and diabetes mellitus (code 250.XX) which, from the literature, emerged as the quantifiable factors most likely to influence treatment decisions. For cases from 2000 and 2002, we were able to generate two models for each diagnosis, each containing a different variable derived from the APR-DRG system. These are severity of illness and risk of mortality. As they are reasonably closely correlated with each other, it was not thought appropriate to use them both in the same model.

Given the likely inter-relationships between potential explanatory variables, multiple logistic regression was used to determine the odds of undergoing revascularization after adjustment for significant covariates, identifying the most parsimonious model. Thus, the regression used backward elimination. The initial model included all variables and those with the highest p values were eliminated at each step until only those that were significant remained.

The two datasets (with and without APR-DRG data) were independently analyzed.

Data management was performed using Access 2003 while all the statistical procedures were carried out using Epi Info version 3.3.2 and Stata version 8.0.

Results

77,974 cases fulfilled the inclusion criteria, equally distributed throughout the study period with a mean of 19,493 cases recorded each year (Standard Deviation (SD) 417) and females comprised around 32% each year. Over the whole study period 80.5% of patients were admitted only once whereas readmissions, accounted for 19.5% of cases.

The distribution of key variables within the cases meeting the inclusion criteria is shown in Table 1, with the numbers in parenthesis representing the percentages of patients undergoing surgical procedures. The length of stay variable, in the lowest rows, shows the mean of the hospitalization days, for each diagnosis and the corresponding standard deviation.

Female patients were, on average, older than male patients, respectively 71.4 (SD 12.5) and 64.2 (SD 12.0). Among those undergoing surgery, males had a mean age of 62.6 (SD 10.2) and females of 67.7 (SD 9.41) (a difference of 5.1 years) $p < 0.001$ while among all those hospitalized who did not have revascularization the figures were respectively 64.9 (SD 12.7) and 72.2 (SD 12.9) years old (a difference of 7.3 years) $p < 0.001$.

There was an upward trend over time in the percentage of patients undergoing interventions, especially among those with diagnoses of angina or myocardial infarction, where operation rates more than doubled, increasing respectively from 14.7% in 1999 to 32.3% in 2002 for males and from 7.2% to 17.1% for females in the same period.

Table 2 shows the crude and adjusted odds ratios (OR) for revascularization procedures among males compared with females, according to potential explanatory variables. Crude analysis demonstrated that, for all diagnoses combined, being male is an important discriminating factor in whether or not one undergoes revascularization (OR=2.11, 95% CI 2.04 – 2.19, $p < 0.001$) (table 3). Broadly comparable findings were obtained within each diagnostic category: myocardial infarction

OR 2.09 (95% CI 1.96 – 2.23); angina OR 1.75 (95% CI 1.61 – 1.90); chronic ischaemia, OR=2.14 (95% 2.01-2.28), and chest pain OR 3.95 (CI 1.52 – 10.24). (table 2).

In all cases, the gap narrows when adjusted for age, but remains highly significant for all diagnostic categories. Adjustment for the presence of heart failure or diabetes makes a relatively small difference. In the sub-set of data with additional information on severity, adjustment again makes relatively little difference.

Table 3 shows the fully adjusted model, before and after backward elimination of variables starting with age group and presence or absence of diabetes or heart failure, plus, for 2000 and 2002, the two severity scores. In the initial analyses, using all four years of data, all three variables were retained, except for chest pain (which had the fewest numbers and the least likelihood of co-morbidity), where they were all eliminated. In the restricted data set from 2000 and 2002, among those with myocardial infarctions the two final models retained only age and severity or age and risk of mortality, suggesting that the presence of heart failure or diabetes as a factor that would reduce the probability of intervention is viewed similarly in males and females. In contrast, heart failure (and diabetes for those with angina after adjustment for mortality risk) do seem to have a different meaning in males and females with angina or chronic ischaemia. Chest pain does not seem to be influenced by any of the studied variables although this is likely to reflect both the low numbers of cases, with some age groups containing no patients, and the low rate of co-morbidity. Comparing the crude male female odds ratios (table 2) with those in the fully adjusted models (table 3), it can be seen that, with the exception of the small number of patients with a diagnosis of chest pain, inclusion of all significant correlates reduces the gender gap (myocardial infarction 2.09 to 1.33; angina 1.75 to 1.61 and chronic ischaemia 2.14 to 1.48; but in all cases the greater probability of intervention in males remains highly statistically significant. Similar results were obtained using the sub-set of 2000 and 2002 data, when additional severity measures were included.

Discussion

In the early 21st century, Italian females remain less likely than males to undergo surgical procedures for coronary artery disease. This inequity is consistent with previous research undertaken in the USA, United Kingdom and elsewhere. However, this study goes beyond many of the earlier studies by showing that the disparity persists after taking account of differences in severity of illness. We believe that this is particularly relevant as earlier research has given rise to the term “Yentl Syndrome”,^[23] named after the film in which a woman disguises herself as a man to become a rabbi. It is suggested that “*Once a woman showed that she was just like a man, by having severe coronary artery disease . . . , then she was treated as a man would be*”. Thus, we would expect that the odds ratios would narrow markedly after adjustment for severity. There was some narrowing but, even after the inclusion of all significant variables, including sophisticated measures of disease severity, females were still less likely to undergo revascularization.

It is beyond the scope of this study, which is based on administrative data, to explain why these inequities exist. Authors of previous studies proposed two sets of hypotheses to explain these gender differences. The first focuses on patient preferences. This is based on the premise that some females do not see themselves as candidates for revascularization because they do not consider themselves to be at risk of coronary heart diseases. A study addressing this issue reported that all the females involved expressed surprise when informed of their diagnosis. The most obvious consequence of this finding is that, if a woman does not see herself as being at a risk of cardiovascular disease, she will tend to interpret symptoms as being due to other reasons.^[24] It has also been reported that females are more likely than males to decline a major procedure.^[20, 25]

Another factor is that females may present with symptoms that are “atypical”, which has been defined largely on the basis of studies among males. It has been found that females experiencing an acute myocardial infarction are more likely than males simultaneously to experience gastrointestinal symptoms such as indigestion,^[26] while it has also been suggested that there may be difficulty

distinguishing some forms of angina from breast tenderness secondary to cyclical hormonal changes in pre-menopausal females or to hormone replacement therapy in postmenopausal females.^[27] Another study concluded that, while there were no major differences in the language used by females and males to report pain, females may localize it differently and may be more likely to report supplementary symptoms.^[8] We tried to consider this issue, at least partially, by including patients initially diagnosed as having chest pain rather than a specific cardiac cause. Even in this much broader group of patients, females continue to be less likely to be operated on.

The second group of hypotheses focuses on decisions made by physicians. There is now clear evidence that the gender of a patient may influence the response of physicians.^[10] Writers from a feminist perspective have argued that this reflects perceptions of the value that each gender contributes to society, with some male physicians believing that men contribute more than females.^[28]

It has also been suggested, as noted above, that bias may arise from the use of diagnostic criteria developed and validated on men, which may thus be less sensitive and specific in females,^[19] so potentially misleading the diagnostician. For instance, females presenting with chest pain are less likely to have a positive exercise test and, among those who have a positive exercise test, females are still less likely to have coronary artery disease.^[29] In addition, some studies suggest that females have more complications and benefit less following surgery.^[20]

Unfortunately, it is not possible to test these hypotheses using these data. This would require other techniques such as ethnography or decision-analysis methods that could explore the decision-making processes adopted by patients and physicians. Nor is it possible to eliminate entirely some forms of artifact. There may also be valid clinical considerations that were undocumented in the clinical records but these would require a prospective study which would be fraught with methodological and ethical problems. Finally, we cannot exclude the possibility that relevant co-morbidities and contraindications were inadequately documented by medical staff. It is, however, unlikely that there would be selective differences in recording by gender.

Our study exhibits both strengths and weaknesses. A major strength is that retrospective analysis of patients' records ensured that clinical practice was not influenced by the research process, thereby biasing findings but, on the other hand, our findings might be subject to the problems known to affect any study using routine information, such as incompleteness of clinical coding, data inaccuracies and failure to link episodes to individuals. Nevertheless, we believe that these risks are minimized by the presence in each facility of strict quality control systems monitoring the quality of data collection, albeit that the diligence with which these processes are undertaken reflects the need to monitor financial rather than clinical performance. Although we were able to remove all records of people who had records of previous revascularization within the data set, we cannot exclude the possibility that a small number may have undergone treatment either elsewhere or before the study years. It is, however, difficult to see, given the small numbers likely to be involved, how this could materially affect the results.

A second advantage is the inclusion of procedures in both public and private hospitals, especially in light of a recent study from the UK showing how private sector provision exacerbates social inequalities in access to care.^[30]

A third advantage of our study is the inclusion of two measures of risk adjustment. In our study, males and females differ significantly in age. This is to be expected, given that coronary artery disease presents earlier in males^[31] but could introduce a bias into the results. In calculating the severity of illness rank, the APR-DRG software takes into account not only relevant co-morbidities, but also the interaction between these co-morbidities and age or the principal diagnosis.^[32] This is particularly useful as, this procedure allows a meaningful and fairer, comparison of the two groups.

A disadvantage is that we are able to look only at patients once they have reached hospital. We cannot exclude the possibility that women might be disadvantaged further by failure to refer them for investigation. This is less likely in cases of myocardial infarction but is certainly possible for angina and chronic ischaemia.

This study has found a substantial gender disparity in the probability of undergoing revascularization among patients with known cardiovascular disease in a large Italian region. This cannot be accounted for by documented differences in age or severity of illness. This is, to our knowledge, the first time that this phenomenon has been documented in Italy. Furthermore, despite evidence of disparities from elsewhere, we are unaware of any significant discussion about this issue in Italy, either among patient or professional groups. While more research is clearly needed to understand these observations better, we hope that this study will at least enable the discussion to begin.

What this paper adds:

-what is already know of this subject?

Previous studies documented gender differences in the likelihood of being treated surgically for ischemic heart disease. Many researchers suggested that women are less likely to be operated because they usually show a less severe picture. This phenomenon, known as the Yentl syndrome, implies that once a woman assumes certain characteristics associated with men (such as classical clinical presentation) , she will receive the same treatment.

-What does this paper add?

Our results, elaborated on a very large dataset, confirm gender differences for a number of cardiac - related diagnoses. Furthermore, we find that inequalities persist after adjustment for severity of illness and risk of mortality.

Competing interest

All authors declare that the answers to the questions on your competing interest form (<http://bmj.com/cgi/content/full/317/7154/291/DC1>) are all no and therefore have nothing to declare.

Licence statement

The authors declare that they have the permission to use the dataset for the study from the Health Regional Agency of Piedmont Region.

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Table 1 - Number of hospitalized patients and percentage of operated – Length of stay and standard deviation (SD)

		1999				2000				2001				2002			
		Males		Females		Males		Females		Males		Females		Males		Females	
		Number	(%)	Number	(%)	number	(%)	number	(%)	Number	(%)	number	(%)	number	(%)	number	(%)
Diagnosis	Total	13202	(23.4)	6316	(12.4)	13266	(28.6)	6367	(15.7)	13402	(32.5)	6503	(19.3)	12722	(36.7)	6196	(20.7)
	Myoc infarct ^a	4624	(14.7)	2165	(7.2)	5170	(23.2)	2405	(11.4)	5220	(27.8)	2710	(17.3)	5121	(32.3)	2659	(17.1)
	angina	2852	(20.8)	1197	(11.4)	2971	(33.4)	1329	(23.6)	2957	(37.1)	1253	(25.6)	2352	(39.9)	1027	(26.2)
	Chr ischae ^b	4317	(42.1)	2004	(24.5)	3668	(43.4)	1650	(24.7)	3884	(46.1)	1552	(30.0)	3968	(51.8)	1577	(35.3)
Age group	chest pain	1409	(0.1)	950	(0)	1457	(0.4)	983	(0.2)	1341	(0.5)	988	(0.1)	1281	(1.1)	933	(0.2)
	<=40	397	(16.6)	111	(9.9)	361	(19.4)	114	(12.3)	432	(20.8)	92	(12.0)	338	(23.4)	86	(10.5)
	41-50	1397	(27.4)	263	(14.1)	1326	(33.2)	253	(17.4)	1272	(36.7)	232	(20.7)	1214	(42.3)	240	(22.5)
	51-60	3069	(28.0)	727	(15.8)	3140	(33.9)	792	(19.0)	2901	(38.9)	800	(23.0)	2758	(40.9)	685	(24.7)
	61-70	4413	(26.1)	1651	(19.0)	4407	(31.4)	1625	(23.8)	4470	(36)	1666	(27.8)	4118	(40.7)	1564	(29.5)
	71-80	3025	(19.8)	2084	(13.4)	3198	(24.6)	2097	(17.7)	3441	(28.6)	2297	(21.5)	3342	(34.4)	2118	(23.1)
Second diagnosis	>81	901	(4.1)	1480	(1.8)	834	(5.8)	1486	(2.1)	886	(7.8)	1416	(3.9)	952	(12.3)	1503	(6.5)
	CHF ^c yes	671	(11.0)	470	(5.1)	688	(13.2)	554	(6.3)	701	(16.0)	520	(8.8)	673	(16.2)	600	(7.2)
	no	12531	(24.1)	5846	(13.0)	12578	(29.4)	5813	(16.6)	12701	(33.4)	5983	(20.2)	12049	(37.8)	5596	(22.1)
	DM ^d yes	1205	(22.6)	871	(13.6)	1498	(31.4)	994	(19.6)	1614	(31.5)	1100	(20.4)	1394	(32.8)	1021	(20.5)
no	11997	(23.5)	5445	(12.2)	11768	(28.2)	5373	(14.9)	11788	(32.6)	5403	(19.1)	11328	(37.1)	5175	(20.7)	
APR-DRG Sev of iln	1	-	-	-	-	8909	(30.3)	3794	(18.6)	-	-	-	-	8441	(40.4)	3646	(25.4)
	2	-	-	-	-	3348	(24.4)	1852	(11.1)	-	-	-	-	3265	(28.2)	1846	(13.2)
	3	-	-	-	-	802	(28.9)	574	(12.0)	-	-	-	-	798	(35.5)	539	(15.2)
	4	-	-	-	-	206	(20.4)	147	(10.2)	-	-	-	-	218	(23.4)	165	(17.0)
APR-DRG Rsk of mort	1	-	-	-	-	10657	(31.2)	4568	(18.3)	-	-	-	-	10032	(40.5)	4356	(24.7)
	2	-	-	-	-	1749	(17.9)	1164	(9.8)	-	-	-	-	1814	(23.1)	1175	(11.7)
	3	-	-	-	-	359	(18.4)	245	(8.2)	-	-	-	-	408	(19.1)	333	(9.3)
	4	-	-	-	-	500	(16.2)	390	(7.4)	-	-	-	-	468	(22.7)	332	(11.4)
Length of stay	Mean		SD	Mean	SD	mean	SD	Mean	SD	Mean	SD	mean	SD	mean	SD	mean	SD
	myoc infarct ^a	10.1	7.1	11.2	9.6	9.4	6.8	11.3	10.0	8.5	7.2	10.1	9.4	8.1	6.6	9.6	9.9
	angina	7.0	7.5	7.6	7.0	6.6	7.2	7.3	7.9	6.0	6.6	6.7	8.0	5.6	6.7	6.4	7.3
	chr ischae ^b	9.9	24.1	12.7	21.1	8.9	17.8	13.8	18.8	8.1	15.6	12.7	18.1	8.5	14.1	12.1	17.3
chest pain	4.1	6.0	4.5	5.8	4.1	6.0	5.0	11.4	3.9	5.5	4.5	7.4	3.4	3.4	3.8	3.3	
Note:	myocardial infarction ^a : 410.XX ICD9-CM				chest pain: 786.5X ICD9-CM				angioplasty: 36.0X ICD9-CM				coronary artery bypass: 36.1X-36.2X ICD9-CM				
	angina: 413.XX ICD9-CM				congestive heart failure ^c : 428.X ICD9-CM												
	chronic ischaemia ^b : 414.XX ICD9-CM				diabetes mellitus ^d :250.XX ICD9-CM												

Table 2 – Crude and adjusted Odds Ratio for Revascularization in males vs females									
		MYOCARDIAL INFARCTION (ICD-9: 410.XX)		ANGINA (ICD-9: 413.XX)		CHRONIC ISCHAEMIA (ICD-9: 414.XX)		CHEST PAIN (ICD-9: 786.5X)	
		OR (CI 95%)	P-value	OR (CI 95%)	P-value	OR (CI 95%)	P-value	OR (CI 95%)	P-value
Crude		2.09 (1.96-2.23)	<0.0001	1.75 (1.61 - 1.90)	<0.0001	2.14 (2.01– 2.28)	<0.0001	3.95 (1.52 – 10.24)	0.0023
all years adjusted for	<i>Age band</i>	1.35 (1.26-1.45)	<0.0001	1.59 (1.46 - 1.72)	<0.0001	1.49 (1.39-1.59)	0.0001	4.07 (1.60-10.35)	0.0014
	<i>Second. diag. diabetes</i>	2.07 (1.94-2.21)	<0.0001	1.77 (1.64-1.92)	<0.0001	2.13 (2.0-2.26)	<0.0001	3.96 (1.54 – 10.16)	0.002
	<i>Second.diag heart failure</i>	1.99 (1.86-2.13)	<0.0001	1.74 (1.61-1.89)	<0.0001	2.13 (2.0-2.27)	<0.0001	3.95 (1.52-10.25)	0.0023
Crude		2.28 (2.08-2.50)	<0.0001	1.73 (1.55-1.94)	<0.0001	2.15 (1.97-2.35)	<0.0001	3.52 (1.20-10.30)	0.0014
2000 and 2002 adjusted for:	<i>Age bands</i>	1.47 (1.34-1.62)	<0.0001	1.58 (1.41-1.76)	<0.0001	1.52 (1.38-1.67)	<0.0001	3.92	0.0076
	<i>Second. diag diabetes</i>	2.26 (2.06-2.48)	<0.0001	1.77 (1.58-1.97)	<0.0001	2.14 (1.95-2.34)	<0.0001	3.53 (1.22-10.24)	0.0132
	<i>Second.diag heart failure</i>	2.15 (1.96-2.36)	<0.0001	1.73 (1.55-1.93)	<0.0001	2.14 (1.95-2.33)	<0.0001	3.51 (1.20-10.31)	0.0146
	<i>Severity of illness</i>	1.99 (1.81-2.18)	<0.0001	1.76 (1.57-1.97)	<0.0001	2.12 (1.94-2.32)	<0.0001	3.32 (1.12-9.76)	0.0211
	<i>Risk of death</i>	2.01 (1.83-2.20)	<0.0001	1.73 (1.55-1.93)	<0.0001	2.08 (1.90-2.27)	<0.0001	3.49 (1.19-10.24)	0.0152

Table 3 - Models of the relationship between probability of intervention and gender adjusted for possible confounders									
Year	Pathology	starting model	OR	(95% CI)	P	final model	OR	(95% CI)	P
99-00-01-02	All diagnoses		2.11	(2.04-2.19)	<0.001				
	M. infarction	hrtfail+diab+clsage	1.33	(1.24 - 1.43)	<0.001	hrtfail+diab+clsage	1.33	(1.24 - 1.43)	<0.001
	Angina		1.61	(1.48 - 1.75)	<0.001		1.61	(1.48 - 1.75)	<0.001
	Chron isch		1.48	(1.39 - 1.58)	<0.001		1.48	(1.39 - 1.58)	<0.001
	Chest pain		4.30	(1.65 - 11.23)	0.003		-	3.95	(1.52 - 10.24)
2000 and 2002	All diagnoses		2.18	(2.06 - 2.29)	<0.001				
	M. infarction	hrtfail+diab+clsage+sevill	1.43	(1.30 - 1.58)	<0.001	clsage+sevill	1.43	(1.30 - 1.58)	<0.001
		hrtfail+diab+clsage+riskmort	1.44	(1.31 - 1.59)	<0.001	clsage+riskmort	1.45	(1.31 - 1.59)	<0.001
	Angina	hrtfail+diab+clsage+sevill	1.55	(1.38 - 1.12)	<0.001	clsage+sevill+hrtfail	1.56	(1.39 - 1.75)	<0.001
		hrtfail+diab+clsage+riskmort	1.61	(1.44 - 1.80)	<0.001	hrtfail+diab+clsage+riskmort	1.61	(1.44 - 1.80)	<0.001
	Chron isch	hrtfail+diab+clsage+sevill	1.53	(1.39 - 1.69)	<0.001	clsage+hrtfail+sevill	1.54	(1.40 - 1.69)	<0.001
		hrtfail+diab+clsage+riskmort	1.52	(1.38 - 1.67)	<0.001	clsage+hrtfail+riskmort	1.53	(1.39 - 1.68)	<0.001
	Chest pain	hrtfail+diab+clsage+sevill	3.69	(1.24 - 10.69)	0.019	-	3.52	(1.20 - 10.30)	0.022
hrtfail+diab+clsage+riskmort		3.88	(1.31 - 14.47)	0.014	-	3.52	(1.20 - 10.30)	0.022	

Note: hrtfail: congestive heart failure diab: diabetes mellitus clsage: age group
sevill: severity of illness riskmort: Risk of mortality

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