How does income affect mental health and health behaviours? A quasi-experimental study of the earned income tax credit

Laura Shields-Zeeman,1,2 Daniel F. Collin,1 Akansha Batra,3 Rita Hamad 1,4

ABSTRACT

Background Although research has repeatedly demonstrated the association between poverty, mental health, and health behaviours, there is limited evidence on the effects of interventions to improve these outcomes by addressing poverty directly. Moreover, most prior studies are often confounded by unobserved characteristics of individuals, making it difficult to inform possible interventions. We addressed this gap in the literature by leveraging quasi-random variation in the earned income tax credit (EITC)—the largest US poverty alleviation programme for families with children—to examine the effects on overall health, psychological distress, smoking, and alcohol consumption.

Methods We used a large diverse national sample drawn from the Panel Study of Income Dynamics (N=34,824). We first conducted ordinary least squares (OLS) models to estimate the association of income and the EITC with the outcomes of interest. We subsequently employed a quasi-experimental instrumental variables (IV) analysis—in which EITC refund size was the instrument—to estimate the effect of income itself.

Results In OLS models, higher income was associated with reductions in psychological distress, increased drinking, increased smoking, and more cigarettes per day, and larger EITC refunds were associated with reductions in psychological distress. In IV models, higher income was associated with decreased psychological distress.

Conclusion These results suggest that typical correlational studies of the health effects of income may be confounded, although results may not generalise to income distributed in different ways than the EITC. The findings also provide valuable information for policymakers and researchers seeking to address socioeconomic disparities in mental health.

INTRODUCTION

It is well recognised that income, and conversely, poverty, are important social determinants of physical and mental health.1 In particular, poverty is associated with greater risk of common mental disorders.2–4 In the US, the prevalence of depression is higher for those in the lowest (9.8%) compared with the highest (5.7%) income groups.4 There are numerous hypothesised mechanistic pathways that may explain these associations (figure 1). Poverty is thought to reduce individuals’ ‘cognitive bandwidth’, which can result in changes to decision-making abilities and engagement in risky health behaviours.5 Low-income populations may also have greater exposure to chronic stressors, amplifying the risk of developing or exacerbating mental health problems.6 Poverty or low socioeconomic status has been associated with a greater likelihood of smoking and problematic alcohol use,2,7 perhaps because substance use may be a coping mechanism for stress.8 Moreover, increased income can also be used to purchase more cigarettes and alcohol.9 The type of income may also matter: some studies have shown that infrequent cash benefits increase expenditures on ‘temptation’ goods such as alcohol relative to more frequent benefits.10 Income insecurity may impact healthcare utilisation (eg, reduced access to, use of, and quality of mental healthcare).11

Despite the associations between income and health, existing evidence is often correlational, and there are limited studies—especially in the US—that test the effects of interventions or policies to address poverty on mental health and health behaviours. Policies focusing on reducing social and economic risk factors have the potential to improve health disparities, although traditional study designs for poverty alleviation interventions face challenges due to the expense of administering financial interventions and long-term follow-up of participants. Quasi-experimental methods that leverage existing variation in economic policies may be helpful in examining the links between income and health. In the US, the earned income tax credit (EITC) is the largest poverty alleviation programme for families with children, with an estimated 25 million recipients receiving over US$60 billion in 2018.12 Initiated in 1975, the EITC is a federal programme that provides tax refunds to low-income working families. Eligibility is based on earned income, family size, marital status, and age. The average payment size for a family with children is approximately US$3000, distributed as a tax refund.13 The EITC has been shown to reduce poverty,14,15 and improve housing outcomes.16 Several studies have examined the EITC’s health effects, finding improvements in food security,17 health insurance coverage,18 infant birth weight, and other child health outcomes.19,20 Studies of the effect on mental health and health behaviours are limited. One study found that state EITC programmes reduced suicides among low-educated adults,21 while another found improved mental health among EITC-eligible mothers.22 The EITC has been associated with reduced tobacco use,23 particularly among mothers,19 although one study found that higher income from the EITC is associated with an increase in the number...
Two prior studies found that the EITC had no effect on alcohol use. Nearly all these studies used historical data prior to 2000, and most estimated EITC eligibility using proxies such as low educational attainment or ecological measures of state EITC implementation. In the context of a shifting political climate and changes to other safety net programmes, contemporary evidence is needed to understand the effects of income and economic policies on mental health and health behaviours.

The present study helps to fill these gaps in the literature. The aim was to assess the effect of an economic policy on mental health and health behaviours among working low-income families in the United States, leveraging variation in EITC policy through a quasi-experimental analysis in a large national sample of US adults with granular information on EITC eligibility.

METHODS

Sample

Data were drawn from the Panel Study of Income Dynamics (PSID), a large national household cohort study. The PSID has been conducted annually since 1968 and biennially since 1997. We restricted our sample to data collected since 1985 (the earliest year when at least one of the outcomes of interest was available) through 2015 (the most recent survey wave available when data analysis began) (N=65,047). We further restricted our sample to heads of households and spouses, as these were the participants for whom PSID asked the outcomes of interest (N=36,262). Next, we restricted our analysis to those with data on at least one of the outcomes of interest (N=35,481). Finally, we restricted the sample to observations with complete cases (N=34,824). Because PSID is a cohort study, individuals are interviewed repeatedly over time, resulting in 289,877 total observations. We account for the panel nature of the data set with the inclusion of individual-level fixed effects and clustered standard errors (SEs, described below).

Health outcomes

We examined psychological distress, overall health, alcohol use, and smoking. Psychological distress was assessed using the Kessler-6 scale (K6), a widely used screening tool for mental health problems. Higher scores indicate higher levels of psychological distress (range 0–24). Overall health, self-reported as a 5-item scale, was coded as a binary variable indicating very good or excellent health versus good, fair, or poor health. Self-rated health is a validated predictor of mental health, health behaviours, and mortality. Alcohol was assessed as current consumption (binary) and whether the individual consumed 3+ drinks per day (binary, zero for non-drinkers). While the latter does not correspond directly to binge drinking (defined as 4+ drinks per day for women and 5+ for men), PSID did not consistently include a continuous measure of the number of drinks consumed. Smoking was assessed as currently smoking (binary) and the number of cigarettes per day (continuous, zero for non-smokers).

Of note, the K6 was asked for either the head of household or spouse (preferentially the head), while other outcomes were available for both the head and spouse. Also, not all outcomes were assessed in all survey waves (online supplemental table 1), resulting in different numbers of observations for each regression (table 1).

Instrument

The amount of EITC refund for which an individual was eligible, including both federal and state refunds, served as the instrumental variable for the analysis described below. This amount was computed using the Taxsim27 package for Stata, which calculates

---

Table 1: Sample characteristics, Panel Study of Income Dynamics, 1985–2015

<table>
<thead>
<tr>
<th>Panel A: demographics</th>
<th>Mean (SD) or %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female</td>
<td>54.5</td>
</tr>
<tr>
<td>Head of household race</td>
<td></td>
</tr>
<tr>
<td>White</td>
<td>54.3</td>
</tr>
<tr>
<td>Black</td>
<td>33.3</td>
</tr>
<tr>
<td>Hispanic</td>
<td>10.4</td>
</tr>
<tr>
<td>Other</td>
<td>1.9</td>
</tr>
<tr>
<td>Age (years)</td>
<td></td>
</tr>
<tr>
<td>19.4 (15.7)</td>
<td></td>
</tr>
<tr>
<td>Head of household married</td>
<td>53.2</td>
</tr>
<tr>
<td>Education (completed)</td>
<td></td>
</tr>
<tr>
<td>Less than high school</td>
<td>19.4</td>
</tr>
<tr>
<td>Completed high school</td>
<td>37.6</td>
</tr>
<tr>
<td>Completed some college</td>
<td>22.4</td>
</tr>
<tr>
<td>Completed college</td>
<td>20.7</td>
</tr>
<tr>
<td>Number of children</td>
<td>0.9 (1.2)</td>
</tr>
<tr>
<td>Household pre-tax earned income* (US$)</td>
<td>50,423 (73,175)</td>
</tr>
<tr>
<td>Eligible for EITC</td>
<td>31.7</td>
</tr>
<tr>
<td>EITC refund amount* (US$)</td>
<td>1,701 (1,577)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Panel B: outcomes</th>
<th>Mean (SD) or %</th>
<th>No. person-years</th>
</tr>
</thead>
<tbody>
<tr>
<td>Psychological distress</td>
<td>3.5 (4.1)</td>
<td>59,734</td>
</tr>
<tr>
<td>Very good/excellent overall health</td>
<td>53.3</td>
<td>289,611</td>
</tr>
<tr>
<td>Currently drink alcohol</td>
<td>60.2</td>
<td>116,066</td>
</tr>
<tr>
<td>Consume three or more drinks per day</td>
<td>13.9</td>
<td>115,015</td>
</tr>
<tr>
<td>Currently smoke</td>
<td>21.3</td>
<td>127,246</td>
</tr>
<tr>
<td>Cigarettes per day</td>
<td>2.9 (7.2)</td>
<td>126,944</td>
</tr>
</tbody>
</table>

N=38,24 persons, 289,877 person-years. *Inflation-adjusted. EITC, earned income tax credit.
refund amounts using Internal Revenue Service formulas.\textsuperscript{29} Inputs included the head of household’s age and marital status, household pre-tax earned income, number of children, state of residence, and tax year. Since PSID does not query participants about actual EITC refunds received, we assumed that all eligible households received the tax refund for which they were eligible, an approach used in numerous studies of the EITC.\textsuperscript{17,20,30} This approach may result in measurement error since prior work has shown that roughly 80% of families receive the EITC refund for which they are eligible\textsuperscript{31}; however, it is analogous to an intent-to-treat analysis and an alternative to previous studies that used low educational attainment or state EITC implementation as proxies for EITC receipt.\textsuperscript{19,21} Coefficients are presented per $1000 of EITC, for ease of interpretation.

Covariates

Covariates included age and age-squared, marital status, education, number of children, inflation-adjusted pre-tax income and income-squared, and indicator variables for year to account for secular trends. Time-invariant covariates like gender and race were not included given the use of individual-level fixed effects, described below.

Data analysis

We first tabulated participant characteristics. We then carried out two types of analyses: (1) ordinary least squares (OLS), which is subject to confounding of the relationship between income and health by unobserved individual characteristics, and (2) instrumental variables (IV) analysis, which is designed to address such confounding (figure 2).

In the first set of OLS models, we regressed each outcome on self-reported income, adjusting for the covariates above. To account for the panel nature of PSID, in which participants are interviewed repeatedly over time, we included fixed effects (ie, indicator variables) for each individual. This technique leverages within-person (rather than across-person) variation in the EITC, which accounts for unobserved time-invariant individual characteristics and strengthens causal inference. Robust SEs were clustered at the individual level to account for correlated observations.

In a second set of OLS models, we regressed each health outcome on EITC refund size, again adjusting for covariates and clustering SEs at the individual level. This model is considered the ‘reduced form’ of the IV analysis described below.

OLS models are likely to suffer from confounding by unobserved individual factors like socioeconomic status or genetic endowment (figure 2, U1). To address this confounding, we employed IV analysis, a well-established quasi-experimental technique.\textsuperscript{32} IV analysis relies on a quasi-randomly assigned variable or ‘instrumental variable’—in this case, the size of the EITC refund for which an individual is eligible (Z)—which impacts the health outcome (Y) through the exposure (X, post-tax household income). This method is useful when the exposure cannot easily be randomised, as in the case of income. IV analysis assumes that there is not a separate pathway linking the instrument and outcome other than through the exposure, nor any unmeasured characteristics that confound the relationship between the instrument and the exposure or outcome. Although these assumptions cannot be empirically verified, we conducted analyses to test these assumptions. Since variation in EITC refund size over time is driven by exogenous policy changes that act through a small set of measured characteristics, which we include as covariates in regressions (C1, eg, pre-tax earned income, marital status), and because we calculate the EITC refund size for which an individual was eligible rather than actual receipt (analogous to an intent-to-treat design), the analysis is, therefore, less likely to suffer from confounding by other measured and unmeasured individual characteristics (C2, U1).\textsuperscript{33} Additional details and equations are provided in the online supplemental file.
We used linear models for all analyses. While logistic models are often more appropriate than linear models when evaluating binary outcomes, in this case, they failed to converge in the presence of individual-level fixed effects due to the large number of variables. Because fixed effects models rely on within-person comparisons and therefore strengthen causal inference, we opted to use linear models with fixed effects even for binary outcomes. The coefficient in these models is therefore interpreted as the change in per cent risk of the outcome per $1000 of income.

**Secondary analyses**

We performed several additional analyses to test the robustness of results to multiple specifications. First, as prior studies have shown that the EITC may affect single and married individuals differently, we conducted a stratified analysis with separate models for single and married individuals. Second, we conducted an analysis restricted to participants with annual household incomes greater than US$0 and less than US$100,000 (N=29,813), as these are more likely to represent an appropriate comparison group for EITC recipients. Finally, we sought to account for possible residual confounding by unmeasured state-level characteristics (figure 2, U2); approximately half of states provide an additional EITC refund to supplement the federal refund, and these states may have other characteristics leading to differences in health. To eliminate the possibility of state-level residual confounding, we used only the federal EITC refund size as the instrument. In this analysis, any effect is based solely on variation in the federal refund, thereby removing the possibility of confounding by state characteristics.

**RESULTS**

**Sample characteristics**

About half of the participants were women, 54.3% were white, 33.3% were black, and 10.4% were Hispanic (table 1). The mean age was 43.7 years (SD 15.7), and 53.2% were married. Approximately 37.6% completed high school. The average household pre-tax income was US$50,423 (SD 73,175), and a quarter of participants were eligible for the EITC at some point during the study period. The average income for EITC-eligible recipients was US$18,061 (SD 11,975), and the average amount of EITC refund for which participants were eligible was US$1701 (SD 1577). The EITC refund amounted to 9.4% of recipients’ annual pre-tax income on average. The mean psychological distress score was 3.5 (SD 4.1), and 53.3% reported very good or excellent overall health. About 60.2% consumed alcohol, and 13.9% consumed 3+ drinks per day. About 21.3% smoked, and the mean number of cigarettes smoked per day was 2.9 (SD 7.2).

**Association of income and EITC with outcomes (OLS models)**

Higher income was associated with reduced psychological distress (−0.026 per US$1000, 95% CI: −0.035 to −0.018), increased probability of consuming alcohol (0.0098, 95% CI: 0.00047 to 0.0015), and increased cigarettes smoked per day (0.012, 95% CI: 0.0056 to 0.019) (table 2, Model 1).

Meanwhile, larger EITC refunds were associated with reduced psychological distress (−0.070 per US$1000, 95% CI: −0.11 to −0.031) but not other outcomes (table 2, Model 2).

**The effect of income on outcomes (IV models)**

In IV models, higher income was associated with decreased psychological distress (−0.13 per US$1000, 95% CI: −0.21 to −0.059) (table 2, Model 3). This represents a decrease of about 3%–4% of a SD in the K6 score. We were unable to reject the null hypothesis that there was no association of income with overall health or alcohol and tobacco use.

**Secondary analyses**

When stratifying by marital status, EITC refund size was associated with reduced psychological distress among single households (−0.098 per US$1000, 95% CI: −0.15 to −0.047), and this was statistically significantly different from the estimate for married households (0.16, 95% CI: −0.27 to 0.59, p<0.001) (table 3, Models 1–2). For other outcomes, estimates for single and married individuals were similar.

Restricting the sample to households with income over US$0 and under US$100,000 (table 3, Model 3), we were unable to reject the null hypothesis of no association of income with any

---

**Table 2  Association of income and the EITC refund size with mental health and health behaviours**

<table>
<thead>
<tr>
<th>Outcome</th>
<th>Model 1: association of income with health (OLS)</th>
<th>Model 2: association of EITC with health (OLS)</th>
<th>Model 3: effect of income on health (IV)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Psychological distress</td>
<td>−0.026** (−0.035 to −0.018)</td>
<td>−0.070 (−0.11 to −0.031)</td>
<td>−0.13** (−0.21 to −0.059)</td>
</tr>
<tr>
<td>Good/excellent health</td>
<td>0.0013* (0.00085 to 0.0018)</td>
<td>0.0000046 (−0.0027 to 0.0027)</td>
<td>−0.0028 (−0.0096 to 0.0040)</td>
</tr>
<tr>
<td>Currently drink alcohol</td>
<td>0.00098** (0.00047 to 0.0015)</td>
<td>0.0025 (−0.0023 to 0.0025)</td>
<td>0.011 (−0.027 to 0.025)</td>
</tr>
<tr>
<td>3+ drinks per day</td>
<td>−0.00012 (−0.00053 to 0.00030)</td>
<td>−0.0016 (−0.00039 to 0.00078)</td>
<td>−0.0071 (−0.018 to −0.0038)</td>
</tr>
<tr>
<td>Currently smoke</td>
<td>0.0018 (−0.00013 to 0.00050)</td>
<td>0.00086 (−0.0012 to 0.0029)</td>
<td>0.0044 (−0.0060 to 0.015)</td>
</tr>
<tr>
<td>Cigarettes per day</td>
<td>0.012** (0.0056 to 0.019)</td>
<td>0.0022 (−0.0033 to 0.0083)</td>
<td>0.011 (−0.17 to 0.19)</td>
</tr>
</tbody>
</table>

Study sample was drawn from the Panel Study of Income Dynamics for survey years 1985–2015. All models involved multivariable linear regressions, adjusting for age and age-squared, education, marital status, number of children, household inflation-adjusted pre-tax earned income and income-squared, year fixed effects, and individual fixed effects. Robust SEs were clustered at the individual level. Coefficients are presented per US$1000 of income or EITC.

Model 1 represents the OLS model, regressing each outcome on pre-tax inflation-adjusted household income.

Model 2 represents the reduced form of the IV analysis, regressing each outcome on EITC refund size.

Model 3 represents the IV analysis, with EITC refund size as the instrument for post-tax inflation-adjusted household income.

*p<0.05, **p<0.01.

EITC, earned income tax credit; IV, instrumental variable; OLS, ordinary least squares.
outcome, and in several cases (eg, alcohol consumption and smoking), the direction of the effect estimate was different from the primary models. Using federal EITC refund size as the instrument to overcome residual state-level confounding (table 3, Model 4), results were similar to primary models: higher income was associated with lower psychological stress (−0.13 per US$1000, 95% CI: −0.20 to −0.049).

**DISCUSSION**

This study examined the effects of income on mental health and health behaviours in a large national sample of US adults, leveraging variation in the EITC and quasi-experimental analyses to reduce confounding present in prior correlational studies.

OLS models suggested that income and the EITC are associated with reduced psychological distress, while income was also associated with improved self-reported health and increased alcohol consumption and cigarettes smoked per day. Prior studies have found that the EITC has beneficial effects on mental health, including depressive symptoms and suicide. Other work has also suggested increased substance use after receipt of government benefits, though recent studies found no short-term effects of the EITC in particular on smoking or alcohol consumption or reductions in smoking. There may be several explanations for these contrasting findings. For analyses of the EITC, the annual lump-sum EITC tax refund may be spent differently from income obtained more regularly through employment or other sources, such that its effects on overall health and health behaviours may operate through different pathways. Indeed, prior work suggests that the EITC is often spent on large appliances and reducing debt. Alternatively, it may be that the overall sample is not an appropriate control group for EITC recipients and that restricting to those who are more likely to be eligible for the EITC addresses an important source of bias in our primary analyses. Additionally, the finding of improved mental health particularly among single individuals reinforces that the benefit may be particularly salient for this vulnerable group.

Our results corroborate findings that income and the EITC are associated with improved mental health. The size of the effect estimate for psychological distress is about 2%-3% of a SD per US$1000, representing a meaningful change at the population level. Our results are consistent with two prior studies finding no effect on alcohol use and contradicts prior evidence that the EITC reduces smoking. Since our study uses more recent data, it may be that this association has changed over time as smoking rates in the US decline and the social safety net has changed. However, it may be that prior studies do not accurately capture the population of EITC recipients because of the use of crude proxies for EITC eligibility (eg, educational attainment), and our results represent a more accurate estimate of policy effects. Studies have shown that the links between socioeconomic status and health behaviours may depend on appropriate adjustment for confounders, suggesting that more rigorous analytical methods are needed. Future studies should attempt to replicate these findings in more recent cohorts and with different...
study designs, and results should be contrasted with studies that provide income in more frequent disbursements.

There are several limitations to this study. First, we assumed that all participants received EITC refunds if eligible, which is likely to introduce misclassification. Since national surveys rarely include questions on EITC receipt, this technique is commonly used in the literature, analogous to an intent-to-treat approach and an alternative to prior studies that use educational attainment as a proxy for EITC eligibility. Second, IV analyses assume that there is no other pathway linking the instrument and outcomes other than income; in this case, changes in employment as a result of the EITC may represent one such violation of this assumption, so results should be interpreted cautiously. Third, there may be measurement error as demographic and outcome data were self-reported. PSID includes a limited number of health outcomes, and therefore we were unable to examine other operationalisations of mental health such as depressive symptoms or self-sufficiency. The analysis may also suffer from time-varying confounding, if EITC refunds received in one wave affect EITC determinants (eg, marital status) in future waves; unfortunately, while we adjust for these measured characteristics in our models, IV analysis is not compatible with marginal structural models or other types of analyses that account for this type of confounding. Future studies should attempt to validate our findings using alternative modelling strategies. Finally, IV analysis represents a ‘local average treatment effect’, and results therefore may not generalise to income received through other means or across income ranges, for example, to higher-income or unemployed households.

Our study contributes to the literature on the health effects of income and provides evidence on the impact of one of the largest US poverty alleviation policies on health. The findings enforce our theoretical understanding of the effects of income on mental health and health behaviours and help inform policymaking on a population-level intervention that may reduce socioeconomic disparities in health. These findings also complement work showing that state EITC supplements are cost-effective and are associated with improvements in health-related quality of life and longevity, important insights for advancing policy discussions on the EITC.

Future research could examine other mental health outcomes or income supplementation programmes implemented differently (eg, the minimum wage or basic income not linked to employment).

Twitter Laura Shields-Zeeman @shieldsy_ and Rita Hamad @DrRitaHamad

Contributors RH conceived of the study, LS-Z, DC, and AB contributed to data cleaning and analysis. LS-Z created the first draft of the manuscript. All authors contributed to interpretation of the results, critically revised the manuscript and approved of the final version.

Funding This work was supported by grants from the UCSF Hellman Fellows Fund, the UCSF Irene Perrenstein Award and the UCSF National Centre of Excellence in Women’s Health. The collection of Panel Study of Income Dynamics data used in this study was partly supported by the National Institutes of Health (grant number R01HD069609 and R01AG040213), and the National Science Foundation (award numbers SES 1157698 and 1623684).

Disclaimer The study funders had no role in study design; collection, analysis and interpretation of data; writing the report; and the decision to submit the report for publication.

Competing interests None declared.

Patient consent for publication Not required.

Ethics approval Ethical approval was not required, as this study was conducted using public de-identified data.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Data are available in a public, open access repository. The source data for this study are publicly available from the website of the Panel Study of Income Dynamics, https://psidonline.isr.umich.edu.

Supplemental material This content has been supplied by the author(s). It has not been vetted by BMJ Publishing Group Limited (BMJ) and may not have been peer-reviewed. Any opinions or recommendations discussed are solely those of the author(s) and are not endorsed by BMJ. BMJ disclaims all liability and responsibility arising from any reliance placed on the content. Where the content includes any translated material, BMJ does not warrant the accuracy and reliability of the translations (including but not limited to local regulations, clinical guidelines, terminology, drug names and drug dosages), and is not responsible for any error and/or omissions arising from translation and adaptation or otherwise.

ORCID ID Rita Hamad http://orcid.org/0000-0003-0730-4077

REFERENCES


