Body fat and stroke: unmasking the hazards of overweight and obesity

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Abstract

Study objective – It has been frequently noted that overweight and obesity have a stronger relationship to hypertension and diabetes mellitus than to the risk of stroke. The reason for this observation has not been clear. This study aimed to examine the lifelong relation between body fat and stroke to shed light on why the public health risks of overweight and obesity have tended to be obscured in previous epidemiological studies.

Design – Case-control study.

Setting – Eleven general practices in West Birmingham.

Participants – Altogether 125 men and women who had just had their first stroke and were aged 35–74 years and 198 controls frequency matched for age and sex were recruited over 24 months during 1988–90.

Main results – Those in the thinnest and fattest quartiles of subscapular skinfold thickness were at increased risk of stroke compared with those in the middle quartiles (age adjusted odds ratios 2·12 (95% confidence interval CI 1·2–3·0) and 2·0·8 (1·1–3·8) respectively). When lifelong maximum reported body mass index was assessed the hazards of obesity but not leanness were seen (odds ratio for the highest versus the lowest quartile were – age adjusted, 1·54 (0·8–3·0) and multiple risk factor adjusted, 2·5 (1·1–4·5). This lifelong pattern of risk seemed to be established early, the odds ratios for the highest versus the lowest quartile of reported body mass index aged 21 years were – age adjusted, 2·18 (1·1–4·4) and multiple risk factor adjusted 2·13 (1·1–4·2). The risks of both maximum reported body mass index and reported body mass index aged 21 years were more clear in those who had never smoked cigarettes (test for trend in odds ratio, p = 0·009 and p = 0·02 respectively).

Conclusions – Potentially important risks of excessive body fat for stroke can be obscured by both a history of cigarette smoking and thinness associated with deteriorating health. The results seem to explain why excess body fat has previously been consistently related to hypertension and diabetes mellitus but less consistently to stroke. Avoiding overweight and obesity during adult life offers protection against stroke.

The relationship between body fat and stroke has remained surprisingly unclear. Expert committees have not yet agreed that overweight is an established risk factor for stroke, although it is accepted to be causally related to hypertension and diabetes mellitus, both of which are risk factors for stroke. Many cohort studies seem to indicate little or no adverse risk for stroke with increasing levels of body mass index. It is unclear, however, why overweight is less clearly linked to stroke than it is to hypertension.

Assessment of body fat in epidemiological studies has, conventionally, used relative weight or body mass index. There is now increasing interest in the role of central or truncal body fat as a stronger indicator of both cardiovascular and stroke risk than body mass index. These initial observations were based on modest numbers of stroke patients and warranted investigation in studies with larger patient numbers.

There is suggestive evidence from a north Carolina cohort yielding 27 strokes, that the lifelong history of body fat is important in the development of stroke. This observation has not been followed up but suggested the value of a case-control approach inquiring retrospectively into indices of obesity over a lifelong span. To explore further the nature of the relation between body fat and stroke, subscapular skinfold thickness, body mass index, and lifelong patterns of body fat were assessed in a community based case-control study which aimed to examine a range of lifestyle related risk factors for stroke – the West Birmingham stroke project. The case-control approach allows indices of obesity at the time of the stroke to be assessed together with lifelong patterns.

Methods

The detailed methods of the case-control study have been previously described. People registered with 11 general practices in West Birmingham comprised the study population of 100,000.

STROKE PATIENTS

Stroke patients included were male and female patients between the ages of 35 and 74 years, whose first stroke occurred between 1 October 1988 and 30 September 1990. Patients were assessed for inclusion by one of us (RS) using the standard World Health Organization criteria - rapidly developing clinical signs of focal and at times global (applied to patients in
deep coma and to those with subarachnoid haemorrhage) loss of cerebral function, with symptoms lasting more than 24 hours or leading to death, with no apparent cause other than that of vascular origin. To ensure maximum recruitment, a representative in each practice was telephoned weekly and admission records for Dudley Road Hospital were regularly checked for patients with preliminary diagnoses suggestive of stroke. Both hospital and community stroke patients, if fit, were offered a computerised tomogram of the head (without contrast). Cerebral infarction was diagnosed if a head scan or necropsy was performed within 21 days of stroke onset and there was either no evidence of haemorrhage or a pattern of predominant infarction. The following characteristics excluded stroke patients from the study: a prior history of stroke (but not transient ischaemic attack), a history of mitral valvular heart disease combined with atrial fibrillation, a history of primary or metastatic cerebral neoplasm, a history of coagulation disorder or myeloproliferative disease, patients dying before assessment could be made, and those declining participation.

CONTROL SUBJECTS Subjects were randomly selected from the general practice population to match broadly by age and sex the incident stroke case distribution (frequency matching). A high response rate was sought through close liaison with the general practitioner when the subject was approached. The same exclusions applied as for stroke patients.

ASSESSMENT The cases and controls were all anthropometrically assessed and interviewed by a single observer (RS) using a standard structured questionnaire. Stroke patients were seen as soon as possible after the event. When disability prevented an adequate response the closest relative or friend was interviewed. Harpenden skinfold calipers (John Bull British Indicators Limited, England) were used to measure subscapular skinfold thickness by standard methods. Skinfold measurements were, with few exceptions, recorded with the subject seated. Two records for each subject were made and the mean calculated. All subjects, when possible, were weighed in light clothing using recently calibrated portable scales with an electronic display (Soehnle, West Germany). Standing height was also measured, when possible, by standard methods using a portable stadiometer (CMS Weighing Equipment Limited, London). Body mass index was calculated as weight in kg/(height in m)^2. Lying height was measured in all stroke patients admitted to hospital. For lying height the subject lay on his/her back and an assistant ensured the heels were correctly positioned at the base of the stadiometer. A simple regression equation which related standing and lying height among stroke patients who could stand enabled standing height to be estimated for those unable to stand. Cases and controls were asked: “What is the most you have ever weighed?” and “About how much did you weigh when you were 21 years old?” The responses were converted into kg and combined with measured/derived standing height to provide maximum reported body mass index and reported body mass index aged 21.

Measured/derived standing height was combined with the maximum recorded weight, if a record was available, in the general practice or hospital notes to provide maximum recorded body mass index. For all measures of body fat subjects were categorised by quartiles derived from control subjects; these were established separately for males and females. Variables which could confound or interact with the relation of body fat and stroke were assessed by questionnaire and have been detailed previously. Local ethical committee approval had been obtained for the study.

STATISTICAL ANALYSIS A computer software program (EGRET) using methods described by Breslow and Day was used to calculate odds ratios (relative risk), their 95% confidence intervals and test for trend in the odds ratios. Mantel-Haentzel methods were used to provide an estimate of the odds ratio after adjustment for one or two other variables. Adjustment was performed using the multiple logistic regression model if there was a need to adjust for more than 2 variables. Regression models were fitted using exposure levels of a given risk factor as discrete (factored) variables. Agreement between reported weights was assessed using 95% limits of agreement.

Results One hundred and twenty five eligible stroke patients were identified over two years. Of the excluded patients 30 had had a prior event, 8 had died before assessment was possible, 3 had haematological disorders, and 1 had mitral valve disease with atrial fibrillation. No stroke patient (or relative) declined to participate. Cerebral infarction was diagnosed in 81 patients, intracerebral haemorrhage in five, and subarachnoid haemorrhage in four. In the remaining 35 patients, computed tomography within 21 days was either impossible or declined by the patient, relative, or attending physician. Assessment was within 7 days for 85 (68%) and within 14 days for 112 (90%).

<table>
<thead>
<tr>
<th>Table 1</th>
<th>Relation between subscapular skinfold thickness and stroke</th>
</tr>
</thead>
<tbody>
<tr>
<td>Subscapular skinfold thickness</td>
<td>Age adjusted odds ratio (95% CI)</td>
</tr>
<tr>
<td><strong>Males:</strong></td>
<td></td>
</tr>
<tr>
<td>1st quartile</td>
<td>0</td>
</tr>
<tr>
<td>2nd quartile</td>
<td>0.67 (0.25, 1.76)</td>
</tr>
<tr>
<td>3rd quartile</td>
<td>0.38 (0.12, 1.09)</td>
</tr>
<tr>
<td>4th quartile</td>
<td>0.99 (0.37, 2.28)</td>
</tr>
<tr>
<td><strong>Females:</strong></td>
<td></td>
</tr>
<tr>
<td>1st quartile</td>
<td>0</td>
</tr>
<tr>
<td>2nd quartile</td>
<td>0.44 (0.11, 1.36)</td>
</tr>
<tr>
<td>3rd quartile</td>
<td>0.43 (0.13, 1.38)</td>
</tr>
<tr>
<td>4th quartile</td>
<td>1.08 (0.39, 3.00)</td>
</tr>
</tbody>
</table>
Unmasking obesity/stroke relation

The median ages of the patients with stroke and the controls were 66 and 63 respectively. Surrogate responses were required for 30 of the 125 cases and 2 of the 198 controls.

SUBSCAPULAR SKINFOLD THICKNESS

The following values defined quartiles for subscapular skinfold thickness — males; 13.7, 17.9, and 23.3 mm; females; 15.2, 22.5 and 30.1 mm. The age adjusted relation between subscapular skinfold and stroke was “U” shaped in both males and females (table 1). As similar patterns were seen in both sexes the results were pooled (fig 1). Individual adjustment of the crude odds ratio by the following had a negligible effect; race, social class, cigarette smoking, alcohol consumption, exercise aged 15 to 25, recent exercise, dietary fat, history of cardiac ischaemia, and family history of stroke. Adjustments for diabetes and hypertension history reduced the relative risk seen among subjects in the upper quartile: the odds ratio for upper versus lower quartiles were, unadjusted — 0.95, adjusted for history of hypertension — 0.80, and adjusted for history of diabetes mellitus — 0.77. Multiple risk factor adjustment of the relation in both sexes combined made the “U” shape more, rather than less, pronounced (fig 1). The inclusion of hypertension history and history of diabetes in the adjustment was not considered appropriate as it is now widely accepted that these two risk factors are in substantial measure a consequence of overweight. Adjustment would, therefore, only be appropriate in exploring the extent to which hypertension and diabetes were part of a causal pathway involving overweight and stroke. As a pretested objective of the study was to examine a possible increased risk in thin as well as overweight people, an analysis was conducted comparing the upper and lower quartiles with the middle two. The odds ratios, adjusted for age, for the lowest quartile was: 2.12 (95% CI 1.2, 3.9) and for the upper quartile: 2.08 (95% CI 1.1, 3.8).

BODY MASS INDEX AND STROKE

Standing height and weight could be measured in all control subjects and 90 of the less disabled stroke patients. Quartiles of body mass index (kg/m²) were defined by the following values: males — 23.7, 25.9 and 28-4 and females — 23-5, 26-6, and 30-1. The risk of stroke did not seem to rise with increasing body mass index (table 2). The pattern was not appreciably altered by multiple risk factor adjustment (covariables as for the subscapular skinfold analysis). A tendency for heavier patients to be more disabled and thus excluded from this analysis cannot be ignored but any effect was likely to be modest.

MAXIMUM REPORTED BODY MASS INDEX

The excess risk seen among those in the thinnest quartile of subscapular skinfold is quite possibly a result of the fact that stroke is not uncommonly a late event in those who are
unwell from a variety of medical conditions. To explore this possibility the relation of stroke with maximum reported body mass index was assessed in subjects who could recall maximum weights (cases = 124, controls = 196). Quartiles were defined by the following values (kg/m²): male - 25-2, 27-9, and 30-5 and female - 24-8, 28-4, and 31-3. The apparent excess risk of stroke in thin people disappeared when the relation between lifelong maximum body mass index and stroke was examined (table 2). The hazards of obesity, however, were not obscured. Most stroke patients (26/43) (60%) in the thinnest quartile of subscapular skinfold thickness had previously been overweight (maximum reported body mass index quartiles 2 to 4). The corresponding figure in controls was 17/50 (34%). As a guide to the reliability of recalled maximum weights, responses from stroke patients were compared with those of their relative/friend in a subsample of 20 cases. The surrogate responses had a mean weight only 0.41 kg higher than that of the cases themselves. The 95% limits of agreement were 9.1 kg to -9.9 kg (correlation coefficient = 0.95). 18 of the 20 paired responses differed by 6.35 kg (14 lb) or less.

MAXIMUM RECORDED BODY MASS INDEX
Because a substantial number of cases (n = 97) and controls (n = 130) had had prior weights recorded, the relation between past maximum recorded body mass index and stroke could be examined. The interquartile values were lower than for the estimated values (male - 24-2, 27-9, and 29-9 and female - 24-0, 27-5, and 31-0) but the overall relation with the risk of stroke still indicated an increased risk in the most overweight quartile of body mass index (table 2). The "U" shape had reappeared in this analysis, probably because weights tended to be recorded in sicker subjects.

REPORTED BODY MASS INDEX AGED 21
Three stroke patients and one control could not recall their weight when aged 21. The interquartile values (kg/m²) in control subjects were: male - 21-4, 22-5, and 24-2 and female - 20-1, 21-8, and 23-8. The most overweight subjects were at significantly increased risk of stroke (table 2). The unadjusted odds ratio for the upper quartile of body mass index aged 21 (2-25) was reduced somewhat by adjustment for exercise age 15 to 25 (1-98) but not by maximum reported body mass index (2-39). The pattern in both males and females was similar for all the measures of body mass index (results not presented).

INTER-RELATION OF BODY FAT WITH CIGARETTE SMOKING
Because a clearer relation between overweight and stroke had been noted previously the odds ratios were examined after stratification by cigarette smoking status. In controls, 24/50 (48%) of those in the thinnest quartile of sub-

Table 4 Relation between body fat and stroke stratified by smoking status

<table>
<thead>
<tr>
<th>Subscapular skinfold thickness:</th>
<th>Quartile 1</th>
<th>Quartile 2</th>
<th>Quartile 3</th>
<th>Quartile 4</th>
<th>Test for trend</th>
</tr>
</thead>
<tbody>
<tr>
<td>Never smoker</td>
<td>(1/0)</td>
<td>0.05</td>
<td>0.54</td>
<td>1.10</td>
<td>p = 0.38</td>
</tr>
<tr>
<td>Former smoker</td>
<td>(6/13)</td>
<td>0.22</td>
<td>0.14</td>
<td>0.93</td>
<td>p = 0.71</td>
</tr>
<tr>
<td>Current smoker</td>
<td>(1/8)</td>
<td>0.84</td>
<td>0.96</td>
<td>0.64</td>
<td>p = 0.72</td>
</tr>
<tr>
<td>Maximum reported body mass index:</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Never smoker</td>
<td>(1/0)</td>
<td>5.13</td>
<td>3.58</td>
<td>12.00</td>
<td>p = 0.009</td>
</tr>
<tr>
<td>Former smoker</td>
<td>(1/0)</td>
<td>0.98</td>
<td>0.59</td>
<td>1.74</td>
<td>p = 0.25</td>
</tr>
<tr>
<td>Current smoker</td>
<td>(1/0)</td>
<td>0.95</td>
<td>0.55</td>
<td>0.93</td>
<td>p = 0.67</td>
</tr>
</tbody>
</table>

* Age groupings (35 to 54, 55 to 64 and 65 to 74 years)
Unmasking obesity/stroke relation

observed link between variability of body weight and coronary heart disease also noted in the Framingham study could be partly explained by the results of our study – those who become obese and then lose weight are at particularly high risk of cardiovascular disease.

Cigarette smoking tended to obscure the true risks of overweight in this study. These observations are consistent with data from the Whitehall and Framingham cohorts for both stroke and cardiovascular disease respectively.21,29 This interaction of effects is plausible as some smokers, particularly women, report that they smoke to avoid overweight. It must be emphasised, however, that although smoking may help in the short term with overweight it will increase the risk stroke.

The above findings may together explain why a considerable number of studies have concluded that overweight (derived from weight/height) is of minimal importance as a risk factor for stroke.2,10 These studies have either had relatively short follow up, body fat stratified by only two levels, or a preponderance of cigarette smokers.

It seems unlikely that the relation between excessive body fat and stroke is due to either study biases or the influence of confounding variables. Separate methods of assessment (subscapular skinfold thickness and maximum reported weight yielding maximum body mass index) both indicated an increased risk. Significant measurement bias by the investigator would seem unlikely because of the high level of repeatability of skinfold measurements. Recall bias is a potential problem in a study such as this, particularly with the inclusion of surrogate responses. Previously recorded weights, however, and the lack of appreciable bias noted when patients and surrogate responses were compared did not indicate particular problems in this study. Loss of height with age is unlikely to have produced an important bias – the impact on body mass index is small and any error introduced is likely to be non-differential. The effect of possible weight loss following stroke could only be small and would only have the effect of underestimating the importance of excessive body fat in this study. The potential problem of residual confounding in adjusted analyses seems unlikely to be important as there was no tendency for adjustment to attenuate the relations that were found (apart from the predicted effects of hypertension or diabetes).

The inter-relationships of diet, exercise, body fat, blood pressure, diabetes mellitus, and other related variables is complex and controversial. This study aimed to explore risks which were related to lifestyle and were potentially under personal control. It was, therefore, a prestaged decision not to present the results after adjustment for hypertension and diabetes mellitus. The relative risks of the upper quartiles of all measures of body fat were attenuated by adjusting for a history of diabetes or hypertension. This is likely to reflect the fact that dietary excess and inadequate exercise lead to overweight and its consequent increase in both blood pressure and risk of diabetes. It is being
realised increasingly that adjustment by variables in the causal pathway is inappropriate except to explore that pathway.10 If a hereditary tendency did tend to cause obesity along with hypertension and diabetes the overweight/stroke relation would have been expected to be reduced by adjustment for a family history of stroke. This was not observed.

CONCLUSION
This study provides support for the hypothesis that excessive body fat is linked to an increased risk of stroke. This is probably through mechanisms such as hypertension and diabetes mellitus. Intercurrent illness, cigarette smoking and inadequate length of follow up have all tended to obscure these risks in my previous studies. If the prevalence of cigarette smoking falls, the avoidance of excess body fat via dietary prudence and adequate exercise, could emerge more clearly as a lifestyle related risk factor with substantial preventive potential for stroke.

We thank the patients and staff of the 11 participating general practices (see below) for their cooperation, Dr R P Fletcher and Dr R Skinner for their advice and Mrs Beryl Bryant for clerical help. The project was funded by the Welcome Trust. The participating practices were Bearwood Road Surgery; Smethwick, Birchfield Medical Centre, Cape Hill Medical Centre, Church Lane Surgery; Handsworth, Hamstead Road Surgery; Handsworth/Great Barr, Lee Bank Health Centre, Newton Health Centre, Plough and Harrow Medical Centre; Ladywood, Sherwood House Surgery; Bearwood, Tower Hill Health Centre, Victoria Health Centre; Smethwick.

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