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Psychological distress, neuroticism, and cause-specific mortality: early prospective evidence from UK Biobank

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► Additional material is published online only. To view please visit the journal online (<http://dx.doi.org/10.1136/jech-2016-207267>).

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Received 25 January 2016

Revised 16 June 2016

Accepted 3 July 2016

Published Online First

12 August 2016

ABSTRACT

Background It is well established that psychological distress (depression and anxiety) is related to an increased risk of mortality. The personality trait of neuroticism, reflecting a relatively stable tendency towards negative emotions, has also been associated with elevated rates of death in some studies. Accordingly, we tested the possibility that it is the neuroticism trait itself, rather than the distress state, that is generating an increased risk of mortality.

Methods We used data from the UK Biobank study, a UK-wide prospective cohort study (2006–2010) in which distress was ascertained using the Patient Health Questionnaire and neuroticism using the Eysenck Personality Questionnaire-Revised Short Form.

Results A mean of 6.2 years of follow-up of 308 721 study members gave rise to 4334 deaths. Higher neuroticism was weakly associated with total mortality (age-adjusted and sex-adjusted HR per SD increase; 95% CI 1.05; 1.02 to 1.09), and moderately strongly correlated with distress symptoms ($r=0.55$, $p<0.0001$). Distress symptoms were positively related to risk of total mortality (age-adjusted and sex-adjusted HR per SD increase in distress; 95% CI 1.23; 1.20 to 1.26). This gradient was, in fact, slightly strengthened after adding neuroticism to the multivariable model (1.30; 1.26 to 1.34) but markedly attenuated after taking into account other covariates which included health behaviours and somatic disease (1.16; 1.12 to 1.20). Similar results were apparent when cardiovascular disease, cancer and external cause of death were the end points of interest.

Conclusions While there was good a priori reasons to anticipate the neuroticism would at least partially explain the relation between distress symptoms and cause-specific mortality, we found no such evidence in the present study.

INTRODUCTION

Individual-participant and literature-based meta-analyses reveal dose-response relationships between higher levels of psychological distress (depression/anxiety) and the risk of premature mortality and selected chronic diseases.^{1–3} These observations have led to the speculation that treatment for distress could usefully occur in individuals at lower levels of distress than is currently recommended. The personality trait of neuroticism, reflecting a relatively stable tendency towards negative emotions, has also been associated with elevated rates of death and cardiovascular disease (CVD) in some studies.⁴ These inter-relationships raise the possibility that it is the

neuroticism trait itself, rather than the distress state, that is generating an elevated risk of mortality. With no empirical examination of this hypothesis, we assessed the impact of controlling for neuroticism on the distress–mortality relation alongside a series of more traditional explanatory variables, including health behaviours.

METHODS

UK Biobank, a UK-wide, on-going, prospective cohort study, has been described in detail.⁵ In brief, between 2006 and 2010, 502 649 participants aged 37–73 years attended various geographically distributed research clinics. Members of the public visited an assessment centre completed a questionnaire, underwent an interview and took part in various physical assessments. Ethical approval was obtained from the National Health Service National Research Ethics Service and all participants providing written informed consent.

Assessment of psychological distress and neuroticism

Psychological distress was measured using the four-item version of the Patient Health Questionnaire (PHQ-4).⁶ Items are rated on a four-point Likert scale from 0 (not at all) to 3 (nearly every day) such that possible total scores range from 0 to 12 (higher scores denote greater distress). Scores on the PHQ-4 show good agreement with longer scales, and correlate with demographic risk factors for depression and anxiety.⁷ Neuroticism was measured with the 12-item Eysenck Personality Questionnaire-Revised Short Form.⁸ Other covariate data were collected using standard protocols, including: health behaviours (smoking status, alcohol intake, physical activity, dietary characteristics), physical attributes (body mass index, systolic blood pressure, forced expiratory volume in 1 min, grip strength), existing disease (physician diagnoses of vascular or heart problems, diabetes, cancer, asthma, chronic lung disease, deep vein thrombosis or pulmonary embolism at baseline) and socioeconomic status (highest attained educational qualification).

Study participants were linked to the National Health Service's Central Registry at Southport, UK, which provided vital status data and, where applicable, cause of death. Having ascertained that the proportional hazards assumption had been met, we used Cox regression analyses with accompanying 95% CIs to summarise the association between psychological distress and mortality experience. In our

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To cite: Batty GD, McIntosh AM, Russ TC, et al. *J Epidemiol Community Health* 2016;**70**:1136–1139.

analyses, using the PHQ-4, psychological distress was categorised into three groups (score): 1 (0), 2 (1–2), 3 (≥ 3). The selection of these categories was data driven: we wanted sufficient numbers of deaths in each distress category to conduct robust statistical analyses.

RESULTS

In online supplementary table S1, we show the relation between categories of the distress score and study member characteristics at baseline. In general, people with a higher distress score had a less favourable risk factor profile as evidenced by poorer health behaviours, a higher prevalence of chronic disease and elevated neuroticism scores ($r=0.55$, $p<0.0001$). This was not a universal observation, however, in that levels of blood pressure and alcohol intake were somewhat *lower* in study members reporting a greater degree of distress.

A total of 308 721 people (142 983 women) had data on distress, neuroticism, other potential confounding variables and

mortality. During a mean follow-up period of 6.2 years, 4334 people died. Table 1 shows the relation of distress scores with total and cause-specific mortality. Irrespective of the mortality outcome, there was a positive distress–death relationship such that an elevated mortality risk was apparent in people with higher distress scores. Higher neuroticism was weakly associated with total mortality (age-adjusted and sex-adjusted HR per SD (2.92) increase; 95% CI 1.05; 1.02 to 1.09), however, in none of our analyses did controlling for neuroticism have an attenuating impact on the distress–mortality end point associations. Indeed, positive confounding, whereby the distress–disease relationship was strengthened, was generally apparent. In contrast, separate adjustment for each cluster of covariates led to partial attenuation, with the greatest impact apparent for health behaviours, irrespective of the outcome interest. Controlling simultaneously for an additional 15 covariates led to more marked attenuation of risk although a dose–response association remained, as we illustrate for total mortality in

Table 1 HRs (95% CI) for psychological distress in relation to cause-specific mortality: UK Biobank (N=308 721)

	Distress categories			p Value for trend	SD increase in distress
	1 (N=128 505)	2 (N=112 635)	3 (N=67 581)		
Total (4334 deaths)					
Age-adjusted and sex-adjusted	1.0 (ref)	1.22 (1.14 to 1.31)	1.77 (1.64 to 1.91)	<0.0001	1.23 (1.20 to 1.26)
Age-adjusted, sex-adjusted and neuroticism-adjusted	1.0	1.28 (1.19 to 1.38)	2.01 (1.84 to 2.20)	<0.0001	1.30 (1.26 to 1.34)
Age-adjusted, sex-adjusted and health behaviours-adjusted	1.0	1.16 (1.09 to 1.25)	1.49 (1.38 to 1.61)	<0.0001	1.14 (1.11 to 1.18)
Age-adjusted, sex-adjusted and physical attributes-adjusted	1.0	1.18 (1.10 to 1.27)	1.60 (1.48 to 1.72)	<0.0001	1.18 (1.15 to 1.21)
Age-adjusted, sex-adjusted and existing disease-adjusted	1.0	1.15 (1.07 to 1.23)	1.57 (1.45 to 1.69)	<0.0001	1.18 (1.15 to 1.21)
Age-adjusted, sex-adjusted and SES-adjusted	1.0	1.23 (1.15 to 1.32)	1.72 (1.59 to 1.85)	<0.0001	1.21 (1.18 to 1.24)
Multiply-adjusted	1.0	1.16 (1.08 to 1.25)	1.54 (1.40 to 1.68)	<0.0001	1.16 (1.12 to 1.20)
Cancer (2807 deaths)					
Age-adjusted and sex-adjusted	1.0 (ref)	1.19 (1.10 to 1.29)	1.59 (1.44 to 1.75)	<0.0001	1.17 (1.13 to 1.21)
Age-adjusted, sex-adjusted and neuroticism-adjusted	1.0	1.27 (1.16 to 1.36)	1.90 (1.69 to 2.12)	<0.0001	1.25 (1.20 to 1.31)
Age-adjusted, sex-adjusted and health behaviours-adjusted	1.0	1.14 (1.05 to 1.24)	1.38 (1.25 to 1.52)	<0.0001	1.10 (1.07 to 1.14)
Age-adjusted, sex-adjusted and physical attributes-adjusted	1.0	1.17 (1.07 to 1.27)	1.50 (1.36 to 1.65)	<0.0001	1.14 (1.10 to 1.19)
Age-adjusted, sex-adjusted and existing disease-adjusted	1.0	1.13 (1.04 to 1.23)	1.46 (1.32 to 1.61)	<0.0001	1.14 (1.10 to 1.18)
Age-adjusted, sex-adjusted and SES-adjusted	1.0	1.19 (1.09 to 1.30)	1.55 (1.41 to 1.70)	<0.0001	1.16 (1.12 to 1.20)
Multiply-adjusted	1.0	1.16 (1.07 to 1.28)	1.54 (1.37 to 1.72)	<0.0001	1.15 (1.10 to 1.20)
Cardiovascular disease (890 deaths)					
Age-adjusted and sex-adjusted	1.0 (ref)	1.25 (1.08 to 1.46)	1.85 (1.56 to 2.18)	<0.0001	1.27 (1.20 to 1.34)
Age-adjusted, sex-adjusted and neuroticism-adjusted	1.0	1.31 (1.12 to 1.54)	2.09 (1.71 to 2.55)	<0.0001	1.34 (1.25 to 1.34)
Age-adjusted, sex-adjusted and health behaviours-adjusted	1.0	1.17 (1.00 to 1.36)	1.40 (1.18 to 1.66)	<0.0001	1.13 (1.07 to 1.20)
Age-adjusted, sex-adjusted and physical attributes-adjusted	1.0	1.18 (1.01 to 1.37)	1.53 (1.29 to 1.82)	<0.0001	1.18 (1.11 to 1.25)
Age-adjusted, sex-adjusted and existing disease-adjusted	1.0	1.13 (0.97 to 1.32)	1.47 (1.24 to 1.75)	<0.0001	1.17 (1.10 to 1.24)
Age-adjusted, sex-adjusted and SES-adjusted	1.0	1.26 (1.08 to 1.47)	1.76 (1.48 to 2.08)	<0.0001	1.24 (1.17 to 1.31)
Multiply-adjusted	1.0	1.13 (0.96 to 1.32)	1.32 (1.07 to 1.65)	0.009	1.11 (1.03 to 1.20)
External (406 deaths)					
Age-adjusted and sex-adjusted	1.0 (ref)	1.17 (0.92 to 1.48)	2.06 (1.62 to 2.62)	<0.0001	1.31 (1.21 to 1.41)
Age-adjusted, sex-adjusted and neuroticism-adjusted	1.0	1.16 (0.91 to 1.48)	2.01 (1.51 to 2.69)	<0.0001	1.31 (1.20 to 1.34)
Age-adjusted, sex-adjusted and health behaviours-adjusted	1.0	1.12 (0.88 to 1.41)	1.73 (1.35 to 2.21)	<0.0001	1.22 (1.12 to 1.32)
Age-adjusted, sex-adjusted and physical attributes-adjusted	1.0	1.13 (0.89 to 1.43)	1.85 (1.45 to 2.37)	<0.0001	1.26 (1.16 to 1.36)
Age-adjusted, sex-adjusted and existing disease-adjusted	1.0	1.09 (0.86 to 1.38)	1.77 (1.39 to 2.27)	<0.0001	1.24 (1.15 to 1.35)
Age-adjusted, sex-adjusted and SES-adjusted	1.0	1.17 (0.93 to 1.49)	1.97 (1.56 to 2.51)	<0.0001	1.28 (1.18 to 1.38)
Multiply-adjusted	1.0	1.06 (0.83 to 1.36)	1.53 (1.14 to 2.06)	0.009	1.17 (1.06 to 1.29)

Multiple adjustment is adjustment for the following covariates: health behaviours (smoking status, alcohol intake frequency, number of types of physical activity performed in last month, whether five or more portions of fruit and vegetables eaten per day); physical attributes (body mass index, systolic blood pressure, forced expiratory volume in 1 min, grip strength); existing disease (diagnoses of vascular or heart problems, diabetes, cancer, asthma, chronic lung disease, deep vein thrombosis or pulmonary embolism at baseline) and socioeconomic status (highest educational qualification).

One SD in psychological distress is 2.0 units; a higher distress category denotes greater distress.

SES, socioeconomic status.

Figure 1 HRs for the full range of psychological distress scores in relation to total mortality: UK Biobank (N=308 721). PHQ, Patient Health Questionnaire.

figure 1. Finally, we disaggregated the PHQ-4 into its four component parts (depressed mood, unenthusiasm/disinterest, tenseness/restlessness and tiredness/lethargy) and related these to total and CVD mortality (see online supplementary table S2). The same pattern of association was evident as for the total distress score.

DISCUSSION

We observed a dose–response relationship between a measure of psychological distress symptom severity and death from diseases of public health importance in men and women in a large, contemporary and well-characterised cohort study. Neuroticism, contrary to our prediction, had no explanatory power in this association.

Study strengths and limitations

We excluded 140 352 participants from our analytical sample owing to missing data. Relative to the analytical sample, the excluded group had slightly higher scores for distress (mean 1.87 vs 1.49), and neuroticism (mean 3.78 vs 3.57), and higher mortality (1.51; 1.45 to 1.58). The distress–total mortality relation (age-adjusted and sex-adjusted HR per SD increase in distress; 95% CI) was also somewhat higher (p value for interaction 0.009) in the excluded group (1.31; 1.27 to 1.34) than in the analytical sample (1.23; 1.20 to 1.26), suggesting the HRs in our analytical sample may underestimate the true levels.

We assessed distress symptoms in the present study using a very brief inventory, the four-item PHQ, which was designed as a screening tool rather than a diagnostic device. Investigators on the largest other study to date to examine the distress–total mortality link administered a more detailed distress questionnaire,¹ yet the age-adjusted and sex-adjusted HR per SD increase in distress (1.21; 1.15 to 1.27) was almost identical to that found in the present study (1.23; 1.20 to 1.26). Also in accordance with that study, here we found that the weakest relationship was apparent for cancer and the strongest for external causes of death. Additionally, the correlates of scores from the PHQ-4 in the present study—gender, health behaviours, somatic illness—are similar to those for depression and anxiety, suggesting some concurrent validity. Taken together, that a single administration of a very brief inventory of distress revealed relatively strong effects may be testimony to the robustness of these relationships.

Inevitably, we have failed to capture all confounding factors, or assessed crudely some of those included here. Residual confounding is a perennial concern in observational epidemiology that should, in principle, be circumvented by using the randomised controlled trial design. While it would be unethical to precipitate prolonged bouts of distress and observe effects on mortality, an alternative approach is to reverse depression using a pharmacological and/or spoken therapy and evaluate the impact of mortality experience, anticipating a lower risk in the intervention arm. In one of the few such trials conducted, based on a population of cardiac patients, people successfully treated for depression did not experience a reduction in event-free survival relative to the usual care group.⁹ In a cluster randomised trial of older people recruited from primary care settings, however, improved survival was, however, apparent in the treatment arm.¹⁰

Finally, the response proportion in UK Biobank, at 10%, is very low by comparison with other studies. While this has impact on accurately estimating disease prevalence and incidence—any calculations are likely to be underestimates¹¹—it has little implications for understating the aetiological role of risk factors for a given chronic disease for which UK Biobank was established.¹² The original Whitehall study of London-based, non-industrialised civil servants (raised blood glucose as a risk factor for heart disease¹³), the British Doctors' study of registered physicians (cigarette smoking and selected cancers¹⁴) and Framingham study based in a single, affluent Massachusetts town (elevated serum cholesterol as a risk factor for heart disease¹⁵) have all yielded findings of major public health importance despite being obviously unrepresentative of the general population. If a study has a large enough sample and succeeds in capturing the range of values within the exposure of interest, the results should be transportable. This logic has recently been supported by comparing results from an occupationally based study of civil servants (the second Whitehall study) with those from the geographically diverse British Regional Heart Study where near identical HRs across an array of known risk factors for coronary heart disease were reported.¹¹ Similarly, in another cohort study, the relation of risk factor data collected at study induction to future CVD mortality was the same as that in a smaller, select group resurveyed 8 years later.¹⁶

Plausible mechanisms

That adjustment for an array of covariates, including neuroticism, did not eliminate the impact of distress on mortality risk inevitably raises speculation as to other mechanisms that may explain this relationship. Any such mechanisms are likely to be outcome-specific. Thus, bouts of acute anxiety may lead to acute coronary ischaemia that have been precipitated by coronary vasospasm¹⁷ and/or episodic elevations in blood pressure. Recurrent exposure to emotional disorder may also inhibit natural killer cell function which is implicated in immune system function,¹⁸ and therefore immunity-related cancers. The symptoms of fatigue, poor concentration and sleep disturbance, which characterise even moderately distressed individuals, may impact unfavourably on decision-making, risk perception, coordination and response time, so precipitating external causes of death such as accidents.¹⁹

In conclusion, while there was good a priori reasons to anticipate the neuroticism would at least partially explain the relation between distress symptoms and cause-specific mortality, we found no such evidence in the present study.

What is already known on this subject

- ▶ Meta-analyses of observational studies have shown that psychological distress (depression and anxiety) is related to an increased risk of total mortality and cardiovascular disease.
- ▶ These gradients seem to be robust to control for various confounding factors, including health behaviours and socioeconomic status.
- ▶ The personality trait of neuroticism, reflecting a relatively stable tendency towards negative emotions, has been associated with elevated rates of death and cardiovascular disease in some studies.

What this study adds

- ▶ For the first time, to the best of our knowledge, we tested the possibility that it is the neuroticism trait itself, rather than the distress state, that is generating an elevated risk of mortality.
- ▶ Using data from UK Biobank, psychological distress was, as expected, associated with a range of mortality end points.
- ▶ We found no evidence supporting an explanatory role for neuroticism; rather, the greatest attenuating effect across multiple mortality outcomes was apparent for health behaviours.

Acknowledgements This research has been conducted using the UK Biobank Resource. TCR, GDB, CRG and IJD are members of the University of Edinburgh Centre for Cognitive Ageing and Cognitive Epidemiology, funding for which comes from the UK Biotechnology and Biological Sciences Research Council and the UK Medical Research Council. IJD and AMM are supported by the Wellcome Trust.

Contributors CRG, AMM, GDB were involved in study concept and design; CRG was involved in acquisition and preparation of the data set (including mortality linkage); CRG was involved in statistical analysis; all authors were involved in interpretation of the data; GDB and CRG were involved in drafting of the manuscript; all authors were involved in critical revision of the manuscript for important intellectual content. CRG had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. All authors saw and agreed on the final manuscript as well as the decision to submit for publication.

Competing interests None declared.

Ethics approval Ethical approval was obtained from the National Health Service National Research Ethics Service.

Provenance and peer review Not commissioned; externally peer reviewed.

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