Coefficients of relationship by isonymy among registrations for five common cancers in Scottish males

Susan M Holloway, J A Sofaer

Abstract

Study objective—The aim was to assess the relative importance of genetic factors in carcinoma of the stomach, colon, rectum, prostate, and bladder in Scottish males.

Design—Cancer cases and controls were compared in terms of the coefficient of relationship by isonymy (R_i).

Setting—Surname distributions for cancer cases were derived from the Scottish Cancer Register for the years 1959–85. Control distributions were derived from all births, marriages and deaths in Scotland for 1976.

Subjects—Analysis was carried out on a total of 60 933 cancer registrations and 101 836 births, marriages, and deaths over the 12 local government regions of Scotland.

Main results—Comparisons of R_i within and between regions indicated that inherited susceptibility was of greatest importance in carcinoma of the prostate and colon, of intermediate importance in carcinoma of the rectum and stomach, and of minimal importance in carcinoma of the bladder. Familial aggregation of cancers was most pronounced in Highland, Tayside, and Borders Regions. For Highland, this appeared to be the result of region-specific familial influences, while Tayside and the Borders shared genetic factors contributing to cancer aetiology with neighbouring regions in south east Scotland.

Conclusions—Surname analysis is a simple but useful tool for studying population genetic structure and its relationship to disease incidence.

J Epidemiol Community Health 1992; 46: 368–372

example increased susceptibility to sunlight induced skin cancer in patients with the autosomal recessive disorder xeroderma pigmentosum and their heterozygous relatives.⁵ Genetic variation in the response to environmental carcinogens may also be related to metabolic polymorphisms, as shown by the increased risk of smoking induced lung cancer in extensive metabolisers of debrisoquine.⁶ At the epidemiological level, however, because common genes and common environmental exposure to carcinogens among relatives can both cause familial aggregation of malignant disease, demonstration of a hereditary component in the absence of a known disorder or genetic marker can be demanding, both in terms of study design and resources.

Surnames provide a simple approach, using data that are readily available, which may help to shed light on the genetic contribution to cancer at different sites. This approach allows estimation of the degree of genetic relationship between groups of patients with the same disease in geographically distant locations, thus reducing the confounding of common genes and common environment that plagues conventional family studies. A disadvantage is that many surnames have multiple origins, but the magnitude of any bias can be assessed by a parallel analysis using only rare surnames, each of which is more likely to have a unique (monophyletic) origin.^{7 8}

Methods

Registrations among males for carcinoma of the stomach (ICD 151), colon (ICD 153), rectum (including rectosigmoid junction and anus, ICD 154), prostate (ICD 185), and bladder (ICD 188) were obtained from the Scottish Cancer Register (1959–85). Data for each case included an identification (ID) number, surname code, date of birth, date of registration, date of death and local government region in which registration occurred. To maintain confidentiality, a key matching surname codes with surnames was held separately and used only when required. Codes for some of the earliest registrations referred to alphabetical groupings of surnames and cases with these codes were excluded. Duplicate registrations, defined as those with the same surname and ID number, or those with the same surname, date of birth, date of registration, and/or date of death, were also excluded. Females were not studied because maiden name was not generally available. As controls, surnames of all males who were born, married, or who died in each of the 12 local government regions of Scotland in 1976 were supplied by the Office of the Registrar General for Scotland. The numbers of controls and cancer

Human Genetics Unit,
University of
Edinburgh, Western
General Hospital,
Crewe Road,
Edinburgh EH4 2XU,
United Kingdom
S M Holloway
Department of Oral
Medicine and Oral
Pathology, Old
Surgeons Hall, High
School Yards,
Edinburgh, United
Kingdom
J A Sofaer

Correspondence to:
Dr Holloway

Accepted for publication
September 1991

Cancer has a spectrum of aetiologies. At one extreme are tumours such as retinoblastoma, a substantial proportion of which are hereditary.¹ At the other are largely non-hereditary tumours, particularly those associated with occupational exposure, such as mesothelioma with asbestos and angiosarcoma of the liver with vinyl chloride.² Almost all instances of carcinogenesis, however, probably involve a degree of interaction between genotype and the environment, comparable exposure to the same carcinogenic agent causing malignancy in some individuals but not in others.³ Indeed, 8% of known human genes appear to influence susceptibility or resistance to cancer.⁴

Examples of genotype-environment interaction occur among monogenic disorders known primarily for their association with neoplasia, for

Table 1 numbers of male controls (births, marriages, and deaths for 1976) and male cancer registrations with known surname (1959–85) for each of the regions of Scotland

Region	Controls	Cancer registrations				
		Stomach	Colon	Rectum	Prostate	Bladder
Highland	3826	395	598	323	586	293
Grampian	8927	1116	1329	975	1622	1076
Tayside	7524	1138	1175	737	1246	900
Fife	6671	914	754	547	1083	858
Lothian	14 327	2071	1855	1269	2578	1823
Borders	1842	258	300	239	405	247
Central	5090	586	527	392	764	642
Strathclyde	49 608	6117	5701	3738	5663	5575
Dumfries ^a	2623	329	361	242	509	280
Orkney	339	35	29	31	70	53
Shetland	430	22	50	35	63	28
Western Isles	629	86	76	42	123	54
Scotland	101 836	13 067	12 755	8570	14 712	11 829

^aDumfries and Galloway

cases used in the analysis are listed by region in table I.

Surname distribution in cancer cases and controls was compared using the coefficient of relationship by isonymy, *R_i*, which can be considered as half the probability that two people selected at random will have the same surname. The factor of $\frac{1}{2}$ is introduced so that the relationship corresponds with the expected proportion of shared autosomal genes on the assumption of monophyly of surnames. Values of *R_i* were calculated within and between groups of cancer cases and controls according to Lasker.^{9 10}

Within a group $R_i = \sum S_i(S_i - 1) / 2n(n - 1)$, where *S_i* is the number of occurrences of surname *i*, and *n* is the total number of individuals in the group, with summation over all surnames. Between groups $R_i = \sum S_{i1}S_{i2} / 2n_1n_2$, where *S_{i1}* and *S_{i2}* are the numbers of occurrences of surname *i* in groups 1 and 2, and *n₁* and *n₂* are the total numbers of individuals in groups 1 and 2, with summation over all surnames.

The significance of the difference between two *R_i* values was assessed by the method of Fox and Lasker.¹¹ These authors showed that *R_i* is effectively independent of sample size and, using empirical estimates of sampling variation for *R_i*, that *R₁* and *R₂*, two independent values of *R_i*, differ significantly at the 5% level if:

$$\frac{R_1 - R_2}{\sqrt{(\text{var } R_1 + \text{var } R_2)}} > 2$$

or

$$\frac{R_1 - R_2}{\sqrt{(R_1^2 + R_2^2)}} > \frac{1}{2}$$

Table II Coefficients of relationship by isonymy within cancers, for all surnames and rare surnames only, derived from the median difference between cancer cases and controls over the 12 regions or 66 pairs of regions and expressed as percentages of the listed median control values.

	Within regions		Between regions	
	All	Rare	All	Rare
Control births, marriages, and deaths				
Median control <i>R_i</i> ($\times 10^5$)	123	62	93	21
% Control <i>R_i</i>				
Stomach	128*	123	120†	100
Colon	142*	129	128†	110*
Rectum	123	121	117‡	119*
Prostate	126*	129†	130‡	119†
Bladder	118	102	125‡	105
Control deaths only				
Median control <i>R_i</i> ($\times 10^5$)	136	74	96	23
% Control <i>R_i</i>				
Stomach	116	100	117‡	91
Colon	115	101	118‡	100
Rectum	109	107	108‡	104
Prostate	114*	112	120‡	109
Bladder	100	76	113‡	96

Where the median difference was significantly greater than zero: **p* < 0.05, †*p* < 0.01, ‡*p* < 0.001.

Values of *R_i* were calculated for controls, within cancers, between cancers and for all cancers combined, both for Scotland as a whole and within each local government region. Values of *R_i* were also calculated for controls, within cancers and for all cancers combined between regions. Two sets of comparisons were made, first with all controls and second with control deaths only. Control deaths, which accounted for 32% of all controls, had an age distribution more similar to that of the cancer cases.

The Wilcoxon matched-pairs signed ranks test¹² was used to compare sets of the 12 within region or 66 between region *R_i* values for each cancer with controls, and sets of the 12 within region *R_i* values between cancers with controls. For each of these comparisons, the median difference (*R_i* for cancer cases minus *R_i* for controls) was calculated, added to the median control value, and the result expressed as a percentage of the median control value. To compare individual regions or pairs of regions with each other, *R_i* for all cancer cases combined was expressed simply as a percentage of the corresponding control value. All calculations were repeated using only rare surnames, considered to be those with an occurrence of less than 1 per 1000 among births, marriages, and deaths for 1976 in Scotland as a whole.¹³ Rare surnames accounted for 59% of controls and 56% of cancer cases. In expressing statistical significance, account was taken of the large number of comparisons through multiplying individual probabilities by the number of comparisons made at each stage.

Results

Calculations for Scotland as a whole, both within cancers and between cancers, did not reveal any significant increase of *R_i* in cancer cases relative to either group of controls, either for all surnames or for rare surnames only. For individual regions, only Highland showed significant differences. In this region, using rare surnames only, *R_i* values for carcinoma of the prostate, and for carcinoma of the prostate with carcinoma of the colon, were both significantly greater than *R_i* for all controls. The difference between *R_i* for carcinoma of the colon and that for all controls was bordering on statistical significance.

Within region and between region *R_i* values for the different cancers, expressed as percentages of their median control values, are given in table II. Within regions, using all controls and all surnames, the median difference between cancer cases and controls was significantly greater than zero for carcinoma of the stomach, colon, and prostate. Using rare surnames only, significance persisted for carcinoma of the prostate. Relative to control deaths only, the median difference was significantly greater than zero only for carcinoma of the prostate using all surnames. Between regions, using all surnames, the median differences for all five cancers were significantly greater than zero using either all controls or control deaths only. Significance persisted for carcinoma of the colon, rectum, and prostate relative to all controls when the analysis was restricted to rare surnames, but there were no significant differences relative to control deaths when rare surnames were used.

Table III Coefficients of relationship by isonymy between cancers within regions, for all surnames and rare surnames only, derived from the median difference from controls over the 12 regions and expressed as percentages of the median control values.

	Control births, marriages, and deaths		Control deaths only	
	All	Rare	All	Rare
Stomach—colon	121	120	108	97
Stomach—rectum	115	108	106	96
Stomach—prostate	119*	108	110	97
Stomach—bladder	116	107	95	92
Colon—rectum	132†	124	119*	105
Colon—prostate	132†	115	114	105
Colon—bladder	124	105	107	95
Rectum—prostate	132†	134†	120*	111
Rectum—bladder	119	118	107	96
Prostate—bladder	127	118	112	100

Where the median difference was significantly greater than zero: * $p < 0.05$, † $p < 0.01$.

Between-cancer R_i values within regions, expressed as percentages of the median control values, are given in table III. Using all controls and all surnames, the median differences between R_i values for pairs of cancers and for controls were significantly greater than zero for all pairs of cancers except carcinoma of the stomach and colon or rectum, and all pairs that included carcinoma of the bladder. One of the four significant differences persisted when rare surnames were used. Relative to control deaths only, two differences were significant for all surnames but none was significant for rare surnames.

Within region R_i values for all cancers combined, expressed as percentages of their control values, are given in table IV. The results for rare surnames using control deaths only, as well as all those for Orkney, Shetland, and the Western Isles, were based on small sample sizes and so were individually ignored. In none of the individual regions was there a significant difference between R_i for all cancer cases and R_i for either set of controls using all surnames or rare surnames only. However, the median difference over all 12 regions was significantly greater than zero for all cancer cases relative to all controls or control deaths only, and for cancer cases relative to all

Table IV Coefficients of relationship by isonymy within regions for all five cancers combined, for all surnames and rare surnames only, expressed as percentages of the control within region values.

	Control births, marriages, and deaths		Control deaths only	
	All	Rare	All	Rare
Highland	184	177	133	111
Grampian	121	155	99	109
Tayside	137	125	135	111
Fife	128	124	101	91
Lothian	130	112	116	94
Borders	135	126	128	92
Central	108	121	89	114
Strathclyde	126	114	113	104
Dumfries ^a	107	101	103	103
Orkney	116	113	112	93
Shetland	143	112	111	59
Western Isles	125	144	113	110

^aDumfries and Galloway

Table V Coefficients of relationship by isonymy between regions for all five cancers combined, for all surnames, and rare surnames only, expressed as percentages of the corresponding control between region values. Only the 10 pairs of regions (excluding Orkney, Shetland, and Western Isles) showing the highest percentages are listed.

Control births, marriages, and deaths		Control deaths only	
All	Rare	All	Rare
TAYS—BORD 146	GRAM—TAYS 130	TAYS—BORD 137	BORD—DUMF 147
TAYS—FIFE 137	GRAM—FIFE 130	TAYS—LOTH 130	HIGH—BORD 122
TAYS—LOTH 135	BORD—CENT 130	TAYS—STRA 126	TAYS—BORD 119
LOTH—BORD 134	GRAM—CENT 128	HIGH—TAYS 121	BORD—CENT 119
TAYS—STRA 133	GRAM—STRA 123	HIGH—BORD 121	GRAM—TAYS 117
FIFE—BORD 132	BORD—STRA 122	HIGH—DUMF 121	OOTH—DUMF 117
FIFE—LOTH 130	TAYS—FIFE 121	TAYS—FIFE 119	HIGH—TAYS 114
BORD—STRA 129	TAYS—BORD 121	TAYS—DUMF 119	GRAM—FIFE 114
GRAM—TAYS 129	LOTH—BORD 121	LOTH—BORD 119	STRA—DUMF 110
GRAM—BORD 128	GRAM—LOTH 118	LOTH—DUMF 117	GRAM—CENT 109

TAYS = Tayside; BORD = Borders; GRAM = Grampian; DUMF = Dumfries and Galloway; LOTH = Lothian; HIGH = Highland; CENT = Central; STRA = Strathclyde.

controls with rare surnames only. Highland, Tayside, and Borders showed comparatively high R_i values relative to controls for all surnames and rare surnames only, indicating that both types of surname may be associated with cancer in these regions. For Grampian, R_i was low using all surnames but high for rare surnames only, suggesting that particular rare surnames may be associated with cancer in this region.

The 10 highest R_i values between regions for all cancers combined, expressed as percentages of their control values, are shown for the four categories of analysis in table V. All regions with an individually high R_i value in table IV are represented in these pairs, although between region values involving Highland were relatively low. There was no significant difference from controls for any one pair of regions in any category. However, the median difference between cancer cases and controls over the 66 pairs of regions was significantly greater than zero for all four categories of analysis.

Discussion

Analysis was carried out for all surnames and for rare surnames, using either all controls or control deaths only. All controls provided a greater number of surnames for comparison, but mean age, taking fathers' age in the case of births, was only 40 years. Mean age for control deaths was 67 years, compared with 67 to 74 years between year of birth and 1976 for cancer patients (some of whom died before 1976). Age is of relevance because within region R_i among controls was greater for deaths than for births and marriages, as previously reported¹³ and as can be inferred from table II. Comparisons using all controls might therefore have tended to overestimate the increase in R_i shown by cancer cases. On the other hand, there was also a tendency to underestimate any difference of R_i between cancer cases and controls because about 4% of the cancer patients died in 1976 and were consequently included in the Registrar General's figures. However, for each cancer, patients who died in 1976 contributed only about 0.5% of the total number of controls and about 1.5% of control deaths. The effect of any net bias is likely to have been small but in any event should not have affected the validity of the analysis since this was aimed primarily at disclosing the relative importance of genetic factors in cancer at the different sites.

The results from rare surnames were generally similar to those based on all surnames, although rather less pronounced, despite the fact that rare surnames are expected to reflect genetic affinity more closely than all surnames, where multiple origins dilute the degree of relationship. The most likely explanation for this is that in many regions the number of cancer cases with rare surnames was small.

The lack of a significant difference between R_i values for cancer cases and controls in Scotland as a whole was probably the result of important genetic factors being associated with a much larger number of surnames in the total population than within regions, making it more difficult to assess family relationship between affected individuals.

Excluding comparisons involving both rare surnames and control deaths, where sample sizes were small, the pattern of median differences between the within region or between region *Ri* values for cancer cases and controls suggested that the five cancers fall into three different classes. For carcinoma of the colon and prostate, within region and between region differences were, respectively, 14–42% and 10–30% of their median control values. All differences were significantly greater than zero except those for within region *Ri* for carcinoma of the colon using control deaths or rare surnames. For carcinoma of the stomach and rectum, within region and between region differences were respectively 9–28% and 0–20% of their median control values. Only one of the six within region differences but five of the six between region differences were significantly greater than zero. For carcinoma of the bladder, the within region and between region differences were, respectively, 0–18% and 5–25% of their median control values. None of the within region differences and two of the three between region differences were significantly greater than zero (table II). Furthermore, *Ri* between cancers within regions disclosed some familial association between all cancers except carcinoma of the bladder (table III). Isonymy among cases of bladder cancer was therefore not only least pronounced but showed relatively little overlap with that among patients suffering from cancer at the other sites.

Within regions, a significantly higher *Ri* for cancer cases can be taken as evidence of familial factors in the aetiology of malignant disease. Between regions, it suggests that these familial factors are likely to be genetic rather than environmental, the larger number of significant comparisons reflecting the difference in size between the sets of within region and between region *Ri* values. The results therefore suggest that genetic factors have a role in the aetiology of carcinoma of the colon and prostate and, probably to a lesser extent, in carcinoma of the stomach and rectum. Inherited susceptibility appears to be least important for carcinoma of the bladder.

Cleek¹⁴ has also used surname analysis to investigate the contribution of genetic factors to cancer aetiology. For all five cancers studied here, he found some surnames which occurred more frequently than expected, the greatest number of such surnames occurring in carcinoma of the colon and prostate. Certain surnames were also associated with cancer at more than one site, in keeping with the significantly increased *Ri* between some cancers within regions found in the present study (table III).

Using a related but more rigorous approach, Hill¹⁵ was able to link 16 820 cases on the Utah Cancer Registry to the computerised Mormon genealogy records, comparing Malécot coefficients of kinship among cancer cases with those among matched controls chosen at random from the genealogy records. Kinship was greater among individuals with cancer at each of several sites, the ranking by site in order of decreasing kinship being colon, breast, gastrointestinal, and urogenital.

Several heritable syndromes are associated with a high risk of bowel cancer but these account for

only a small proportion of cases. Relatives of patients with so called sporadic disease have a two- to threefold increased risk compared with the general population¹⁶ but it is likely that diet is of particular importance. High fat ingestion is currently thought to be one of the most relevant factors, raising the possibility of inherited differences of susceptibility through genetic variation in fat metabolism.³ Familial aggregation of carcinoma of the prostate, thought to have at least a partially genetic basis, has been known for several years.^{17 18} However, since persons migrating from low risk countries to the high risk USA have shown intermediate incidence, environmental factors must also be involved.¹⁹

Familial aggregation of gastric carcinoma is rare and this tumour has fallen in incidence by 60% from 1950 to the 1980s in the USA.²⁰ This fall, together with the observation that populations migrating from high to low risk countries show intermediate incidence, demonstrates the importance of environmental influences. On the other hand, the weak but significant association between carcinoma of the stomach and blood type A, first reported more than 35 years ago,^{21 22} indicates that inherited susceptibility contributes to the disease. Autosomal dominant transmission of carcinoma of the bladder has been reported in five families but this is rare and the risk to a relative of the usual 60–70 year old patient can be considered negligible.²³ Environmental factors such as cigarette smoking and employment in the dye, rubber, leather, paint, and organic chemical industries are known to increase the risk of developing the disease.²⁴

The results of the present study are therefore in agreement with previous findings using other methods. In addition, within region *Ri* for all cancers combined indicated that familial factors were most important in Highland, Tayside, and Borders regions. In particular, in Highland Region, there was evidence for the importance of familial factors in the aetiology of carcinoma of the prostate and colon. There was also evidence that particular rare surnames may be associated with cancer in Grampian Region (table IV). Pairs of regions with high between region *Ri* (table V) were often those with individually high within region values (table IV), suggesting that pairs of regions with high values are those likely to share genetic factors important in cancer aetiology. For Highland, the high within region but less pronounced between region *Ri* values may reflect particular familial factors confined to this region.

As in the study of Hill,¹⁵ it is not possible to use the present results to draw inferences of relative risk in the families of cancer patients. It may, however, be feasible to investigate this matter further, either using empirical data from common disorders with known modes of genetic transmission or through computer simulation. Future investigations into the possible contribution of social class differences in surname distribution to isonymy among cancer cases would also seem to be indicated.

The authors are grateful to the Information and Statistics Division of the Scottish Health Service Common Services Agency and to the General Register Office for Scotland for making the data available.

- 1 Roberts DF, Aherne GES. Retinoblastoma. In: Emery AEH, Rimoin DL, eds. *Principles and practice of medical genetics*, 2nd ed. Edinburgh: Churchill Livingstone, 1990: 705–21.
- 2 Doll R, Peto R. The causes of cancer: quantitative estimates of avoidable risks of cancer in the United States today. *J Natl Cancer Inst* 1981; **66**: 1191–308.
- 3 Levine EG, King RA, Bloomfield CD. The role of heredity in cancer. *J Clin Oncol* 1989; **7**: 527–40.
- 4 Mulvihill JJ. Prospects of cancer control and prevention through genetics. *Clin Genet* 1989; **36**: 313–9.
- 5 Swift M, Chase C. Cancer in families with xeroderma pigmentosum. *J Natl Cancer Inst* 1979; **62**: 1415–26.
- 6 Ayesh R, Idle JR, Ritchie JC, Crothers MJ, Hetzel MR. Metabolic oxidation phenotypes as markers for susceptibility to lung cancer. *Nature* 1984; **312**: 169–70.
- 7 Lasker GW. Relationships among the Otmoor villages and surrounding communities as inferred from surnames contained in the current register of electors. *Ann Hum Biol* 1978; **5**: 105–11.
- 8 Lasker GW, Mascie-Taylor CGN. Surnames in five English villages: relationship to each other, to surrounding areas and to England and Wales. *J Biosoc Sci* 1983; **15**: 25–34.
- 9 Lasker GW. A coefficient of relationship by isonymy: a method for estimating the genetic relationship between populations. *Hum Biol* 1977; **49**: 489–93.
- 10 Lasker GW. *Surnames and genetic structure*. Cambridge: Cambridge University Press, 1985.
- 11 Fox WR, Lasker GW. The distribution of surname frequencies. *Int Stat Rev* 1983; **51**: 81–7.
- 12 Siegel S. *Nonparametric statistics*. New York: McGrawHill, 1956: 75–83.
- 13 Holloway SM, Sofaer JA. Coefficients of relationship by isonymy within and between the regions of Scotland. *Hum Biol* 1989; **61**: 87–97.
- 14 Cleck RK. Surnames and cancer genes. *Hum Biol* 1989; **61**: 195–211.
- 15 Hill JR. A survey of cancer sites by kinship in the Utah Mormon population. In: Cairns J, Lyon JL, Skolnick M, eds. *Banbury Report 4, Cancer incidence in defined populations*. New York: Cold Spring Harbor Laboratory, 1980: 299–318.
- 16 Cohen AM, Shank B, Friedman MA. Colorectal cancer. In: DeVita VT, Hellman S, Rosenberg SA, eds. *Cancer—principles and practice of oncology*, 3rd edn. Philadelphia: JB Lippincott, 1989: 895–964.
- 17 Morganti G, Gianferrari L, Cresseri A, Arrigoni G, Lovati G. Recherches clinico-statistiques et génétiques sur les néoplasies de la prostate. *Acta Genet Stat Med* 1956; **6**: 304–5.
- 18 Woolf CM. An investigation of the familial aspects of carcinoma of the prostate. *Cancer* 1960; **13**: 739–44.
- 19 Perez CA, Fair WR, Ihde DC. Carcinoma of the prostate. In: DeVita VT, Hellman S, Rosenberg SA, eds. *Cancer—principles and practice of oncology*, 3rd ed. Philadelphia: JB Lippincott, 1989: 1023–58.
- 20 Macdonald JS, Steele G, Gunderson LL. Cancer of the stomach. In: DeVita VT, Hellman S, Rosenberg SA, eds. *Cancer—principles and practice of oncology*, 3rd ed, Philadelphia: JB Lippincott, 1989: 765–99.
- 21 Aird I, Bentall HH, Roberts JAF. A relationship between cancer of the stomach and the ABO blood groups. *BMJ* 1953; **i**: 799–801.
- 22 Mourant AE, Kopec AC, Domaniewska-Sobczak K. *Blood groups and diseases*. Oxford: Oxford University Press, 1978.
- 23 Schimke RN. Dominant inheritance in human cancer. In: Chaganti RSK, German J, eds. *Genetics in clinical oncology*. Oxford: Oxford University Press, 1985: 103–21.
- 24 Richie JP, Shipley WU, Yagoda A. Cancer of the bladder. In: DeVita VT, Hellman S, Rosenberg SA, eds. *Cancer—principles and practice of oncology*, 3rd ed. Philadelphia: JB Lippincott, 1989: 1008–22.