Ventilatory function, height, and mortality among lifelong non-smokers

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Abstract

Study objective—The aims were to determine the relationship between spirometric indices and mortality among lifelong non-smokers, and to investigate whether the association of short stature with increased risk of death is explained by reduced levels of ventilatory function in shorter men.

Design—The study was a nested (within cohort) case-control analysis of an 18 year prospective study of mortality.

Subjects—Participants were 3452 male civil servants aged 40–64 years at entry who denied ever having smoked tobacco.

Measurements and main results—408 men who died were matched to 2874 controls of the same age and height. Reduced one second forced expiratory volume (FEV₁) was associated with mortality from non-respiratory causes (rate ratio per litre decrease 1·44, 95% confidence interval 1·19–1·73). The ratio of FEV₁ to forced vital capacity was a weak predictor of mortality. Among 397 case-control sets matched for age and FEV₁, mortality was unrelated to height. Comparing mortality differentials across age adjusted tertiles of each risk factor, height adjusted FEV₁ was a stronger predictor of death than height, body mass index, or plasma cholesterol. FEV₁ adjusted for age but not for height was almost as strong a predictor as systolic blood pressure.

Conclusions—The determinants of ventilatory function in lifelong non-smokers may include causes of premature death. FEV₁ may be a more sensitive indicator than height of early life influences upon mortality.

Other studies, however, suggest that a real association may exist. In Baltimore, all cause mortality during 24 years of follow up was significantly and inversely related to FEV₁ adjusted for age and height among lifelong non-smokers. Similar findings have been reported for vital capacity among healthy non-smokers in the Framingham study, and the Copenhagen City Heart Study. A longitudinal study in Papua New Guinea found that the local methods of smoking tobacco were not associated with increased risk of death but that reduced ventilatory function was nevertheless an important predictor of all cause mortality.

Men of short stature are at increased risk of death and coronary heart disease. These associations have attracted considerable interest because height, in so far as it is influenced by environmental factors, is mainly determined during childhood and adolescence. A recent analysis of incident cases of myocardial infarction in the British Regional Heart Study suggested that the effect of height on risk of myocardial infarction may be explained by the tendency for short men to have lower levels of ventilatory function. However, that study included both smokers and non-smokers and therefore could not exclude the possibility that the closer relationship of myocardial infarction to forced expiratory volume than to height was a reflection of residual confounding by unmeasured aspects of smoking behaviour.

In order to clarify these issues, this paper reports the effect of ventilatory function and height upon 18 year mortality among a large cohort of British civil servants who denied ever having smoked tobacco.

Methods

The Whitehall study obtained questionnaire information, clinical measurements, and samples of blood from 19 018 male civil servants who were examined during 1967–1969. Subsequent deaths among these men have been identified through the National Health Service Central Register and the underlying cause of death has been coded throughout to the 8th revision of the International classification of diseases. This analysis is limited to deaths occurring on or before 31 January 1987 among men aged 40–64 years at examination who answered negatively to each of the following questions:

1. Do you smoke cigarettes now?
2. If you do not smoke cigarettes now, did you ever smoke them?
3. Have you ever smoked a pipe?
4. Have you ever smoked cigars regularly?
The examination at entry included measurement of height, weight, and blood pressure (using the London School of Hygiene sphygmomanometer with the subject seated). Three forced expiratory manoeuvres were recorded by Vitalograph spirometer and the means of the highest two measurements of forced expiratory volume in one second (FEV₁) and the corresponding measurements of forced vital capacity (FVC) were coded for computer analysis. Capillary blood samples were analysed for plasma cholesterol concentration. A standardised history of cardiorespiratory symptoms was obtained by questionnaire. A six limb lead electrocardiogram (ECG) was recorded and classified by the Minnesota code.

Subjects with a history of angina or severe anterior chest pain lasting an hour or more, or abnormal ECG tracings (Minnesota codes for Q/QS waves 1-1-1-3, ST depression 4-1-4-4, T wave inversion or flattening 5-1-5-3, or left bundle branch block 7-1) were grouped together as providing evidence of prevalent coronary heart disease at entry to the study.

STATISTICAL ANALYSIS

Statistical Analysis System (SAS)²⁰ was used for data processing. Two “nested” (within cohort) case-control studies were generated. Cases were defined as men dying during the 18 years of follow up. Matched controls were selected for each case from among the men who survived longer than the case from their date of examination. In the first case-control analysis, the matching criteria were age at examination (single years), and height (measured to the nearest half inch [1-3 cm]). In the second, the matching criteria were age (single years) and FEV₁ (recorded to the nearest 100 ml). In each study, all eligible controls were included, resulting in variable numbers of controls per case.

The matched case-control sets were analysed by conditional logistic regression using the Epidemiological Graphics Estimation and Testing package (EGRET).²¹ Case-control status was the outcome variable and measurements of ventilatory function and other cardiovascular risk factors were entered as explanatory variables. Coefficients from such regression models are equivalent to those obtained by proportional hazards regression²² and estimate the logarithm of the rate ratio per unit increase in an explanatory variable. In this instance these are mortality ratios and are all adjusted for the variables matched in the design (age and height in the first study, age and FEV₁ in the second), in addition to other explanatory variables in the model.

In order to compare the relative importance of various risk factors as predictors of mortality in the whole group of non-smoking men, the distribution of each risk factor was adjusted for age and standardised to a common scale. This was achieved by calculating the difference between the measurement for each man and the mean of measurements for all lifelong non-smokers of the same single year of age in the study population, and dividing this by the standard deviation of measurements for men of the same age. These age specific “z scores” were then ranked and the upper, middle, and lower thirds of the distribution were used in the analysis. The use of these age adjusted terciles removed any confounding effect of age upon the relationship between the level of each risk factor and subsequent mortality.

Results

STUDY POPULATION

Among 18 403 men aged 40–64 years at examination, 3455 (19%) denied ever having smoked cigarettes, cigars, or a pipe. Three men (all of whom died during follow up) had no spirometric measurements recorded. Among the remaining 3452 men there were 445 deaths during 60 084 person-years of follow up (average 17.4 years per man). These included 11 deaths from lung cancer.

The entry examinations preceded the adoption of internationally accepted criteria for definition of spirometric indices.²² However, a sample of the original spiromgrams from lifelong non-smokers was inspected as part of another study. The sum of the FEV₁ and FVC from each of the two “best tests”, as defined by the American Thoracic Society criteria,²² were within 5% of each other in 84% (73/87) spiromgrams, within 7% in 91% (79/87), and within 10% in 98% (85/87).

Forced expiratory volume in one second (FEV₁) was strongly related to age, height, and employment grade, which together explained 31% of the variance. The effect of employment grade was mainly due to a difference of 300 ml in mean age-height adjusted FEV₁ between the 2717 men in “high” grades (administrative, professional, and executive posts) and the 571 in “low” grades (clerical and messenger posts). One hundred and sixty four men in the British Council and Diplomatic Service (excluded from the above grade categories) had similar adjusted FEV₁ to those in the “high” grades. After adjustment for age, height, and grade, there remained considerable individual variation in FEV₁ (residual standard deviation = 573 ml).

HEIGHT MATCHED CASE-CONTROL ANALYSES

One or more controls of the same age and height were available for 408 (92%) of the 445 deaths. Overall, there were 2874 controls, an average of seven per matched case. The ventilatory function of the matched cases was not extreme: 320 (78%) had an FEV₁ in the range 2.5–4.0 litres.

The relationship between ventilatory function and all cause mortality was almost identical, whether FEV₁ or FVC was used in the analysis. The association was marginally stronger for FEV₁ (rate ratio per 1 litre decrease = 1.52, 95% confidence interval 1.27–1.83, \( \chi^2 = 20.1, \text{df} = 1 \)) than for FVC (rate ratio per 1 litre decline = 1.51, 95% CI 1.25–1.82, \( \chi^2 = 19.3, \text{df} = 1 \)). The ratio FEV₁/FVC was therefore not associated with overall mortality, although it was a significant predictor of death from respiratory disease (rate ratio per 10% decline = 2.12, 95% CI 1.93–2.77, \( \chi^2 = 8.74, \text{df} = 1 \)).

Table 1 shows the rate ratios (per litre decrease in FEV₁) for various causes of death. Reduced levels of FEV₁ (adjusted for age and height by matching) were associated with increased mortality for all groups of causes studied. This excess was most marked for respiratory disease.
but remained highly significant when non-respiratory causes were considered. Among the non-respiratory causes, the association with cancer deaths was weakest, and that with non-cancer, non-circulatory deaths was strongest. However, this heterogeneity of rate ratios could readily be due to chance ($\chi^2$ for interaction = 0.71, df = 2). When circulatory deaths were further subdivided, reduced FEV$_1$ was significantly associated with both coronary and non-coronary mortality. Fatal strokes accounted for half of the latter category.

The effect of FEV$_1$ on all cause mortality was examined by age at death and duration of follow up. There was no substantial variation in the rate ratios for death above and below 70 years ($\chi^2$ for interaction = 0.20, df = 1). The association of FEV$_1$ with death in the first 10 years of follow up (rate ratio per litre decrease = 1.94, 95% CI 1.46-2.57) was stronger than with later deaths (1.28, 1.01-1.64), the interaction being significant ($\chi^2$ = 4.90, df = 1).

Table II shows the mortality rate ratios (per 1 litre decrease in FEV$_1$) for circulatory diseases, before and after adjustment for known cardiovascular risk factors. Adjustment for employment grade (in five categories) increased the coefficient for FEV$_1$ very slightly. This was because matching for age and height in the design effectively removed grade differences in circulatory disease mortality. Further adjustment for systolic blood pressure, plasma cholesterol and body weight did not substantially alter the effect of FEV$_1$.

Inclusion of prevalent coronary heart disease as an additional covariate reduced the independent effect of FEV$_1$ (Table II) but the latter remained significant at the 5% level ($\chi^2$ = 3.99, df = 1). The association of FEV$_1$ with circulatory mortality was not confined to those with prevalent coronary heart disease ($\chi^2$ for interaction = 0.09, df = 1). Among those with no prevalent heart disease, the adjusted rate ratio per litre decline in FEV$_1$ was 1.30 (95% CI 0.96-1.76).

When 45 cases and 233 controls with a history of phlegm, wheeze, or exertional dyspnoea were removed from the analysis, the relationship between FEV$_1$ and circulatory death was strengthened: rate ratio per litre decrease = 1.55 (95% CI 1.16-2.08), compared to 1.49 among all case-control sets.

FEV$_1$ MATCHED CASE-CONTROL ANALYSES

One or more controls of the same age and FEV$_1$ were available for 397 (89%) of the 445 deaths. Overall, there were 2536 controls, an average of 6.4 per matched case.

Table III presents the rate ratios per 10 cm decrease in height for all causes of death and mortality from specific causes, adjusted for age and FEV$_1$ by matching. The effect of height on mortality was generally weak and was non-significant throughout. There was virtually no independent influence of height on mortality from circulatory disease at any given age and measured FEV$_1$.

The effect of employment grade upon mortality after adjustment for age and FEV$_1$ by matching was largely confined to respiratory deaths. Comparing “low” grades to “high” grades, the mortality rate ratio was 1.10 (95% confidence interval 0.83-1.47) for all causes of death, 1.06 (0.79-1.42) for non-respiratory causes, and 3.78 (0.74-19.3) for respiratory disease.

RELATIVE IMPORTANCE OF FEV$_1$ AS A PREDICTOR OF MORTALITY

The importance of measured FEV$_1$ and height adjusted FEV$_1$, as predictors of mortality in the whole population of 3452 lifelong non-smokers was explored by comparing the death rates in age adjusted tertiles of various risk factors measured at the entry examination. Table IV shows that differentials in all cause mortality across the tertiles of height adjusted FEV$_1$ were greater than the equivalent differentials for height, body mass index, or plasma cholesterol. The analysis of case-control sets matched for FEV$_1$ suggested that height had little effect upon mortality if FEV$_1$ was known. The convention of adjusting ventilatory function for height may therefore be considered as “overmatching” which would tend to reduce power to detect effects of reduced FEV$_1$ on mortality. When FEV$_1$ unadjusted for height was considered, the association across age adjusted tertiles was greater than for height adjusted FEV$_1$, and approached the magnitude of the trend for systolic blood pressure, the most powerful single predictor of mortality among these lifelong non-smokers.

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### Table I

<table>
<thead>
<tr>
<th>Cause of death</th>
<th>Cases (n)</th>
<th>Controls (n)</th>
<th>Rate ratio (95% CI) per litre decrease in FEV$_1$</th>
<th>$\chi^2$ (1 df)</th>
</tr>
</thead>
<tbody>
<tr>
<td>All causes</td>
<td>408</td>
<td>2874</td>
<td>1.52 (1.27-1.83)</td>
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<tr>
<td>Respiratory disease</td>
<td>20</td>
<td>78</td>
<td>4.16 (1.69-10.2)</td>
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<tr>
<td>Non-respiratory causes</td>
<td>388</td>
<td>2796</td>
<td>1.44 (1.91-12.3)</td>
<td>13.9</td>
</tr>
<tr>
<td>Malignant neoplasms</td>
<td>105</td>
<td>826</td>
<td>1.27 (0.89-1.82)</td>
<td>1.70</td>
</tr>
<tr>
<td>Circulatory disease</td>
<td>236</td>
<td>1649</td>
<td>1.49 (1.81-1.91)</td>
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<tr>
<td>Other non-respiratory causes</td>
<td>47</td>
<td>321</td>
<td>1.60 (0.96-2.66)</td>
<td>3.14</td>
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<tr>
<td>Coronary heart disease</td>
<td>171</td>
<td>1225</td>
<td>1.38 (1.04-1.84)</td>
<td>4.92</td>
</tr>
<tr>
<td>Other circulatory disease</td>
<td>65</td>
<td>424</td>
<td>1.87 (1.13-3.10)</td>
<td>5.91</td>
</tr>
</tbody>
</table>

*Matched analysis by conditional logistic regression

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### Table II

<table>
<thead>
<tr>
<th>Risk factor</th>
<th>Cases (n)</th>
<th>Controls (n)</th>
<th>Rate ratio (95% CI) per litre decrease in FEV$_1$ before adjustment</th>
<th>Rate ratio (95% CI) per litre decrease in FEV$_1$ after adjustment</th>
</tr>
</thead>
<tbody>
<tr>
<td>Employment grade</td>
<td>236</td>
<td>1649</td>
<td>1.49 (1.16-1.91)</td>
<td>1.50 (1.17-1.94)</td>
</tr>
<tr>
<td>Grade, systolic blood pressure</td>
<td>226</td>
<td>1536</td>
<td>1.39 (1.08-1.80)</td>
<td>1.37 (1.04-1.79)</td>
</tr>
<tr>
<td>Plasma cholesterol, body weight</td>
<td>225</td>
<td>1530</td>
<td>1.42 (1.10-1.83)</td>
<td>1.33 (1.01-1.75)</td>
</tr>
</tbody>
</table>

*Matched analysis by conditional logistic regression. Rate ratios before adjustment are calculated for the subset of cases and controls with complete data for each set of cardiovascular risk factors.

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Discussion
The restriction of this analysis to lifelong non-smokers was based upon self-reported smoking habit, and it is likely that some misclassification of smokers as non-smokers occurred. However, it is unlikely that such errors were common enough to generate an association between ventilatory function and mortality of the magnitude that was observed. It is possible that spirometric performance was influenced by preclinical disease present at examination, but the effect of FEV$_1$ on mortality was found even after 10 years follow up and also among subjects with no respiratory symptoms.

The observed association between reduced ventilatory function and risk of death in this cohort is therefore more likely to be real than spurious. Follow up was limited to fatal events, so it is not possible to distinguish whether the excess mortality among the men with low levels of FEV$_1$ was due to increased incidence of disease or to greater case fatality. The association of reduced ventilatory function with mortality was not confined to any specific cause of death, and it is likely that no single mechanism accounts for all these relationships.

Early interest in the relationship between vital capacity and cardiovascular disease was stimulated by the hypothesis that ventilatory function might be an indicator of physical activity or physical fitness, both of which have been found to predict mortality. However, habitual activity is not an important determinant of ventilatory function in adults and forced vital capacity is a poor predictor of physical fitness, so it is unlikely that these mechanisms could account for more than a small part of the association between reduced FEV$_1$ and all cause mortality.

Diminished lung function may be an indicator of environmental influences upon mortality, or it may be implicated directly in the causal pathway, particularly as a determinant of case fatality. Little is known of the constitutional or environmental determinants of ventilatory function among lifelong non-smokers, and the present study suggests that this forms an important agenda for future research. In general terms, FEV$_1$ may be affected either by the development of the lung or the occurrence of airway obstruction. In this male cohort, as among the mostly female lifelong non-smokers in the Copenhagen City Heart Study, the ratio of FEV$_1$ to FVC was relatively unimportant as a predictor of death, suggesting that the capacity of the lungs, which is in part determined by stature, may be the more important variable.

The analysis of case-control sets matched for age and FEV$_1$ supported the suggestion that the association between height and coronary heart disease is explained by lower absolute levels of ventilatory capacity in shorter men. This conclusion also applied to other causes of death. The convention of adjusting lung function measurements for height in epidemiological studies may deserve reconsideration, particularly where the prediction of subsequent morbidity or mortality is of interest. Height has attracted epidemiological interest as a marker of early dietary and environmental influences upon subsequent disease outcomes. The findings here suggest that such influences, if they exist, may be more adequately indicated by the ventilatory function of the adult lung. Respiratory function also offers a plausible mechanism for a causal relationship between experiences in childhood and mortality in middle age.

When compared in a standardised way to widely recognised risk factors for premature death, FEV$_1$ appeared to be a powerful predictor of subsequent mortality. Hypertension, obesity, and raised plasma cholesterol are considered of clinical relevance because they are potentially reversible following appropriate treatment or modifications to lifestyle. However, the reversibility of the associated risk may be less impressive: for example, treatment of hypertension results in only a partial reduction of the excess mortality. Prevention is likely to be more successful if the emergence of an adverse combination of risk factors can be avoided. Potentially remediably influences upon the development of the lung in childhood and the rate of decline of ventilatory capacity in adult non-

<table>
<thead>
<tr>
<th>Cause of death</th>
<th>Cases (n)</th>
<th>Controls (n)</th>
<th>Rate ratio (95% CI) per 10 cm decrease in height</th>
<th>$\chi^2$ (df)</th>
</tr>
</thead>
<tbody>
<tr>
<td>All causes</td>
<td>397</td>
<td>2536</td>
<td>1.0 (0.91-1.32)</td>
<td>0.94</td>
</tr>
<tr>
<td>Respiratory</td>
<td>16</td>
<td>54</td>
<td>1.33 (0.96-1.21)</td>
<td>0.32</td>
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<tr>
<td>Non-respiratory</td>
<td>381</td>
<td>2482</td>
<td>1.09 (0.90-1.31)</td>
<td>0.72</td>
</tr>
<tr>
<td>Malignant neoplasms</td>
<td>106</td>
<td>792</td>
<td>1.22 (0.87-1.72)</td>
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<tr>
<td>Circulatory</td>
<td>231</td>
<td>1364</td>
<td>0.99 (0.88-1.14)</td>
<td>0.00</td>
</tr>
<tr>
<td>Other non-respiratory causes</td>
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<td>326</td>
<td>1.22 (0.71-2.0)</td>
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<tr>
<td>Coronary heart disease</td>
<td>172</td>
<td>1068</td>
<td>1.00 (0.86-1.16)</td>
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<tr>
<td>Other circulatory disease</td>
<td>79</td>
<td>596</td>
<td>0.98 (0.76-1.28)</td>
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</table>

*Table IV* Mortality rates (per 1000 person-years) for all 3452 lifelong non-smokers by age adjusted tertile of various cardiovascular risk factors, height, FEV$_1$, and height adjusted FEV$_1$.

<table>
<thead>
<tr>
<th>Cause of death</th>
<th>Tertile of risk$^a$</th>
<th>Systolic blood pressure</th>
<th>Plasma cholesterol</th>
<th>Body mass index</th>
<th>Height</th>
<th>Height adjusted FEV$_1$</th>
<th>Unadjusted FEV$_1$ $^b$</th>
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<tbody>
<tr>
<td>All causes</td>
<td>Low risk</td>
<td>5.2</td>
<td>7.0</td>
<td>6.1</td>
<td>6.2</td>
<td>5.9</td>
<td>5.1</td>
</tr>
<tr>
<td></td>
<td>Middle</td>
<td>7.4</td>
<td>6.2</td>
<td>7.4</td>
<td>7.3</td>
<td>7.1</td>
<td>7.9</td>
</tr>
<tr>
<td></td>
<td>High risk</td>
<td>9.0</td>
<td>8.8</td>
<td>8.8</td>
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<td>9.1</td>
<td>9.2</td>
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<tr>
<td>Malignant neoplasms</td>
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<td>7.0</td>
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<td>5.9</td>
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<td>Middle</td>
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<td>7.4</td>
<td>7.3</td>
<td>7.1</td>
<td>7.9</td>
</tr>
<tr>
<td></td>
<td>High risk</td>
<td>9.0</td>
<td>8.8</td>
<td>8.8</td>
<td>8.6</td>
<td>9.1</td>
<td>9.2</td>
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<tr>
<td>Circulatory</td>
<td>Low risk</td>
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<tr>
<td>disease</td>
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<td>5.4</td>
<td>4.4</td>
<td>5.4</td>
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<td>4.5</td>
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<td></td>
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<td>6.1</td>
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<td>Low risk</td>
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<td>1.3</td>
<td>1.5</td>
<td>1.0</td>
<td>0.8</td>
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<td></td>
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<tr>
<td></td>
<td>High risk</td>
<td>1.4</td>
<td>1.4</td>
<td>1.4</td>
<td>1.4</td>
<td>1.7</td>
<td>1.9</td>
</tr>
</tbody>
</table>

$^a$Low risk groups are the lowest age adjusted tertiles of systolic blood pressure, plasma cholesterol and body mass index, and highest age adjusted tertiles of height, adjusted and unadjusted FEV$_1$.

$^b$Includes respiratory disease.
smokers may have long term consequences of considerable importance to public health.

I am grateful to Professor G Rose for encouraging me to analyse 18 mortality data from the Whitehall study, and for his helpful comments upon the paper.


