Relation between the longitudinal development of lipoprotein levels and biological parameters during adolescence and young adulthood in Amsterdam, The Netherlands

J W R Twisk, H C G Kemper, G J Mellenbergh, W van Mechelen

Abstract

Study objectives – To analyse the relation between the longitudinal development of total serum cholesterol (TC), high density lipoprotein cholesterol (HDL), and the TC/HDL ratio and the longitudinal development of the biological parameters body fatness (SSF), lean body mass (LBM), and cardiopulmonary fitness (VO₂-max). The relations were analysed with generalised estimating equations (GEE).

Setting – The relations were investigated with data from the Amsterdam growth and health study, a longitudinal study in which six measurements were carried out within a period of 15 years.

Participants – Altogether 98 females and 84 males aged 13 years at the start of the study.

Main results – Adjusted for lifestyle and other biological parameters, the longitudinal development of TC was inversely related to the development of LBM (standardised regression coefficient β = −0.27; p ≤ 0.01) and positively to SSF (βmales = −0.32; p ≤ 0.01 and βmales = 0.15; p ≤ 0.01). HDL was inversely related to LBM (β = −0.26; p ≤ 0.01) and positively to VO₂-max (β = 0.08; p ≤ 0.05). The TC/HDL ratio was positively related to SSF (βmales = 0.39; p ≤ 0.01 and βmales = 0.13; p ≤ 0.01) and inversely to VO₂-max (β = −0.09; p ≤ 0.05).

Conclusions – The longitudinal analyses showed that body fatness was related to a high risk profile with respect to hypercholesterolaemia, and cardiopulmonary fitness to a low risk profile. Furthermore, it was shown that using body mass index as an indicator of body fatness in relation to lipoprotein values, has some important drawbacks.

Because it is well known that the origin of atherosclerosis lies in early childhood, it is important to start prevention of atherosclerosis as soon in life as possible. Therefore risk factors regarding atherosclerosis, such as hypercholesterolaemia, obesity, and hypertension have to be identified at a young age. Early identification is only useful, however, if the predictability of early measurements of these risk factors for values later in life (ie, tracking analysis) has been assessed. If some degree of tracking is observed then one has to investigate which parameters influence the longitudinal development of these risk factors. In preventive strategies these influencing factors can than be manipulated in order to prevent the development of atherosclerosis later in life.

One of the most important risk factors for atherosclerosis is hypercholesterolaemia (ie, high values of total serum cholesterol (TC), low values of high density lipoprotein cholesterol (HDL), or (perhaps) more important, high values of the TC/HDL ratio.23 It is often shown that TC, HDL, or the TC/HDL ratio track well from childhood or adolescence into adulthood,24 indicating that influencing these parameters during adolescence or young adulthood can have important implications for the prevention of atherosclerosis later in life. Because of this, many studies have analysed the relation between TC, HDL, and/or the TC/HDL ratio and potential influencing factors during these younger age periods.7-22

Potential parameters that influence the longitudinal development of lipoprotein values can be divided into biological and lifestyle factors. The latter consist of smoking, alcohol consumption, dietary intake, and daily physical activity. Relevant biological parameters are cardiopulmonary fitness and body weight (ie body fatness). However, lean body mass (LBM), another component of body weight, seems to be related to lipoprotein concentrations.23-24 One problem is that most studies do not take into account the fact that biological and lifestyle parameters not only influence lipoprotein values, but also each other (eg, increasing daily physical activity enhances cardiopulmonary fitness, which influences body fatness or the inverse relation between smoking behaviour and body fatness, etc). Thus, analysis of the relation between the longitudinal development of lipoprotein concentrations and biological parameters has to be investigated correcting for relevant lifestyle parameters. In this paper we correct for dietary intake, smoking behaviour, alcohol consumption, and daily physical activity.

Another problem with studies investigating the relation between lipoprotein levels and biological parameters is that most of them have a cross sectional design,7 10 12 15-22 which limits the interpretation of the results. The few longitudinal studies investigating this problem are either experimental23 24 or observational.8 9 11 13-14
The results of the observational studies have limitations, however, because measurements at only two points in time are used in the analysis. In these studies the relation between changes in lipoprotein levels and changes in biological parameters, where change is defined as the difference between two measurements, is investigated. In this paper the relation between lipoprotein concentrations and relevant biological parameters is investigated longitudinally, not using two points in time but all the available longitudinal data. The data are derived from the Amsterdam growth and health study, which includes six longitudinal measurements between 13 and 27 years of age.

A comparison is made between univariate analyses, in which the influences of the biological parameters are investigated separately, and multivariate analyses in which the relation between the development of the lipoprotein concentrations and biological parameters is investigated correcting for relevant lifestyle parameters.

Methods

STUDY POPULATION
The Amsterdam growth and health study is a longitudinal study which started in 1977 with 307 subjects (148 males and 159 females) from the first and second forms of a secondary school in Amsterdam. At the start of the study the mean (SD) age of both the boys and the girls was 13.0 (0.8) years. From 1977 onwards four annual measurements were carried out. Altogether 233 subjects (102 males and 131 females) completed these four measurements. In 1985, 200 subjects returned to the laboratory for a fifth measurement and in 1991, 181 subjects were measured for the sixth time. The population used in this study consisted of these 181 subjects. The total number of missing observations in this population over all longitudinal measurements was about 2%. At the last measurement, in 1991, the mean age of both the males and females was 27.1 (0.8) years. For the variables of interest in this study, no drop-out effects were observed during the longitudinal period.

BIOPARAMETERS
Body fatness was quantified as the sum of the bicipital, tricipital, subcapular, and suprailiac skinfold thickness (SSF), measured in mm. Lean body mass (LBM) was determined from SSF and total body weight using regression equations. Body mass index (weight/height², BMI) was also calculated. Cardiopulmonary fitness was quantified as maximum oxygen uptake (VO₂-max), using a standard running test on a treadmill. VO₂-max was expressed in ml/min kg⁻¹.

LIFESTYLE PARAMETERS
Habitual food intake was measured by a modification of the cross check diet history interview, which was specially developed for the Amsterdam growth and health study. In this diet history, the subjects were asked to recall their usual food intake (frequency, amounts as well as the methods of preparation) during the previous month. From this information food characteristics were calculated with the Dutch food and nutrition table. The following food characteristics were assessed:

1. Energy intake relative to body weight (kcal/kg),
2. The intake of saturated fatty acids, expressed as a percentage of total energy intake, and
3. Cholesterol intake relative to total energy intake.

At the same dietary interview, the total amount of consumed alcohol (expressed in grams) was assessed.

Daily physical activity was measured with a structured interview covering the three months before the interview. It was expressed in a total weighted activity score, which combines both the duration and intensity of daily physical activities during work, school, sport, transportation, and leisure time. The weighted score was expressed in METs/week. Regular smoking behaviour (expressed in grams of tobacco smoked per week) was determined by questionnaire at the time of each measurement.

Extensive information about the methods used in the Amsterdam growth and health study has been given by Kemper.

ANALYSIS
The statistical model which was used to assess the relation between the longitudinal development of TC, HDL, and the TC/HDL ratio on the one hand and biological parameters on the other hand, correcting for certain lifestyle parameters, is explained in the Appendix. The parameters of the model were estimated with generalised estimating equations (GEE).

In the presented study three statistical analyses were carried out:

1. A "univariate" analysis, in which the longitudinal developments of TC, HDL, and the TC/HDL ratio were related to the development of each of the biological parameters separately, correcting only for gender, time, and biological age,
Table 1: Mean (SD) of total serum cholesterol (TC), high density lipoprotein cholesterol (HDL), and the TC/HDL ratio in relation to calendar age and gender.

<table>
<thead>
<tr>
<th>Calendar age (y)</th>
<th>TC (mmol/l)</th>
<th>HDL (mmol/l)</th>
<th>TC/HDL</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>δ</td>
<td>β</td>
<td>δ</td>
</tr>
<tr>
<td>13</td>
<td>4.5 (0.6)</td>
<td>4.4 (0.7)</td>
<td>1.5 (0.3)</td>
</tr>
<tr>
<td>14</td>
<td>4.3 (0.7)</td>
<td>4.4 (0.7)</td>
<td>1.4 (0.3)</td>
</tr>
<tr>
<td>15</td>
<td>4.1 (0.7)</td>
<td>4.4 (0.7)</td>
<td>1.3 (0.2)</td>
</tr>
<tr>
<td>16</td>
<td>4.0 (0.6)</td>
<td>4.4 (0.7)</td>
<td>1.3 (0.2)</td>
</tr>
<tr>
<td>21</td>
<td>4.5 (0.8)</td>
<td>4.9 (0.7)</td>
<td>1.2 (0.2)</td>
</tr>
<tr>
<td>27</td>
<td>5.0 (0.9)</td>
<td>5.3 (0.9)</td>
<td>1.2 (0.2)</td>
</tr>
</tbody>
</table>

Table 2: Mean (SD) of biological parameters lean body mass (LBM), sum of four skinfolds (SSF), and maximal oxygen uptake (VO_{2-max}) in relation to calendar age and gender.

<table>
<thead>
<tr>
<th>Calendar age (y)</th>
<th>LBM (kg)</th>
<th>SSF (mm)</th>
<th>VO_{2-max} (ml.min^{-1}.kg^{-1})</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>δ</td>
<td>β</td>
<td>δ</td>
</tr>
<tr>
<td>13</td>
<td>37.6 (5.5)</td>
<td>36.5 (4.8)</td>
<td>28.4 (10.9)</td>
</tr>
<tr>
<td>14</td>
<td>34.5 (6.1)</td>
<td>38.9 (4.6)</td>
<td>27.8 (10.7)</td>
</tr>
<tr>
<td>15</td>
<td>44.6 (6.0)</td>
<td>40.6 (4.3)</td>
<td>27.8 (10.3)</td>
</tr>
<tr>
<td>16</td>
<td>52.1 (5.7)</td>
<td>41.7 (4.4)</td>
<td>29.1 (9.7)</td>
</tr>
<tr>
<td>21</td>
<td>60.9 (5.8)</td>
<td>45.1 (4.8)</td>
<td>34.9 (11.9)</td>
</tr>
<tr>
<td>27</td>
<td>64.5 (6.0)</td>
<td>47.5 (4.9)</td>
<td>36.5 (13.5)</td>
</tr>
</tbody>
</table>

Table 3: Standardised regression coefficients (β) (95% confidence intervals) according to univariate and multivariate analysis regarding the longitudinal development of total serum cholesterol (TC), high density lipoprotein cholesterol (HDL), and the TC/HDL ratio and the longitudinal development of lean body mass (LBM), sum of four skinfolds (SSF), and maximal oxygen uptake (VO_{2-max}).

When a significant interaction was found between one of the biological parameters and gender, separate coefficients were calculated for males and females. In addition, interactions between time and each of the biological parameters were added to the model in order to investigate if the relations between lipoprotein levels and biological parameters were stronger at the beginning (adolescence) or at the end of the longitudinal period. The interaction terms were formed by multiplication of the two parameters involved.

All GEE analyses were carried out with the Statistical Package for Interactive Data Analysis (SPIDA).6

Results

Descriptive data on TC, HDL, and the TC/HDL ratio are shown in table 1. Mean values are given for males and females for each of the six measurements. In table 2, similar descriptive information is given for the biological parameters LBM, SSF, and VO_{2-max}.

The results of the longitudinal GEE analysis regarding the longitudinal relation between TC, HDL, and the TC/HDL ratio on one hand and the biological parameters SSF, LBM, and VO_{2-max} on the other are shown in table 3. In the univariate analysis all relations were significant except for the relation between LBM and the TC/HDL ratio. Correcting for other biological parameters (the first multivariate analysis) and correcting for the lifestyle parameters (the second multivariate analysis) made some of these relations - the inverse relation between VO_{2-max} and TC and the inverse relation between SSF and HDL - disappear. The results of the final multivariate analysis showed that LBM was inversely related to both TC and HDL. SSF was positively related to TC and the TC/HDL ratio. Both relations were stronger for males than for females. VO_{2-max} was positively related to HDL and negatively to the TC/HDL ratio. With regard to TC, a positive interaction (p≤0.05) was found between SSF and time and between LBM and time, indicating that the relation with SSF was stronger at the end of the longitudinal period and the relation with LBM was stronger at the beginning of the longitudinal period. With regard to the TC/HDL ratio, there was a strong negative interaction (p≤0.01) between VO_{2-max} and time, also indicating a stronger relation at the end of the longitudinal period. Because BMI is most frequently used as an indicator for body fatness, the longitudinal relation between BMI and the lipoproteins was also analysed. The correlation matrix between SSF, LBM, and BMI (table 4) showed that BMI was closely related to both SSF and LBM. Therefore in the multivariate analyses only a correction for VO_{2-max} was made. The results showed that the development of BMI was positively related to the development of TC, negatively to the development of HDL, and positively to the development of the TC/HDL ratio in both univariate and multivariate analyses (table 5). In the relationship with TC, a strong positive interaction with time was found (p≤0.01).

(2) A multivariate analysis in which the developments of TC, HDL, and the TC/HDL ratio were related to each of the biological parameters, correcting for the other biological parameters, gender, time, and biological age, and

(3) A multivariate analysis, in which the longitudinal development was related to each of the biological parameters, not only correcting for the other biological parameters, gender, time and biological age, but also for the lifestyle parameters.

After assessing the main effects, for each of the biological parameters, an interaction term with gender was added to the statistical model.

* p<0.05; ** p<0.01.
† When no significant interaction was found with gender, only one coefficient was calculated for males and females.
‡ Correcting for biological parameters; § Correcting for biological and lifestyle parameters.
Table 4  Pearson correlation matrix of lean body mass (LBM), sum of four skinfolds (SSF), and body mass index (BMI) for males and females at each of the six longitudinal measurements

<table>
<thead>
<tr>
<th>Age (y)</th>
<th>LBM/SSF</th>
<th>LBM/BI</th>
<th>SSF/BMI</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(\rho)</td>
<td>(\beta)</td>
<td>(\gamma)</td>
</tr>
<tr>
<td>13</td>
<td>0.13 0.38*</td>
<td>0.66* 0.70*</td>
<td>0.62* 0.79*</td>
</tr>
<tr>
<td>14</td>
<td>0.15 0.31*</td>
<td>0.65* 0.68*</td>
<td>0.67* 0.78*</td>
</tr>
<tr>
<td>15</td>
<td>0.22 0.32*</td>
<td>0.66* 0.66*</td>
<td>0.68* 0.78*</td>
</tr>
<tr>
<td>16</td>
<td>0.08 0.13 0.65*</td>
<td>0.59* 0.74*</td>
<td>0.46* 0.74*</td>
</tr>
<tr>
<td>21</td>
<td>-0.13 0.15</td>
<td>0.55* 0.59*</td>
<td>0.46* 0.74*</td>
</tr>
<tr>
<td>27</td>
<td>0.06 0.03</td>
<td>0.57* 0.51*</td>
<td>0.60* 0.73*</td>
</tr>
</tbody>
</table>

* \(p<0.01\).

Discussion

In this study the longitudinal relation was investigated between TC, HDL, and the TC/HDL ratio and the biological parameters, SSB, LBM, VO2-max, and BMI over a period of 15 years in subjects aged from 13 to 27 years. The relation was investigated using all available longitudinal data and correcting for the lifestyle parameters: smoking behaviour, alcohol consumption, dietary intake, and daily physical activity. The statistical model used (equation 1), however, is certainly not the only way to analyse these longitudinal data. In the model presented here, for instance, Yi is related to \(X_{i0}\) while a time lag between the dependent and the independent variables could also be modelled (eg, relating Yi to \(X_{i0-1}\)). Another possibility is to model changes in parameters between the different time points instead of the actual values. The results of this study should be interpreted with respect to the limitations of the statistical model used.

Before discussing our results it should be stated that these should be interpreted with some caution, because they were based on a study with a relatively small number of subjects (181). Furthermore, it should be remembered that it is difficult to compare our results with those of other studies, because we used a new method which combines cross sectional and longitudinal effects into one coefficient. To our knowledge, there are no longitudinal studies investigating the relation between lipoproteins and biological parameters which have used longitudinal data analysing techniques like GEE.

Body fatness, operationalised as SSB, was highly positively related to TC and the TC/HDL ratio. This relation was stronger for females than for males. The positive interaction between SSB and time regarding the relation with TC indicates that the relation between SSB and TC is stronger at the end of the longitudinal period than during adolescence.

Most studies investigating the relation between body fatness and lipoprotein values during adolescence and young adulthood had used BMI rather than SSB as an indicator of body fatness. In cross sectional studies, BMI was found to be positively associated with TC, inversely with HDL, and positively with the TC/HDL ratio in both males and females,\(^7\) which is comparable to our results. In the study of Garn et al\(^10\) body fatness was, as in our study, quantified as SSB. They found a positive relation between skinfold thickness and TC among boys and girls from 15 to 19 years old. Kikuchi et al\(^2\) used subscapular skinfold thickness as indicator for body fatness and they found a negative relation between skinfold thickness and HDL.

There are only a few longitudinal studies investigating the relation between a change in body fatness and a change in lipoprotein levels during adolescence and young adulthood. In the Bogalusa heart study,\(^4\) a positive relation between a change in body fatness (body fatness quantified as triceps skinfold thickness) and a change in TC was found for boys, but not for girls. In addition to that a negative relation between a change in body fatness and a change in HDL was also found. The relation was investigated over a period of five years, with subjects aged between 5 and 12 years at the beginning of the study. In the Muscatine study,\(^11\) a change in body fatness (body fatness operationalised as BMI) over a period of 10 years was significantly related to TC in boys and girls with an initial age from 8 to 18 years. The relation between body fatness and TC probably results from an increased flow of free fatty acids to the liver, which is followed by an increased production of very low density lipoprotein, ie, higher values of TC.\(^4\)

Thus, positive relations between body fatness and TC and the TC/HDL ratio have been found in both cross sectional and longitudinal studies, which is equivalent to our results. In most studies, however, a negative relation between body fatness and HDL has been found, while in our study only in a univariate analysis was this relation found. The reason for this contradictory result is the fact that in the multivariate analyses the relation between body fatness (ie, the SSB) is analysed correcting for LBM, which seems to be related more to the development of HDL than SSB (table 3). LBM was not only negatively related to HDL, but also negatively to TC. With regard to the TC/ HDL ratio, often mentioned to be more important with respect to atherosclerosis than the single values of TC or HDL,\(^23\) there was no relation with LBM in our study.

As mentioned before, in most studies body fatness is quantified as BMI. BMI, however,
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not only reflects body fatness, but also LBM. This is shown in the correlation matrix between the SSF, LBM, and BMI at each of the six longitudinal measurements (table 4). From the correlation matrix along with the results of the longitudinal relations it can be concluded that the inverse relation between BMI and HDL probably reflects the relation between LBM and HDL, rather than the relation between body fatness (ie, SSF) and HDL, which is often suggested. The reason why LBM was inversely related to TC and HDL is not clear. Perhaps it has something to do with low plasma glucose and insulin concentrations, both of which are not only related to high LBM, but also to lower TC values.

Cardiopulmonary fitness, represented in our study by VO\(_2\)max, was positively related to HDL and negatively to the TC/HDL ratio. This indicates that a higher VO\(_2\)max is related to a lower risk profile in terms of hypercholesterolaemia. There are many studies investigating the relation between lipoprotein levels and physical fitness. The results, however, are controversial. Dwyer et al\(^{22}\) studied children of 9, 12, and 15 years of age in the Australian schools health and fitness survey. They used the physical work capacity on a bicycle ergometer at a heart rate of 170 beats per minute (PWC\(_{170}\)) as the indicator for cardiopulmonary fitness and they did not find a cross sectional relation between fitness and TC and between fitness and HDL. Sallis et al\(^{12}\) and Armstrong et al\(^{13}\), however, did find a cross sectional relation in the same age groups between cardiopulmonary fitness (defined as VO\(_2\)max) and TC, HDL, and the TC/HDL ratio. But when they corrected their results for body fatness, the relationships disappeared. Hofman and Walter\(^{11}\) investigated the relation between physical fitness and lipoprotein levels longitudinally over a period of five years in boys and girls. The initial age of the boys and girls was 9, and fitness was defined as a heart rate recovery index after a step test. They found a relation between the change in fitness and the change in HDL for boys, but not for girls. Dwyer et al\(^{8}\), on the other hand, did not find a relation between a change in fitness (determined as running time on a treadmill) and a change in HDL and the TC/HDL ratio among men aged 21 years over a period of nine weeks.

The differences in our results compared with others are partly due to the fact that in most studies the relation is investigated without correction for any lifestyle parameter. Only in the study of Dwyer et al\(^{8}\) was the relation investigated longitudinally, correcting for both lifestyle parameters and body fatness. However, these authors studied the relation over a time period of nine weeks, which is very short.

The fact that subjects with high cardiopulmonary fitness showed higher HDL levels and therefore lower values of the TC/HDL ratio is probably caused by an increased activity of lipoprotein lipase (LPL) and lecithin: cholesterol acyltransferase (LCAT). Both activities are increased by high levels of cardiopulmonary work and both are known to increase the concentration of HDL.\(^{42,44}\) This effect is irrespective of the effect of body fatness, because the amount of body fatness is not associated with LPL activity.\(^{42}\)

In this study a comparison was made between univariate analyses, in which each of the biological parameters was analysed separately correcting for time, gender, and biological age, and multivariate analyses in which the biological parameters were analysed correcting for each other and correcting for lifestyle parameters. The most striking result of this comparison was the fact that in a multivariate analysis body fitness (ie, SSF) was not related to HDL, which was the case in the univariate analysis and which was also reported in many other studies.

In addition to the comparison between univariate and multivariate analyses, the impact of the control for lifestyle parameters could also be assessed. Comparison of the two multivariate analyses revealed that the influence of the adjustment for lifestyle parameters is only obvious in the longitudinal relation between VO\(_2\)max and HDL and the TC/HDL ratio. With correction for lifestyle parameters the standardised regression coefficient for VO\(_2\)max was 0.19 and in relation to the TC/HDL ratio it was 0.08. Without adjustment, these coefficients were both 0.11. So without the adjustment for lifestyle parameters the coefficients for VO\(_2\)max were overestimated more than 20%. The fact that all other standardised regression coefficients were not or only slightly influenced by the adjustment for lifestyle parameters indicates that the correction for other biological parameters is more important than the correction for lifestyle parameters.

Although this study is quite complete with regard to the inclusion of both biological and lifestyle parameters, there are also potential confounding parameters which were not taken into account. Parameters like the use of oral contraceptives, pregnancy, and breastfeeding are parameters which could be related to body fatness as well as to the lipoprotein levels. Perhaps the lack of controlling for these gender specific parameters helps to explain the relatively low regression coefficients found for females compared with males in terms of the longitudinal relation between SSF and TC and the TC/HDL ratio.

CONCLUSIONS
From 13 to 27 years of age, body fatness was positively related to TC and the TC/HDL ratio. Thus, body fatness is related to a high risk profile with regard to hypercholesterolaemia. No relations were found between body fatness (ie, SSF) and HDL, which contradicts published findings. The unexpected results probably arose because of lack of correcting for confounding variables in most studies and the use of BMI as indicator for body fatness. BMI not only reflects body fatness, but also LBM, which was found to be inversely related to HDL. Cardiopulmonary fitness was related to a positive risk profile in respect of hypercholesterolaemia, as indicated by the positive effect.
longitudinal relation between V0.7-max and HDL and the inverse relation with the TC/HDL ratio. V0.7-max was not related to TC. In the longitudinal relation between lipoprotein levels and biological parameters, the adjustment for other biological parameters seemed to be more important than the correction for lifestyle parameters.

Appendix
To analyse the longitudinal relations the following statistical model was used:

\[ Y_{ij} = \beta_0 + \sum_{j=1}^{J} \beta_{ij} X_{ij} + \sum_{k=1}^{K} \beta_{ik} Z_{ik} + \beta_{jt} t + \sum_{m=1}^{M} \beta_{jm} G_m + \epsilon_{ij} \]

where: \( Y_{ij} \) = observations of subject \( i \) at time \( t \); \( \beta_{ij} \) = regression coefficient of independent variable of interest \( j \); \( J \) = number of independent variables; \( Z_{ik} \) = time dependent covariate \( k \) of subject \( i \) at time \( t \); \( \beta_{ik} \) = regression coefficient of time dependent covariate \( k \); \( K \) = number of time dependent covariates; \( t \) = time; \( \beta_{jt} \) = regression coefficient of time; \( G_m \) = time independent covariate \( m \) of subject \( i \); \( \beta_{jm} \) = regression coefficient of time independent covariate \( m \); \( M \) = number of time independent covariates; and \( \epsilon_{ij} \) = measurement error of subject \( i \) at time \( t \).

In this model the coefficients of interest are \( \beta_{ij} \) because these standardised regression coefficients reflect the magnitude of the relation between the longitudinal development of the lipoprotein levels (\( Y_{ij} \)) and the development of the different biological parameters of interest ie, SSF, LBM, V0.7-max, and BMI (\( X_{ij} \)). This relation (between \( Y_{ij} \) and \( X_{ij} \)) is not only analysed under correction of time (\( t \), but also under correction of time dependent covariates (\( Z_{ik} \)) and time independent covariates (\( G_m \)).

The lifestyle parameters (ie, total energy intake, the intake of saturated fatty acids, cholesterol intake, smoking, alcohol consumption, and the amount of daily physical activity) are also added to the model as time dependent covariates. Another time dependent covariate is biological age, whereas gender (\( G_t \)) is the only time independent covariate in the model.

The parameters of the statistical model are estimated with generalised estimating equations (GEE). With GEE, the relations between the variables of the model at different time points (ie from \( t_1 \) to \( t_2 \)) are tested simultaneously. In the model, both the dependent variables (\( Y_{ij} \)) and the independent variables of interest (\( X_{ij} \)) were transformed into z-scores to abolish scale differences between the variables. Thus, as a result of the GEE analysis for each biological parameter of interest (\( X_{ij} \)), an adjusted standardised regression coefficient was calculated (\( \hat{\beta}_{ij} \)), which can be interpreted as an adjusted longitudinal correlation coefficient.

The estimated \( \hat{\beta}_{ij} \)’s reflect the relation between the longitudinal development of TC, HDL, or the TC/HDL ratio and the development of the corresponding biological parameter, with the development measured over a period of 15 years using all available longitudinal data at the six time points. With GEE analysis a pooled analysis of cross sectional and longitudinal relations is carried out. This means that the standardised regression coefficient \( \hat{\beta}_{ij} \) combines the within subjects (ie, longitudinal) effects with the between subjects (ie, cross sectional) effects into one coefficient.

Because the repeated observations within one subject are not independent of each other, in GEE analysis, a correction is made for the within subject correlations. This correction is carried out by assuming a priori a certain correlation structure for the repeated measurements of the dependent variable (ie, the lipoprotein measurements). First of all, an independent structure can be considered. In this structure all within subject correlations are assumed to be zero. However, the within subject correlations in the data set under consideration are far from zero. The lowest Pearson correlation coefficient between two longitudinal measurements for TC, for instance, is around 0.5. Secondly, an exchangeable correlation structure can be considered. In this structure all correlations are assumed to be equal, irrespective of the time period under consideration. In addition, this structure is not suitable for the data set of the Amsterdam growth and health study. This is because the total time period under consideration is quite long. For instance for TC, the Pearson correlation coefficients between two longitudinal measurements vary between 0.5 and 0.8. The third possible structure is a stationary \( m \)-dependent structure, which means that correlations \( k \) measurements apart are the same for \( k = 1, \ldots, m \), while the within subject correlations between measurements more than \( m \) occasions apart are assumed to be zero. Because of the high values of the within subject correlations and the fact that the within subject correlations vary between the different interperiods, a \( m \)-dependent correlation structure is assumed to be the best estimate of the “real” correlation structure for TC, HDL, as well as the TC/HDL ratio. Therefore all GEE analyses are carried out with this a priori assumed correlation structure.

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6 Guo S, Becker L, Chumlea WC, Roche AF, Siervogel RM.
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