Mortality and temperature in Sofia and London

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Accepted for publication 15 January 2003

Mortality has long been known to be seasonal, and has been associated with the effects of both heat and cold. Although the seasonality of deaths is probably attributable to a combination of factors including exercise, diet, social interactions, pollution effects, etc, hot and cold outdoor temperatures have been associated with increased mortality independently of season. In multi-country studies, cold related mortality has been found to be greater in less affluent countries, and in those with warmer winters. A multi-city study in the United States found that estimated heat effects were greater in the cooler, northern states, and that cold effects were greater in the south.

Methods used to assess temperature related mortality have varied considerably, and so have the way in which results have been presented. In most countries, temperature and season are clearly related to each other, so that many studies have controlled for season, though others have not. Season has been controlled for in various ways, for example, by including “month” as a variable in the model, or by digital filtering. Longer term trends have been added, for example a simple linear function of date. More recently, smoothed functions of date, for example as produced by generalised additive models, have been used to capture both seasonal and longer term trends in a flexible way. These smooth functions may also partially or wholly capture variation attributable to seasonal epidemics such as influenza. Air pollution may play a confounding part in mortality-temperature relations, and some researchers have controlled for this, others not.

Unlike air pollution, with its presumed monotonic, if not linear, dose-response relation, both the extremes of temperature have adverse effects upon health. This causes complications in modelling, and some researchers have dealt with the problem by concentrating only upon cold effects or heat effects—sometimes but not always by splitting the year into cold and hot seasons—sometimes even by comparing total mortality in two places, one colder than the other. In other studies, heat and cold have been modelled simultaneously—graphically, using a smoothed function of temperature, produced from a generalised additive model, or as slopes, linear estimates of effects above and below some threshold. Where these slopes are found to differ, for example between countries, it may not be clear how well these differences reflect actual differences in the overall impact of temperatures on mortality, given that the effects may operate over different numbers of days in different countries, or that temperatures may be reach greater extremes in one than another.

Our aim in this small study was to explore, in as transparent a way as possible, associations between mortality and temperature in two very different European capitals—London, a northern maritime city, and Sofia, at the opposite end of the continent, considerably less affluent and with a more extreme climate. We assess the associations graphically, and also produce estimates of linear effects (for heat and cold), exploring the effects of lag and of differing choices of cut off points. We also attempt to assess and compare the overall impact of heat and cold on mortality, under some simple assumptions.

METHODS

Mortality and temperature data

Four years of daily mortality were collected, from the Office for National Statistics for London, 1993–1996, and from the Civil Register of the Administrative Service for Sofia, 1996–1999. Temperature data over the same periods were collected from The British Atmospheric Data Centre and the Bulgarian National Institute for Hydrology and Meteorology. Associations were examined between all cause mortality and both two day and two week average temperatures. For the two day mean, the average was taken of the current day’s mean temperature and the previous day’s mean temperature. Similarly for the two week mean—an average of mean temperatures from the current and the previous 13 days.
Potential confounders
Data were collected on daily relative humidity, and on particulate matter levels (black smoke for London, TSP for Sofia).

Analysis
All analyses were done in STATA v6, using generalised additive (Poisson) models (GAMs). These allow some variables to be entered as linear terms, and others to be entered as smoothed functions, whose smoothness is determined by the number of degrees of freedom (the more degrees of freedom, the more flexibility is allowed in the shape of the function). The convergence tolerances of the GAM function were set to a low value, of 10−6, with a limit of 1000 iterations. We controlled for day of week, public holiday, particulate matter (with a one day lag), relative humidity and season (a smoothed function of date with seven degrees of freedom per year). The residual from the previous day was entered to allow for temporal autocorrelation, under the transitional model paradigm. Vari- ances were multiplied by the overdispersion factor (a factor which, if greater than 1, indicates more variability than would be expected from a Poisson process). This overdispersion factor was itself estimated from the Poisson model.

(a) The shape of the mortality-temperature relation:
Smoothed mortality-temperature functions were estimated to assess visually the shape of the relations between mortality and temperature. This was done separately for the two day and the two week mean temperatures to describe prompt and possibly delayed effects respectively. As the temperature ranges differed between the two cities, the temperature terms were entered into the model with one degree of freedom allowed per five temperature degrees, rather than specifying a fixed number of degrees of freedom.

(b) Quantitative estimates, and lags, of temperature effects at upper and lower extremes:
In a second model, instead of using a smoothed mortality-temperature function, we used linear terms to represent “heat” and “cold”. Thus we obtained quantitative effect estimates of heat and cold that could be compared between cities. The “heat” term was the average, over two days, of the daily number of degrees above the 90th centile of the two day mean temperature. Similarly “cold” was the average, over two weeks, of the daily number of degrees below the 10th centile of the two week mean. By using these centile cut off points, instead of defining the same absolute cut off points in both cities, we attempt to create relative definitions of cold and heat. This allows for possible acclimatisation, for example, by which a temperature considered low in one climate may be considered moderate in a colder climate. The “heat” and “cold” terms were present simultaneously in the model. Differences in estimates between the two cities were tested for significance using z tests. We also counted the number of days and hot “day degrees”—the total numbers of degrees above the 10th centile of the two day mean temperature, appears stronger in Sofia than in London.

(c) A more absolute, and more inclusive, definition of “cold” and “heat”
The choice of 10th and 90th temperature centiles as cut off points for “cold” and “heat” represented a relative definition of heat and cold with regard to the different climates of the two cities. A second definition was used, as a contrast with the first. Here we found a single cut off point, the same in both cities, thus imposing a “V” shaped relation across the whole temperature range. The cut off point was found by trying all integer possibilities from 15°C to 25°C in both cities and identifying the one which maximised the joint likelihood. As with the centile model, the “heat” and “cold” terms were the averages over two days and two weeks respectively of the daily number of degrees above and below the cut off point. Analyses of linear associations and of lags were done, using this V model.

(d) Attributable fractions
For each set of cut off points above, the fractions of mortality attributable to heat and to cold were estimated as the mean of the fractions (RR−1)/RR over deaths, where RR is the relative risk due to heat(cold) on the day of death according to the logistic regression model.

(e) The role of particulate matter in the model
Heat and cold effects (using the centiles as cut off points) were re-estimated without particulate matter in the model, to see whether it had a confounding effect on the mortality-temperature relations.

RESULTS
Temperatures ranged from −6°C (the minimum of daily minimums) to 34°C (the maximum of daily maximums) in London, and from −17°C to 37°C in Sofia. Mean winter temperature (October to March) was 8.0°C in London and 3.9°C in Sofia, summer means 17.0°C and 17.4°C respectively.

(a) Shape of mortality-temperature relation:
Figure 1 shows fitted mortality-temperature curves, for two day mean temperature and two week mean temperature. The curves show, for values across the temperature range, the fitted number of deaths (and 95% confidence intervals) re-scaled as a percentage of the mean number of daily deaths for that city. For reasons that will become clear below (in the results about lags of effect) the two day mean better captures heat effects and the two week mean better captures cold effects. This is apparent from figure 1.

The heat effect, as depicted in the graphs for the two day mean temperature, appears stronger in Sofia than in London. Both cities show a rise in mortality as the two week mean temperature drops, and this extends—most strikingly in Sofia—over much of the temperature range. Wider confidence intervals in Sofia reflect lower numbers of deaths—an average of 31 per day, compared with 176 in London.

(b) Quantitative estimates, and lags, of temperature effects:
The 10th and 90th centiles of the two week and two day mean temperatures respectively were used as cut off points for cold and heat effects. These gave cold cut off points of 5.25°C in London and −0.46°C in Sofia. The heat cut off points were similar in the two cities: 21°C and 21.55°C respectively. All the estimated linear effects were statistically significant, and differed noticeably between the two cities. For a one degree average drop below the 10th temperature centile over two weeks, mortality in London increased by 4.24% (95% CI

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3.41 to 5.07) after controlling for season and other potential confounders (table 1). In Sofia, mortality increased by a smaller amount, 1.83% (0.59 to 3.90). The total number of cold day degrees, according to this definition of “cold”, was larger in Sofia than in London (492 vs 426).

Conversely, estimated heat coefficients were larger in Sofia than London. For a one degree average rise above the 90th temperature centile over two days, mortality in London increased by 1.86% (1.36 to 2.36), and in Sofia by 3.49% (2.23 to 4.76). The total number of hot day degrees was higher in London.

Before controlling for season and the other potential confounders, apparent heat effects in both cities were weaker—in fact, were inverse—and not statistically significant (table 1). The opposite was the case for the associations seen between cold and mortality—associations appeared stronger before controlling for potential confounders. The degree of confounding appeared to be greater in Sofia than in London.

Figure 2 shows the estimated cold and heat effects, all mutually adjusted, across a range of lags. In Sofia, there was a strong heat effect on the current day, representing an increase in mortality of around 3.5% for an average one degree rise above the 90th centile, after controlling for potential confounders and the effects of previous days’ hot (or cold) weather. Between a lag of five days, and a lag centred around 22 days, estimated heat coefficients were all considerably below 1—that is, apparently beneficial. In London, a smaller heat effect of around 1.5% operated on the current day. Heat coefficients in London then dropped towards and then slightly below 1, where they remained.

Cold effects in London operated over a considerable period, most strongly after a lag of three days, and were still very evident at a lag of 22 days. The individual cold coefficients in Sofia, though also above 1 for much of this same period, were smaller than in London, and not individually significant, and dropped slightly below 1 at a lag of 22 days. In Sofia there was an apparent beneficial effect of cold on the current day, after allowing for effects of cold (and/or heat) on previous days.

Table 1  Total numbers of degrees above (below) cut off point, on days defined as hot (cold) using the given cut off points

<table>
<thead>
<tr>
<th>Sofia</th>
<th>London</th>
</tr>
</thead>
<tbody>
<tr>
<td>Day degrees</td>
<td>% increase in mortality per °C (95% CI)</td>
</tr>
<tr>
<td>Heat—unadjusted</td>
<td>347</td>
</tr>
<tr>
<td>Heat—adjusted*</td>
<td>347</td>
</tr>
<tr>
<td>Cold—unadjusted</td>
<td>492</td>
</tr>
<tr>
<td>Cold—adjusted*</td>
<td>492</td>
</tr>
</tbody>
</table>

Using the V model with common cut off point in both cities:
| Heat—unadjusted | 1282 | 0.61 (-0.05 to 1.27) | 1044 | 0.91 (0.58 to 1.23) |
| Heat—adjusted* | 1282 | 2.21 (1.55 to 2.87) | 1044 | 1.30 (0.99 to 1.62) |
| Cold—unadjusted | 11776 | 1.16 (0.98 to 1.34) | 8892 | 1.55 (1.39 to 1.71) |
| Cold—adjusted* | 11776 | 0.70 (0.51 to 0.88) | 8892 | 1.43 (1.28 to 1.58) |

Estimated percentage rise (95% CI) in mortality for a daily average of one degree over two days (two weeks) above (below) the cut off point. Heat and cold effects are mutually adjusted. Adjusted models (*) also contains terms for season and long term trends, day of week, public holiday, relative humidity, particulate matter.

Figure 1  Fitted deaths [scaled to be a percentage of mean daily deaths] in Sofia (left) and London. Plotted against temperature, two day mean (top) and two week mean.
(c) Quantitative estimates assuming a V shaped mortality-temperature relation with common cut off point

The cut off point that maximised the joint likelihood was at 18°C. The V shaped relation meant that heat or cold effects operated on all days. In particular, many more days contributed to the “cold” effects, and per degree estimates of both heat and cold coefficients were correspondingly lower. The increase in mortality in London, for a one degree rise above 18°C, was 1.30% (0.99 to 1.62), after adjusting for season and the other confounders (table 1). The increase in mortality for each degree fall below 18°C was 1.43% (1.28 to 1.58)—much smaller than the estimated linear effect below 5.25°C, the 10th centile. In Sofia, the increase in mortality for a one degree rise above the cut off point was 2.21% (1.55 to 2.87), and for a one degree fall below the cut off point was 0.70 (0.51 to 0.88).

Although coefficients were smaller than seen in the previous model, in overall pattern the results were similar. Patterns seen for cold and heat effects across a range of lags, using this V shape model, were essentially the same as those seen previously. The greatest difference was that heat effects in London did not remain below zero at greater lags, but dropped to and stayed around zero. In Sofia, as before, heat effects beyond a lag of three days were all below zero.

(d) Crude estimates of the attributable fractions of mortality attributable to heat and cold

Around 1.4% of deaths in London were estimated as being attributable to cold (table 2), using the centile cut off points, and 0.67% of deaths in Sofia. The numbers attributable to heat were far fewer than for cold in London, 0.44%, but 0.72% of deaths were attributable to heat in Sofia—similar to the fraction for cold.

The sizes of linear coefficients estimated by the V shaped model were smaller, but all days contributed to heat or cold effects, not just the extremes, and the estimated attributable fractions of deaths were larger. This was particularly so for cold, with 5.75% of deaths attributable to cold in Sofia, and 8.67% in London. Attributable fractions for heat were 1.66% and 0.82% respectively.

(e) The role of particulate matter in the model

The estimated PM effect in Sofia was highly significant, after controlling for temperature, season and the other potential confounders: an increase of 0.06% (0.03 to 0.10) in mortality per 1 µg/m³ increase in TSP. The estimated increase in mortality in London per unit of black smoke was of a similar order, but of no statistical significance: 0.08% (−0.67 to 0.84).

Excluding particulate matter from the model increased estimates of both heat and cold coefficients in both cities, but only very slightly. The largest difference was in the cold coefficient for Sofia, which was 1.83% (0.59 to 3.90) with PM in the model (table 1), and 2.18% (95% CI 0.98 to 3.40) without it.

### Table 2

Percentage (95% CI) of deaths attributable to cold and heat, estimated using different cut off points

<table>
<thead>
<tr>
<th></th>
<th>Sofia</th>
<th>London</th>
</tr>
</thead>
<tbody>
<tr>
<td>Heat</td>
<td>0.72 (0.47 to 0.96)</td>
<td>0.44 (0.33 to 0.56)</td>
</tr>
<tr>
<td>Cold</td>
<td>0.67 (0.22 to 1.11)</td>
<td>1.39 (1.13 to 1.64)</td>
</tr>
<tr>
<td>Heat</td>
<td>1.66 (1.19 to 2.13)</td>
<td>0.82 (0.63 to 1.01)</td>
</tr>
<tr>
<td>Cold</td>
<td>5.75 (4.29 to 7.18)</td>
<td>8.67 (7.82 to 9.52)</td>
</tr>
</tbody>
</table>

DISCUSSION

For two capital cities of Europe, Sofia and London lie far apart both geographically and economically. Their climates are

**Key points**

- Cold and heat are both strongly associated with mortality in both cities
- Stronger heat effects in Sofia may be partly attributable to “harvesting”
- Winter is harmful, and summer beneficial, independent of temperature
we cannot rule confounding out, we have attempted to
adjust for previous days’ cold (and hot) tempera-
ture, and two week averages to represent
inverse effects of cold were seen in Sofia after about three
weeks, and of heat after about four days—these may indicate
harvesting, the bringing forward of imminent death. Braga
remarked that, from their findings, “the hot temperature
effect appears to be primarily harvesting”. Inverse heat
effects at later lags were less evident in London, and not at all
for cold.
Heat coefficients were greater in the full than in the unad-
justed model, whereas cold coefficients were reduced by
controlling for confounders. These changes were largely
attributable to season. Adjusting, singly, for the other
covariates made little difference (results not shown). This
controlling in opposite directions may indicate harmful
winter effects and beneficial summer effects, both acting
independently of temperature, for example, because of differ-
ces in diet, sunlight, etc.
There were no major influenza epidemics in either city dur-
ing the study period. However, we checked, using London
influenza notifications, the impact of controlling for flu
(results not shown above). Influenza was an independent pre-
dictor of mortality, but heat and cold coefficients were not sig-
nificantly changed by including it in the model. We feel it
unlikely that including flu would have changed our conclu-
sions. Nor did measures of particularity matter greatly
confound any mortality-temperature associations.
Possible permutations of temperature-mortality models
seem almost endless, and the biological and epidemiological
evidence upon which to base the choices is scarce. We have
attempted to explore a few simple variations and hence to
reflect upon our way forward, drawing on current time series
methodologies developed by Schwartz, Dominici, and
others. Among quantitative estimates of effect, linear tempera-
ture-mortality slopes seem the type most often presented in previ-
ous studies. These may not fully compare the impacts of tem-
perature upon mortality. For example, city A may have a
steeper cold-mortality slope than city B, but if temperatures
frequently drop lower in city B, the overall impact of cold may
be greater in city B. We therefore also presented fractions of
mortality attributable to cold and heat, under assumptions of
linearity and of the choice of cut off points. The choice of cut
off point changes both types of estimate. In the event, the pat-
terms of higher cold effects in London, and higher heat effects
in Sofia, seemed robust both to the choice of cut off points and
to the type of estimate, but this may not always be the case.
We used two a priori definitions of cut off points, intended
to provide direct comparisons between the cities—one of
effects of absolute temperature, one of relative temperature.
The cut off point used for the V shaped model was chosen
empirically (using likelihood methods). This has been done
before, sometimes allowing different cut off points for each
city.21 The last idea is attractive—and used extensively by
estimates for modellers, say—though makes resulting esti-
mates less comparable between cities, as the selected cut off
points may vary in both absolute and relative terms. The V
shape model, though imposing some possibly undesirable
constraints, removes other difficulties from analysis and
interpretation associated with a two cut off point model—for
example, finding that the empirically derived “hot” cut off
point is below the “cold” cut off point—and may give good
results for the purposes of comparison.
Using the V shaped model allows all days to contribute to the
temperature-mortality effects, and estimates of attribut-
able fractions of deaths are correspondingly higher, with over
8% of deaths in London attributable to cold. These high figures
are compatible with attributable fractions published from pre-
vious studies.21 22 23 Keatinge and colleagues, albeit not control-
ling for season, show London/UK attributable mortality corre-
sponding to figures of around 5% to well over 10%. Although
we cannot rule confounding out, we have attempted to
account for obvious factors, and the smoothed GAM function
of time also helps to account for unmeasured effects. It should
also be mentioned that cold is thought mainly to precipitate
deaths from other causes—usually cardiovascular disease—
rather than cold itself being thought of as the sole cause of all
these excess deaths.24 25
The V shaped model and the centile model estimate overall
and extreme effects of temperature—and, in our use of them
here—give an absolute and a relative comparison. The type of
estimate required would inform the choice of model. Actual fit
of model would not have helped us choose between them in
this instance, though it might do elsewhere. Comparing devi-
ances, the V shaped model fitted the Sofia data better and the
centile model fitted the London data better. Looking at the
GAM graphs, a V shaped model with a cut off point at roughly
18°C seems not a bad approximation. Separate maximum
likelihood estimates of V cut off points for Sofia and London
were 17°C (95% profile interval 15 to 20) and 21°C (18 to 24)
respectively.
Because of the lags of effects, we used two day temperature
averages to represent heat, and two week averages to represent
cold. If there is no mortality-heat association beyond a few
days’ lag, then averaging over a longer period should not bias
the result but is likely to reduce power. If a considerable part
of heat related mortality is actually the short-term bringing
forward of imminent death, then averaging over a longer
period may really be more appropriate, and results using two
day means may be overestimates. This seems plausible from
figure 2. Conversely, for cold, averaging over a shorter time
could underestimate and/or reduce power to detect effects.
Inverse mortality-cold associations seen on the current day,
after controlling for previous days’ cold (and hot) tempera-
tures, are hard to interpret. Keatinge suggests vasoconstric-
tion, delaying deaths among the very ill, and assessment of
this hypothesis would be useful.26
We consider it preferable to look at heat and cold effects
together. Fities we found “cold” days during the sum-
mer months, even with our more extreme definition of cold,
and a few “hot” days in winter. To estimate the effects of one,
Heat and cold have both been found to have harmful effects in both cities. Heat effects were worse in Sofia, though this may be ignoring some harvesting. Cold effects were worse in London. If inverse associations with cold at later lags are signs of harvesting, then the fact that this was seen in Sofia and not London suggest that cold temperatures in London harm the general population and the effects are not concentrated among persons close to death.

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REFERENCES
