Lymphatic and haematopoietic cancer mortality in a population attending school adjacent to styrene-butadiene facilities, 1963–1993

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Abstract

Study objective—To evaluate the risk of mortality from lymphatic and haematopoietic cancers and other causes among students.

Design—The study used school records, yearbooks, and Texas Department of Health records for the school years 1963–64 to 1992–93 to construct a cohort of 15 403 students. Three mortality databases were searched to identify deaths, and mortality rates in the cohort were compared with mortality rates from the United States and Texas. Computed standardised mortality ratios and 95% confidence intervals were used.

Setting—Eastern Texas high school adjacent to facilities that have been producing synthetic styrene-butadiene since 1943.

Main results—338 deaths were identified. The all causes standardised mortality ratio was 0.84 (95% confidence intervals 0.74, 0.95) for men and 0.89 (0.73, 1.09) for women. The standardised mortality ratio for all lymphatic and haematopoietic cancers was 1.64 (95% confidence intervals 0.85, 2.87) for men and 0.47 (0.06, 1.70) for women. The slight male excess in lymphatic and haematopoietic cancers was stronger among men who attended school for two years or less.

Conclusions—The overall mortality from lymphatic and haematopoietic cancer among the students was little different from that of the United States as a whole. A moderate excess for men, predominantly among the shorter-term students, was offset by a deficit among women. These variations are compatible with random fluctuations; the overall pattern is not indicative of an effect of environmental exposure sustained while attending the high school.

(J Epidemiol Community Health 1999;53:283–287)
Table 1  Selected SMRs and 95% confidence intervals by sex, compared with US standard rates

<table>
<thead>
<tr>
<th>Cause of death</th>
<th>Men</th>
<th></th>
<th></th>
<th></th>
<th>Women</th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Obs</td>
<td>Exp</td>
<td>SMR</td>
<td>95% CI</td>
<td>Obs</td>
<td>Exp</td>
<td>SMR</td>
<td>95% CI</td>
</tr>
<tr>
<td>All causes</td>
<td>241</td>
<td>287.72</td>
<td>0.84</td>
<td>0.74, 0.95</td>
<td>97</td>
<td>108.51</td>
<td>0.89</td>
<td>0.73, 1.09</td>
</tr>
<tr>
<td>All malignant neoplasms</td>
<td>31</td>
<td>25.50</td>
<td>1.22</td>
<td>0.83, 1.73</td>
<td>13</td>
<td>25.18</td>
<td>0.52</td>
<td>0.28, 0.88</td>
</tr>
<tr>
<td>Ca digestive organs, peritoneum</td>
<td>5</td>
<td>3.96</td>
<td>1.26</td>
<td>0.41, 2.94</td>
<td>2</td>
<td>2.52</td>
<td>0.80</td>
<td>0.30, 2.87</td>
</tr>
<tr>
<td>Ca bronchus, trachea, lung</td>
<td>5</td>
<td>3.43</td>
<td>1.46</td>
<td>0.47, 3.40</td>
<td>1</td>
<td>2.25</td>
<td>0.44</td>
<td>0.01, 2.48</td>
</tr>
<tr>
<td>Ca breast</td>
<td>0.00</td>
<td>0.02</td>
<td>0.00</td>
<td>0.01, 1.64</td>
<td>4</td>
<td>6.81</td>
<td>0.59</td>
<td>0.16, 1.50</td>
</tr>
<tr>
<td>Ca all lymphatic, haematopoietic</td>
<td>12</td>
<td>7.30</td>
<td>1.64</td>
<td>0.85, 2.87</td>
<td>2</td>
<td>4.24</td>
<td>0.47</td>
<td>0.06, 1.70</td>
</tr>
<tr>
<td>Hodgkin's disease</td>
<td>2</td>
<td>1.37</td>
<td>1.46</td>
<td>0.18, 5.28</td>
<td>1</td>
<td>0.83</td>
<td>1.20</td>
<td>0.03, 6.68</td>
</tr>
<tr>
<td>Leukaemia, aleukaemia</td>
<td>6</td>
<td>3.30</td>
<td>1.82</td>
<td>0.67, 3.96</td>
<td>1</td>
<td>2.24</td>
<td>0.45</td>
<td>0.01, 2.48</td>
</tr>
<tr>
<td>Ca all other lymphopoietic</td>
<td>4</td>
<td>1.95</td>
<td>2.05</td>
<td>0.11, 3.51</td>
<td>0</td>
<td>0.96</td>
<td>0.56</td>
<td>0.02, 1.57</td>
</tr>
<tr>
<td>All other malignant neoplasms</td>
<td>3</td>
<td>2.65</td>
<td>1.13</td>
<td>0.23, 3.31</td>
<td>0</td>
<td>1.78</td>
<td>0.00</td>
<td>0.14, 1.78</td>
</tr>
<tr>
<td>Benign neoplasms</td>
<td>5</td>
<td>0.80</td>
<td>6.27</td>
<td>2.04, 14.63</td>
<td>1</td>
<td>0.64</td>
<td>1.56</td>
<td>0.04, 8.71</td>
</tr>
<tr>
<td>All heart disease</td>
<td>13</td>
<td>22.72</td>
<td>0.57</td>
<td>0.31, 0.98</td>
<td>6</td>
<td>7.40</td>
<td>0.81</td>
<td>0.39, 1.78</td>
</tr>
<tr>
<td>Ischaemic heart disease</td>
<td>7</td>
<td>13.06</td>
<td>0.54</td>
<td>0.22, 1.11</td>
<td>3</td>
<td>2.80</td>
<td>1.07</td>
<td>0.22, 3.13</td>
</tr>
<tr>
<td>All other heart disease</td>
<td>6</td>
<td>7.63</td>
<td>0.79</td>
<td>0.29, 1.71</td>
<td>3</td>
<td>3.46</td>
<td>0.87</td>
<td>0.18, 2.53</td>
</tr>
<tr>
<td>All external causes of death</td>
<td>147</td>
<td>172.22</td>
<td>0.85</td>
<td>0.72, 1.00</td>
<td>56</td>
<td>45.80</td>
<td>1.22</td>
<td>0.92, 1.59</td>
</tr>
<tr>
<td>Accidents</td>
<td>104</td>
<td>112.40</td>
<td>0.93</td>
<td>0.76, 1.12</td>
<td>36</td>
<td>29.88</td>
<td>1.21</td>
<td>0.84, 1.67</td>
</tr>
<tr>
<td>Motor vehicle accidents</td>
<td>79</td>
<td>73.67</td>
<td>1.07</td>
<td>0.85, 1.34</td>
<td>30</td>
<td>22.86</td>
<td>1.31</td>
<td>0.89, 1.87</td>
</tr>
<tr>
<td>All other accidents</td>
<td>25</td>
<td>38.74</td>
<td>0.65</td>
<td>0.42, 0.95</td>
<td>6</td>
<td>7.02</td>
<td>0.85</td>
<td>0.31, 1.86</td>
</tr>
<tr>
<td>Suicides</td>
<td>26</td>
<td>35.81</td>
<td>0.73</td>
<td>0.47, 1.06</td>
<td>10</td>
<td>8.94</td>
<td>1.12</td>
<td>0.54, 2.06</td>
</tr>
<tr>
<td>Homicides</td>
<td>17</td>
<td>24.01</td>
<td>0.71</td>
<td>0.41, 1.13</td>
<td>10</td>
<td>6.98</td>
<td>1.43</td>
<td>0.69, 2.64</td>
</tr>
<tr>
<td>AIDS</td>
<td>28</td>
<td>20.35</td>
<td>1.38</td>
<td>0.91, 1.99</td>
<td>0</td>
<td>1.48</td>
<td>0.00</td>
<td>0.25, 2.38</td>
</tr>
</tbody>
</table>

To obtain data on sex and name changes that occurred after leaving high school, we accessed several ancillary data sources. Computerised birth records for the years 1942–1980 maintained by the Texas Department of Health provided sex for 73% of the students in our cohort. For the students for whom we did not find a match in the birth records, we had two independent assessors assign sex based on the student’s name and yearbook pictures. We were unable to assign sex to three students who did not have a picture in any yearbook and had sexual neutral names.

To identify name changes, we purchased a copy of the Texas Department of Health marriage database for the years 1966 (the first year available) to 1994. Our intent was to ascertain possible married names for female students to improve our ability to identify deaths that might have occurred later in life. We searched the marriage database using the names of the women in the eligible cohort who had complete dates of birth and the three people with unknown sex. Our matching algorithm identified one or more marriages and the resulting married names for 5154 (68%) of the women searched. We then conducted a second round of matching in the marriage database using as a starting point the married names picked up in the first round of matching. We found 838 additional matches. Upon completion of the school record review and the linking with the Texas marriage database, we had a total of 15 722 possible last names for the 7591 women in the eligible cohort.

We used three mortality databases for vital status searching: the National Death Index, the Social Security Administration Death Master Files, and the Texas Department of Health death database. The National Death Index included nearly all United States deaths from 1979 to 1994; the Social Security Administration files included United States deaths reported to the Administration from the early 1930s to 1995; the Texas death database included reported Texas deaths for the years 1964 to 1978. We submitted a total of 23 366 records for mortality searching, which were all of the possible names for the 15 408 unique cohort members who had a complete date of birth. We reviewed the results of the mortality searches and requested copies of death certificates for possible matches from the appropriate states. To be considered a correct match, the name (maiden or married name), date of birth and state of birth on the death certificate had to match the information in the cohort database. In some cases a comparison of the parents’ names was required to confirm the death match for a woman who had married more than once or may have been divorced. A certified nosologist coded all causes of death and underlying cause of death according to the version of the International Classification of Diseases (ICD)9 in effect at the time of death.

From the original cohort of 15 553 students we eliminated 145 with unknown date of birth, two with incongruous dates of birth, and three with unknown sex. A total of 15 403 students remained in the final analytic cohort.

For comparison we used United States and Texas mortality rates derived from United States National Center for Health Statistics data. The rates were stratified by sex, five year age categories, and five year calendar intervals from 1950–1992. We applied the 1990–92 rates to cohort members followed up in 1993. As the student body at the high school has been historically almost exclusively white, we used mortality rates for whites. We used the Occupational Mortality Analysis Program (OCMAP)9 to summarise person time and calculate standardised mortality ratios (SMRs) for all causes of death as well as for selected categories of death, including the lymphatic and haematopoietic cancers. We conducted all analyses using the underlying cause of death from the death certificate.

Results
We identified 338 graduates (241 men, 97 women) who had died. The 7882 men and 7521 women in the analytic cohort contributed a total of 310 254 person years during the follow up period of 1963 to 31 December 1995. Table 1 shows the SMRs for all causes and for
Mortality from environmental styrene-butadiene exposure

compared with US standard rates, men only
6.27 (95% CI 2.04, 14.63) for men and 1.56
deaths from benign neoplasms; the SMR was
0.66, 2.04). There was an unexpected excess of
sexes combined, the SMR was 1.21 (95% CI
0.83, 2.87). For women the corresponding
SMR was 0.47 (95% CI 0.06, 1.70). For both
causes of death among men who attended the
school for \(< 2\) years was 1.42 (95% CI 1.11,
1.80), while that for men who remained in
school for \(\geq 3\) years was 0.72 (95% CI 0.61,
0.83).

Among men who attended the high school for
\(< 2\) years, the greatest relative increase in
mortality from lymphatic and haematopoietic
cancers was for those dying of leukaemia and
aleukaemia. This category includes lymphoid
leukaemia; myeloid leukaemia; monocytic leu-
kaemia; other specified leukaemias such as
acute erythraemia, erythroleukaemia, and
megakaryocytic leukaemia; and leukaemias of
unspecified cell type. The corresponding ICD
codes are in the range 204.0 to 208.9.

Although there are some slight increases in risk
seen among other groups and for other catego-
ries of lymphatic and haematopoietic cancers,
the numbers are too small and the correspond-
ing confidence intervals too wide to draw con-
clusions regarding these subcategories. In gen-
eral, men who remained in school longer faced
smaller relative increases in mortality from
lymphatic and haematopoietic cancers. We
found a similar pattern for deaths resulting
from external causes. Among men who at-
tended the school for \(< 2\) years a large
percentage of cells was attributable to external
causes. For men 147 (61%) of 241 deaths were
caused by external causes and for women 55
(57%) of 97 deaths. The proportion of deaths
caused by external causes, compared with
organic illness, decreased with years since leav-
ing school (fig 1).

The SMRs for all cancers compared with the
US population were, for men, 1.22 (95% CI
0.83, 1.73) and for women 0.52 (95% CI 0.28,
0.88). The SMRs for individual cancer sites in
table 1 vary considerably; the number of
observed cases is small for individual cancer
sites, and the associated confidence intervals
are wide. The SMR for men for all lymphatic
and haematopoietic cancers was 1.64 (95% CI
0.85, 2.87). For women the corresponding
SMR was 0.47 (95% CI 0.06, 1.70). For both
sexes combined, the SMR was 1.21 (95% CI
0.66, 2.04). There was an unexpected excess of
deaths from benign neoplasms; the SMR was
6.27 (95% CI 2.04, 14.63) for men and 1.56
(95% CI 0.04, 8.71) for women. For both sexes
combined, the SMR for benign neoplasms was
4.18 (95% CI 1.53, 9.09).

We conducted additional analyses by sub-
group according to the number of years of
attendance at the high school (\(< 2\) years at the
school and \(\geq 3\) years, table 2). The SMR for all
causes of death among men who attended the
school for \(< 2\) years was 1.42 (95% CI 1.11,
1.80), while that for men who remained in
school for \(\geq 3\) years was 0.72 (95% CI 0.61,
0.83).

Table 2  Selected SMRs and 95% confidence intervals by years of attendance at Port Neches-Groves High School, compared with US standard rates, men only

<table>
<thead>
<tr>
<th>Cause of death</th>
<th>(&lt; 2) years (n=1538)</th>
<th>(\geq 3) years (n=6352)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Obs/Exp SMR 95% CI</td>
<td>Obs/Exp SMR 95% CI</td>
</tr>
<tr>
<td>All causes</td>
<td>71/49.89 1.42 (1.11,1.80)</td>
<td>170/237.83 0.72 (0.61,0.83)</td>
</tr>
<tr>
<td>All cancer</td>
<td>9/4.27 2.11 (0.96,4.00)</td>
<td>22/21.24 1.04 (0.65,1.57)</td>
</tr>
<tr>
<td>Ca all lymphatic, haematopoietic</td>
<td>4/1.25 3.20 (0.87,8.20)</td>
<td>8/6.05 1.32 (0.57,2.60)</td>
</tr>
<tr>
<td>Lymphosarcoma, reticulocarcinoma</td>
<td>0/0.12 0.00 (0.00,31.89)</td>
<td>0/0.07 0.00 (0.00,6.43)</td>
</tr>
<tr>
<td>Hodgkin’s disease</td>
<td>0/0.23 0.00 (0.00,15.75)</td>
<td>0/2.13 1.77 (0.21,6.38)</td>
</tr>
<tr>
<td>Leukaemia, aleukaemia</td>
<td>3/0.57 5.29 (1.09,15.46)</td>
<td>3/2.73 1.10 (0.23,3.21)</td>
</tr>
<tr>
<td>Ca all other lymphopoietic</td>
<td>1/0.31 3.01 (0.08,16.76)</td>
<td>3/1.62 1.86 (0.38,5.14)</td>
</tr>
<tr>
<td>Benign neoplasms</td>
<td>1/0.14 7.35 (0.18,40.94)</td>
<td>4/0.66 6.05 (0.65,15.48)</td>
</tr>
<tr>
<td>All heart disease</td>
<td>4/3.73 1.07 (0.29,2.75)</td>
<td>9/18.99 0.47 (0.22,0.90)</td>
</tr>
<tr>
<td>All external causes of death</td>
<td>48/30.35 1.58 (1.17,2.10)</td>
<td>99/141.88 0.70 (0.57,0.85)</td>
</tr>
<tr>
<td>Accidents</td>
<td>3/19.66 1.73 (1.20,4.42)</td>
<td>70/92.74 0.76 (0.59,0.95)</td>
</tr>
<tr>
<td>Motor vehicle accidents</td>
<td>25/12.95 1.93 (1.25,2.85)</td>
<td>54/60.72 0.89 (0.67,1.16)</td>
</tr>
<tr>
<td>All other accidents</td>
<td>9/4.72 2.34 (0.61,2.54)</td>
<td>16/32.02 0.50 (0.29,0.81)</td>
</tr>
<tr>
<td>Suicides</td>
<td>5/9.39 1.41 (0.64,2.68)</td>
<td>17/29.42 0.58 (0.34,0.93)</td>
</tr>
<tr>
<td>Homicides</td>
<td>5/4.30 1.16 (0.38,2.72)</td>
<td>12/19.71 0.61 (0.32,1.06)</td>
</tr>
<tr>
<td>AIDS</td>
<td>4/3.58 1.12 (0.30,2.86)</td>
<td>24/16.77 1.43 (0.92,2.13)</td>
</tr>
</tbody>
</table>
proportion of their excess mortality was from motor vehicle accidents.

For women the analyses by number of years of attendance show a slightly increased SMR 1.64 (95% CI 1.14, 2.29) for death from any cause among those women who attended school for ≤2 years. As with men this increased risk is accounted for mainly by excess deaths attributable to external causes. Among women who remained in school for ≥3 years we found fewer deaths than expected for these same categories of death.

We also examined mortality by year of graduation, dividing the cohort into those who graduated or were scheduled to graduate before 1980 and from 1980 forward. We found no clear pattern. We repeated all analyses using the state of Texas as the comparison population, with little difference in the results.

Discussion

Port Neches-Groves High School alumni had overall mortality rates that were slightly lower than those of the United States as a whole. Follow up of women in a record linkage study such as this, where we lacked Social Security number, can be problematic because women are likely to change their names. To evaluate success of follow up among women, we can compare the SMR for overall mortality among women with that of men. We assume that no biological reason exists to increase or depress overall mortality to any considerable extent for only one sex among students of the high school. The study hypothesis relates to a cause of death that is but a small proportion of overall deaths, and should not measurably affect the overall mortality. Thus, overall mortality serves as an indirect evaluation of adequacy of follow up. We observed an overall SMR for women (0.89) that was higher than the overall SMR for men (0.84). This finding implies that the follow up for women was comparable with that of men, and thus indicates that the steps we took to find married names or other aliases for female students resulted in a follow up for women that was satisfactory. The small difference between the observed SMRs for total mortality and 1.0 could be because of some underascertainment of deaths in the cohort, or to real differences between the cohort and the population data, or some combination of these factors.

The excess of deaths attributable to benign neoplasms was an incidental finding. All six who died of benign neoplasms had brain tumours. In their cohort study, Divine and Hartman1 reported six deaths from benign neoplasms among butadiene workers, with 3.8 expected. They found no apparent trend with duration of employment. Although Delzell et al did not mention benign neoplasms in their published report,1 their interim data indicate a total of nine deaths from benign neoplasms in their overall cohort, with 11 expected.2 Meinhardt et al3 and Matanoski et al4 do not present findings on deaths from benign neoplasms.

The mortality rate in this cohort for lymphatic and haematopoietic cancer was little different from that of the United States as a whole; from US death rates, we would have expected about 11.5 deaths, slightly less than the 14 we observed. The moderate excess of male deaths was partially offset by a deficit among women. These variations for men and women may be just random fluctuations, as the numbers are small and subject to substantial statistical variability. Another possibility is that the excess in men stems from occupational exposures sustained after leaving school. An effect from exposure that occurred while attending high school should not have been limited to men; thus, the deficit in deaths from lymphatic and haematopoietic cancer among women is not consistent with an environmental effect on students while they were attending school.

This conclusion is reinforced by the analysis comparing short-term male students with longer term male students. The observed excess in lymphatic and haematopoietic cancer deaths among men is stronger for men who attended the high school for no more than two years (SMR = 3.2, based on four observed deaths vs 1.25 expected) than for men who attended the school for three or more years (SMR = 1.3, based on 6 observed deaths vs 0.05 expected). This difference by length of attendance at school is also not consistent with an environmental effect.

A weakness of the study is that we had to evaluate cancer mortality rather than incidence. Neither Texas nor neighbouring Louisiana had statewide tumour registries during the follow up period of this study. In addition, Port Neches is a small, rural community, and residents with serious disease seek diagnosis and treatment in other locations, such as Houston, Galveston, and Baton Rouge. Therefore it was not feasible to assess cancer incidence in the cohort with any hope of achieving the same degree of follow up that we achieved from the three separate mortality sources that we used.

Another weakness of the study is the lack of direct environmental measurements of styrene or butadiene in the school. The close proximity of the school to the plants would presumably have resulted in some environmental exposure from plant emissions, but these would be expected to be less concentrated than occupational exposures experienced by plant workers. Another difference between this study and studies of occupational exposure is the younger age of the population that we studied. Little is known about how age at exposure might influence any potential carcinogenic effect.

Three retrospective cohort studies1–3 and one nested case-control study4 have evaluated the mortality of workers employed in styrene-butadiene or butadiene monomer manufacturing. All three cohort studies found lower than expected mortality rates for all causes of death combined and all cancers combined. A combined analysis of the three cohort studies (using an interim report of one of the studies1) reported 36 cases of leukaemia with 34.2 expected, for an SMR of 1.05 (95% CI 0.74, 1.46).11

Nevertheless, increased death rates from certain lymphopoietic cancers have been re-
ported for some subgroups of employees exposed to styrene-butadiene. Among butadiene workers, Divine and Hartman found an excess of lymphosarcoma deaths (9 observed/4.7 expected).\textsuperscript{1} The increase was concentrated among men with less than five years of employment, first hired during the second world war, and employed in a job with the potential for varied amounts of butadiene exposure on a routine basis. There was no appreciable leukaemia excess (13 observed/11.5 expected).

The results of the one case-control study of styrene-butadiene workers, conducted by Santos-Burgoa et al\textsuperscript{12} within the cohort studied by Matanoski et al,\textsuperscript{3} are inconsistent with the findings in the three cohort studies, including the very cohort that formed the source population for the case-control study. Santos-Burgoa et al reported a seven to ninefold increased rate of leukaemia mortality among employees exposed to butadiene.\textsuperscript{4} In the same study, little association between leukaemia and styrene remained after the association with butadiene was taken into account. The remarkable discrepancy between the null result for leukaemia reported for some subgroups of employees exposed to styrene-butadiene. Among butadiene workers, Divine and Hartman found an excess of lymphosarcoma deaths (9 observed/4.7 expected).\textsuperscript{1} The increase was concentrated among men with less than five years of employment, first hired during the second world war, and employed in a job with the potential for varied amounts of butadiene exposure on a routine basis. There was no appreciable leukaemia excess (13 observed/11.5 expected).\textsuperscript{1} The increase was concentrated among men with less than five years of employment, first hired during the second world war, and employed in a job with the potential for varied amounts of butadiene exposure on a routine basis. There was no appreciable leukaemia excess (13 observed/11.5 expected).\textsuperscript{1}

The results of the one case-control study of styrene-butadiene workers, conducted by Santos-Burgoa et al\textsuperscript{12} within the cohort studied by Matanoski et al,\textsuperscript{3} are inconsistent with the findings in the three cohort studies, including the very cohort that formed the source population for the case-control study. Santos-Burgoa et al reported a seven to ninefold increased rate of leukaemia mortality among employees exposed to butadiene.\textsuperscript{4} In the same study, little association between leukaemia and styrene remained after the association with butadiene was taken into account. The remarkable discrepancy between the null result for leukaemia reported for some subgroups of employees exposed to styrene-butadiene. Among butadiene workers, Divine and Hartman found an excess of lymphosarcoma deaths (9 observed/4.7 expected).\textsuperscript{1} The increase was concentrated among men with less than five years of employment, first hired during the second world war, and employed in a job with the potential for varied amounts of butadiene exposure on a routine basis. There was no appreciable leukaemia excess (13 observed/11.5 expected).\textsuperscript{1}

A fourth cohort study, recently completed, assessed the mortality of 15,649 men employed at eight North American styrene-butadiene plants during the period 1943–1991.\textsuperscript{5} The study improved upon some shortcomings of the previous studies with its large size, long follow up period, use of objective classifications of exposure indices, and evaluation of mortality patterns by employment factors such as payroll status, duration of employment, time since hire, work location, and estimated cumulative exposure to styrene and butadiene monomer. The authors found a small to moderate increase in deaths from leukaemia for the overall group of styrene-butadiene industry workers (48 observed/37 expected). The excess was concentrated among “ever hourly” subjects with 10+ years worked and 20+ years since hire and among subjects in three job classifications with potential for high exposure to butadiene or styrene monomers. The authors inferred that the associations between employment factors and forms of lymphatic and haematopoietic cancer other than leukaemia were not causal.

To summarise, the conflicting evidence from occupational mortality studies of styrene-butadiene workers raised concern that students who attended a high school that borders a styrene-butadiene plant were at increased risk for certain haematopoietic cancers. Our long term follow up of these students indicates that their mortality rates from these cancers does not indicate any important increased risk for the cancers in question. We did find an unexpected increase in deaths from benign neoplasms.

We thank Ms. Galina Savikovsky and Ms. Rita Tsang for their programming support. We are grateful to the staff at Port Neches-Groves High School for their help and encouragement.

Funding: This study was funded by a group of styrene-butadiene manufacturers including: BASF Corporation, Chevron Chemical Company, Conoco Inc, The Dow Chemical Company, Exxon Company, USA, Goodyear Tire and Rubber Co, Mitsubishi, Canadian OXY Offshore Production Company, PeroTex, Philips Petroleum Company, Quantum Chemical Corporation, Solvita Inc, Texaco, Inc, Texas Petrochemicals Corporation, Union Carbide Corporation, and Uniroyal Goodrich Tire Company. This study was conducted under a contract that, at the outset, guaranteed the investigators the unfettered right to publish the results.

Conflicts of interest: None.