

# Excess mortality due to indirect health effects of the 2011 triple disaster in Fukushima, Japan: a retrospective observational study

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## ABSTRACT

**Background** Evidence on the indirect health impacts of disasters is limited. We assessed the excess mortality risk associated with the indirect health impacts of the 2011 triple disaster (earthquake, tsunami and nuclear disaster) in Fukushima, Japan.

**Methods** The mortality rates in Soma and Minamisoma cities in Fukushima from 2006 to 2015 were calculated using vital statistics and resident registrations. We investigated the excess mortality risk, defined as the increased mortality risk between postdisaster and predisaster after excluding direct deaths attributed to the physical force of the disaster. Multivariate Poisson regression models were used to estimate the relative risk (RR) of mortality after adjusting for city, age and year.

**Results** There were 6163 and 6125 predisaster and postdisaster deaths, respectively. The postdisaster mortality risk was significantly higher in the first month following the disaster (March 2011) than in the same month during the predisaster period (March 2006–2010). RRs among men and women were 2.64 (95% CI 2.16 to 3.24) and 2.46 (95% CI 1.99 to 3.03), respectively, demonstrating excess mortality risk due to the indirect health effects of the disaster. Age-specific subgroup analyses revealed a significantly higher mortality risk in women aged ≥85 years in the third month of the disaster compared with predisaster baseline, with an RR (95% CI) of 1.73 (1.23 to 2.44).

**Conclusions** Indirect health impacts are most severe in the first month of the disaster. Early public health support, especially for the elderly, can be an important factor for reducing the indirect health effects of a disaster.

## INTRODUCTION

Unprecedented ageing societies, coupled with increasing disaster risks due to climate change and political and economic instability,<sup>1,2</sup> make the management of disaster risk a public health issue in terms of prevention, preparation, response, recovery and reconstruction.<sup>3</sup> This is because the elderly are particularly susceptible to health risks in disaster settings.<sup>4</sup> For example, the elderly often face a higher mortality risk than younger populations in disaster settings, as seen in the Hanshin-Awaji earthquake and the Chicago heat wave, both in 1995.<sup>5</sup> In order to develop effective design and delivery of public health countermeasures against

the health impact of disasters in ageing populations, it is essential to understand actual health impacts and their potential risk factors.<sup>6</sup>

The Great East Japan Earthquake, ensuing tsunami and nuclear disaster (hereinafter referred to as the triple disaster), which occurred on 11 March 2011, is a powerful example of a disaster that occurred in a rapidly ageing society. Residents of Soma and Minamisoma cities, located around 10–45 km north of the Fukushima Daiichi Nuclear Power Plant (see online supplementary figure S1), experienced substantial health consequences from the triple disaster. More than 1000 residents died from direct injuries from the earthquake and tsunami. The disaster also inflicted serious damage to community healthcare systems, resulting in various health consequences. For example, mass evacuation following the disaster strained essential health services and infrastructure and disrupted social capital and networks.<sup>7</sup>

After the nuclear disaster, mandatory evacuation zones were issued on 12 March 2011 for the 20 km radius surrounding the plant, and voluntary evacuation zones were additionally put into place for the 20–30 km radius. As a result, the population of the study area dropped from its predisaster level of nearly 100 000 to 40 000 within a month of the triple disaster.<sup>8</sup> Five of the eight emergency hospitals in the district closed after the disaster because of forced evacuation or workforce shortage.<sup>7</sup>

The health effects of natural and man-made disasters can be both direct and indirect.<sup>9</sup> The direct health effects of a disaster are defined as external causes such as traumas and burns that can be attributed directly to the disaster.<sup>10</sup> Elderly people are at greater risk of increased direct mortality after a disaster.<sup>11</sup> In the case of Hurricane Katrina in 2005 in the USA, approximately 75% of all direct mortality in New Orleans occurred among the elderly, who represented only 15% of the city's total population before the disaster.<sup>12</sup> The high vulnerability of the elderly to direct disaster impacts could be due to their impaired physical mobility, diminished sensory awareness and pre-existing health conditions.<sup>13</sup> It is equally important to recognise indirect health effects. Major disasters damage primary healthcare systems and food and water supplies, and cause displacement by evacuation and relocation.<sup>14,15</sup> Previous studies have reported indirect health effects such as increases in the mortality



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risk of cardiovascular diseases<sup>16</sup> or infectious diseases<sup>17</sup> after disasters. A report on Hurricane Katrina suggested that indirect psychological health effects of disaster might be implicated in the persistent increase in mortality observed for at least 6 months after Hurricane Katrina in 2006.<sup>18</sup> Elderly people can also be at greater risk of indirect health effects of disasters due to disruption to ongoing medical care or aggravation of medical conditions; for example, older age was reported to be a risk factor of the disruption of access to medical care after Typhoon Haiyan in 2013.<sup>19</sup> Disasters can, therefore, exert powerful influences on individual vulnerability to physical and psychological stress, and may also impact health.<sup>20</sup>

However, evidence on the indirect impacts of disasters is limited,<sup>21</sup> partly because they are, by definition, not directly measurable and partly because it is difficult to prove an indirect causal relationship between postdisaster mortality and morbidity as time passes. The major objectives of the present study were twofold: (1) to assess the excess mortality risk due to the indirect health impacts of the triple disaster in Soma and Minamisoma cities in Fukushima; and (2) to identify the postdisaster risk periods over a multiyear follow-up period and populations at risk of the indirect health effects.

## METHODS

### Data sources

We used vital registration records from a national survey between January 2006 and December 2015 (predisaster: 2006–2010; postdisaster: 2011–2015). Secondary use of the vital registration records was approved by the Ministry of Health, Labour and Welfare in compliance with the Statistics Act. Mortality data were classified by age, sex and cause of death in the International Classification of Diseases and Health-Related Problems, 10th Revision (ICD-10), for each month during the study period. Population data were obtained from the Basic Resident Registers—the nationwide resident-registry network administered by municipality unit (city/town/village).

### Sample selection

We defined a ‘direct death’ as a death attributed to external causes, primarily the physical forces of the triple disaster.<sup>22 23</sup> We considered the following causes of death that occurred in March 2011 to be direct deaths due to the disaster: S00–S09 (injuries to the head), S12 (fractures of neck), S20–S29 (injuries to the thorax), T00–T07 (injuries involving multiple body regions), T14 (injuries of unspecified body region), T29 (burns and corrosions of multiple body regions), T71 (asphyxiation) and T79 (certain early complications of trauma). We additionally considered T75.1 (drowning) that occurred on the disaster date as direct deaths. Our data included no mortality due to exposure to forces of nature (X30–X39).

### Inclusion and exclusion criteria

All mortality data from Soma City and Minamisoma City, Fukushima Prefecture, Japan, in 2006–2015 were eligible for inclusion in the study. In order to investigate the indirect health effects of the triple disaster, we compared the excess mortality risk postdisaster with predisaster mortality levels after excluding direct deaths.

### Data analysis

Age-adjusted mortality rates from 2006 to 2015 were calculated separately for men and women using the 1985 Japanese model population (see online supplementary table S1).<sup>24</sup> We then

compared postdisaster age-adjusted mortality rates (2011–2015) and the average predisaster age-adjusted mortality rates between 2006 and 2010 using  $X^2$  tests.

In order to assess the temporal trends in excess mortality risk, we compared mortality rates in each month postdisaster (2011–2015) with the average predisaster mortality rate (2006–2010) using a Poisson regression model with age (0–64, 65–74, 75–84 and 85+ years), city (Soma City or Minamisoma City) and year (baseline (2006–2010), 2011, 2012, 2013, 2014 and 2015) as covariates. In order to reduce any seasonal/monthly effect on the data, we constructed regression models for each month (January–December) separately. Our major interests were comparisons between postdisaster years (2011–2015) versus the predisaster baseline with respect to the variable ‘year’. Baseline risk was defined as the monthly mortality rate per year from 2006 to 2010, which was also calculated for each month per year (eg, every January across the 5 years). In this regression, to avoid 0 or small numbers of death counts for each age group in each month, the original age groups at 5-year intervals were reclassified into four age groups (0–64, 65–74, 75–84 and 85+ years). For a detailed explanation of this approach, see online supplementary text S1.

We performed the cause of death-specific analysis using the same models to evaluate the potential difference in the post-disaster mortality trends by cause of death. We considered the four leading causes of death identified in the first month of the disaster (March 2011). We also separately constructed a regression model for each age group using similar methods to assess how monthly trends in potential postdisaster excess mortality risk differed by age.

We used Stata/MP V.14.1 for all analyses. P Values less than 0.05 were considered statistically significant. R V.3.30 was used to generate graphics. This study was approved by the Ethics Board of Soma Central Hospital (ID: 2016-A1).

## RESULTS

### Sample characteristics

Of the total number of deaths in the predisaster period ( $n=6163$ ) and in the postdisaster period ( $n=7215$ ), we excluded 1090 deaths in 2011 as direct deaths. As a result, 6163 (women: 2938, 48%) deaths in the predisaster period and 6125 (women: 2953, 48%) deaths in the postdisaster period were included for further analysis. The numbers of indirect and direct deaths by age (0–39, 40–64, 65–74, 75–84, 85+), sex (male or female) and year (2006–2015) are shown in [table 1](#).

### Trends in age-adjusted mortality rates

[Figure 1](#) illustrates the yearly trends in age-adjusted total mortality rate per 100 000 from 2006 to 2015 for men and women. Comparisons of postdisaster mortality rates with that in the predisaster baseline period (2006–2010) demonstrated a statistically significant decline in 2014 (598 vs 493,  $p<0.01$ ) and 2015 (598 vs 528,  $p=0.04$ ) in men and in 2015 in women (300 vs 247,  $p=0.03$ ).

### Predisaster and postdisaster comparison of monthly mortality risks

[Figure 2](#) shows the estimated relative risks (RRs) of mortality postdisaster in comparison with the predisaster baseline, which were based on month-specific regression models. After adjusting for city and age, the postdisaster mortality risk for both men and women was significantly higher in the first month of the disaster (March 2011) than that in the same month during

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**Table 1** Age-specific, sex-specific and year-specific counts of direct and indirect deaths in the predisaster and postdisaster periods

| Age   | Predisaster period (male) |      |      |      |      |       | Postdisaster period (male) |      |      |      |      | Total indirect deaths | Total direct deaths |
|-------|---------------------------|------|------|------|------|-------|----------------------------|------|------|------|------|-----------------------|---------------------|
|       | 2006                      | 2007 | 2008 | 2009 | 2010 | Total | 2011                       | 2012 | 2013 | 2014 | 2015 |                       |                     |
| 0–39  | 16                        | 11   | 19   | 19   | 10   | 75    | 11                         | 18   | 10   | 6    | 7    | 52                    | 88                  |
| 40–64 | 91                        | 97   | 112  | 96   | 116  | 512   | 89                         | 92   | 86   | 78   | 95   | 440                   | 171                 |
| 65–74 | 120                       | 142  | 140  | 130  | 133  | 665   | 96                         | 114  | 123  | 109  | 142  | 584                   | 104                 |
| 75–84 | 222                       | 227  | 225  | 207  | 243  | 1124  | 217                        | 220  | 216  | 213  | 204  | 1070                  | 121                 |
| 85–   | 149                       | 158  | 168  | 190  | 184  | 849   | 194                        | 167  | 217  | 210  | 238  | 1026                  | 38                  |

| Age   | Predisaster period (female) |      |      |      |      |       | Postdisaster period (female) |      |      |      |      | Total indirect deaths | Total direct deaths |
|-------|-----------------------------|------|------|------|------|-------|------------------------------|------|------|------|------|-----------------------|---------------------|
|       | 2006                        | 2007 | 2008 | 2009 | 2010 | Total | 2011                         | 2012 | 2013 | 2014 | 2015 |                       |                     |
| 0–39  | 13                          | 8    | 10   | 12   | 8    | 51    | 3                            | 5    | 7    | 5    | 4    | 24                    | 79                  |
| 40–64 | 48                          | 40   | 40   | 47   | 42   | 217   | 36                           | 27   | 42   | 51   | 36   | 192                   | 151                 |
| 65–74 | 60                          | 62   | 64   | 68   | 48   | 302   | 38                           | 62   | 40   | 58   | 48   | 246                   | 126                 |
| 75–84 | 168                         | 187  | 172  | 157  | 182  | 866   | 158                          | 164  | 151  | 134  | 140  | 747                   | 142                 |
| 85–   | 275                         | 297  | 323  | 272  | 335  | 1502  | 352                          | 307  | 333  | 355  | 397  | 1744                  | 70                  |

The counts of direct deaths are shown in the far right column of the table. All direct deaths were observed in 2011.

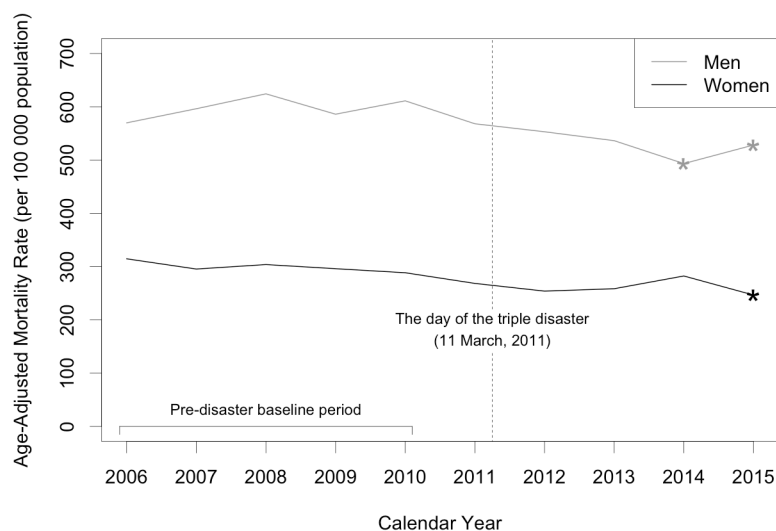
the predisaster period (March 2006–2010): RR: 2.64, 95% CI 2.16 to 3.24; RR: 2.46, 95% CI 1.99 to 3.03, respectively. This finding indicates the excess mortality risk due to the indirect effects of the disaster in the first month. The mortality risk dramatically declined from the second month postdisaster (April 2011), and statistical significance was observed in the 10th and 11th month of the disaster in both men and women. The estimated RRs for the first 12 months of the disaster are shown in online supplementary table S1.

### Cause-specific mortality trends

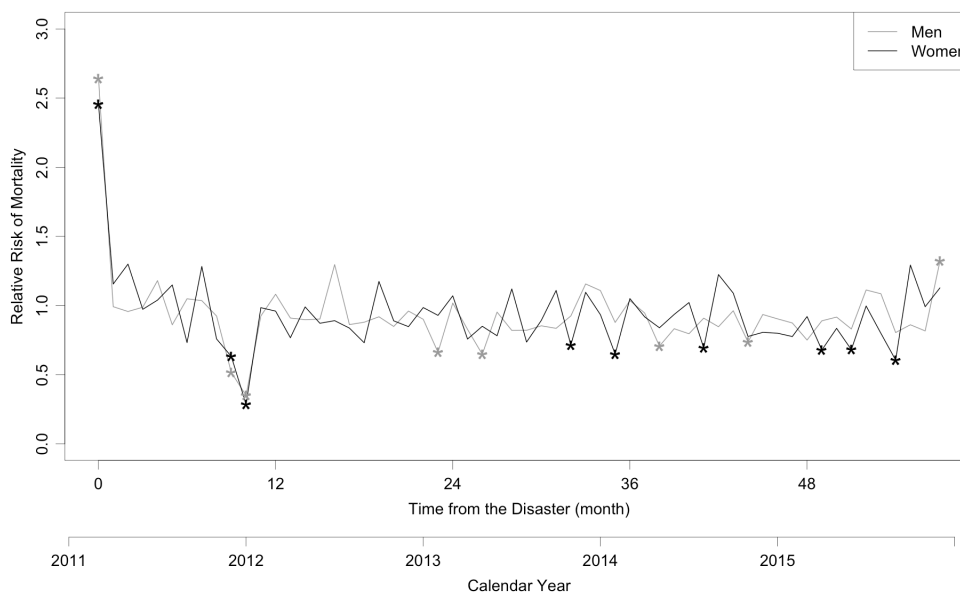
In the first month of the disaster (March 2011), when excess mortality risk was identified, the four leading causes of deaths were pneumonia (ICD-10: J12–18, 28% of the overall mortality; n=47), stroke (ICD-10: I60–I69, 15%; n=25), coronary heart disease (ICD-10: I21–I25, 10%; n=16) and cancer (ICD-10:

C00–97, 9%; n=15). In the same month of the predisaster baseline period (March 2006–2010), mortality due to pneumonia, stroke, coronary heart diseases and cancer accounted for 16% (n=1049), 14% (n=877), 9% (n=580) and 26% (n=1659) of the overall mortality, respectively. As for pneumonia, we investigated medical records in the study area and obtained information on 23 of the 47 patients who died from pneumonia. Among them, 15 (65%) were hospitalised or living in nursing care facilities as of 11 March 2011.

Figure 3 shows the cause-specific trends in the RR of mortality during the postdisaster period. Mortality due to pneumonia, stroke, coronary heart disease and cancer were all significantly higher in the first month of the disaster (March 2011) than in the same month of the predisaster period (March 2006–2010). Estimated cause-specific RRs for the first 12 months from the disaster are presented in online supplementary table S3.



**Figure 1** Trends in age-adjusted mortality rates (direct and indirect). The age-adjusted mortality rates after 2011 were compared with those in the control period (2006–2010) using  $\chi^2$  tests. The asterisk (\*) indicates statistical significance at the 0.05 level compared with the control period.



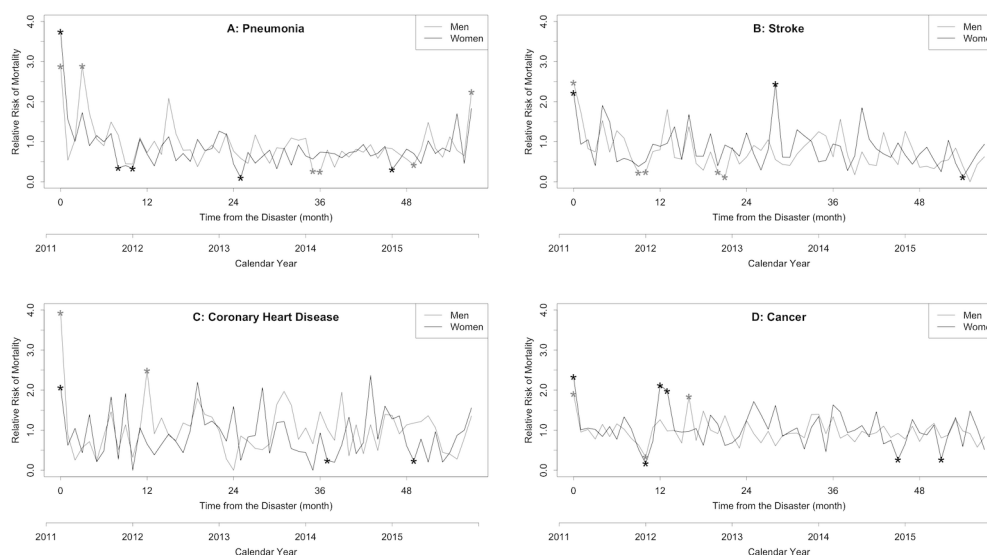
**Figure 2** Monthly trends of the relative risks of indirect mortality predisaster and postdisaster. Baseline risk was defined as the monthly mortality rate per year from 2006 to 2010. The relative risks were adjusted for city and age, and the asterisk (\*) indicates statistical significance at the 0.05 level.

### Age-specific predisaster and postdisaster comparison of monthly mortality risks

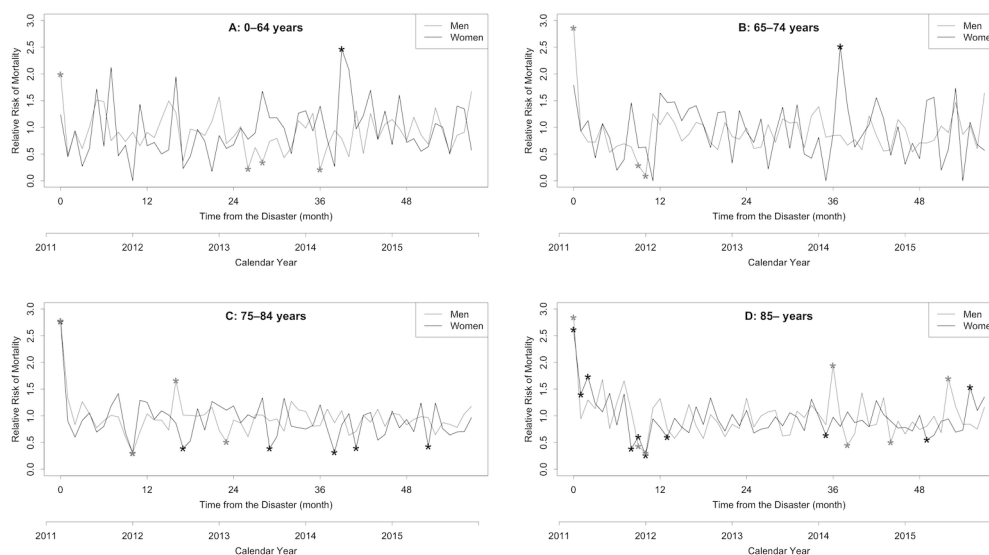
The age-specific RRs of mortality during the postdisaster period compared with the predisaster baseline are plotted in [figure 4](#). For both sexes, in the first month of the disaster (March 2011), the mortality risks in all age groups were significantly higher than those in the same month of the predisaster baseline period (March 2006–2010). Among women aged 85 years and over, significantly higher mortality risk was also identified in the third month of the disaster (May 2011). The estimated RRs in each age group for the first 12 months of the disaster are shown in online supplementary table S4.

### DISCUSSION

To our knowledge, this is the first study to comprehensively estimate the short-term and long-term indirect health impacts of the triple disaster (earthquake, tsunami and nuclear disaster) in Fukushima through the measurement of excess mortality risk. The results of this study showed that postdisaster excess mortality risk was greatest in the first month of the 5-year postdisaster period. This mortality increase may be due to the indirect health effects of the disaster. Previous studies reported increased burdens of chronic diseases and psychological disorders after the disaster in Fukushima, which are consistent with our findings.<sup>25 26</sup> In order to mitigate such indirect health



**Figure 3** Cause-specific monthly trend of the relative risks of indirect mortality predisaster and postdisaster: (A) pneumonia, (B) stroke, (C) coronary heart disease and (D) cancer. Baseline risk was defined as the monthly mortality rate per year from 2006 to 2010. The relative risks were adjusted for city and age, and the asterisk (\*) indicates statistical significance at the 0.05 level.



**Figure 4** Age-specific monthly trends of the relative risks of indirect mortality predisaster and postdisaster: (A) 0–64 years, (B) 65–74 years, (C) 75–84 years and (D) 85+ years. Baseline risk was defined as the monthly mortality rate per year from 2006 to 2010. The relative risks were adjusted for city, and the asterisk (\*) indicates statistical significance at the 0.05 level.

effects of disasters, initial public health support, such as facilitating community development, restoring livelihoods, providing psychosocial support to inform people and mobilising community partnership, is important.<sup>27–29</sup> This study also suggested the importance of initial public health support, particularly within the first month postdisaster, in reducing indirect health effects of a disaster.

The increase in mortality risks persisted for 3 months after the disaster among women 85 years or older (figure 4). This persistent increase in mortality risk is consistent with that observed after Hurricane Katrina.<sup>18</sup> Our findings suggest that mortality risk due to the indirect health effects of the disaster may be more severe and persistent among the elderly than among younger groups. This larger effect among the elderly may be explained by their increased risks of impaired physical mobility, diminished sensory awareness, poor chronic health conditions, and social and economic limitations,<sup>30</sup> which may hinder their adaptability after disasters.<sup>13</sup> Several studies reported increased health burdens among the elderly associated with the indirect health effects of a disaster.<sup>31–32</sup> Our findings suggest that the public health support to mitigate the indirect health effects of disasters may be more important among the elderly than in the general population.

The four leading causes of death included pneumonia, coronary heart disease, stroke and cancer. These increases in mortality risk after disasters are consistent with past studies.<sup>33–34</sup> Patients with chronic conditions such as cancer and those with acute health conditions such as pneumonia, coronary heart disease and stroke are dependent on medical care, which may explain their vulnerability to the impacts of a disaster, including the loss of access to medical care, environmental change, and physical and psychological stress.

In particular, pneumonia accounted for 28% (47/165) of the deaths during the first month after the earthquake. A previous study hypothesised that postdisaster increases in pneumonia were caused by the swallowing of dust and water carried by the tsunami.<sup>33</sup> However, in our findings, more than half of the patients who died from pneumonia were in inpatient medical facilities at the time of the earthquake, making them unlikely

to be directly affected by the tsunami. A previous study on patients with pneumonia after the disaster suggested that most cases of pneumonia were aspiration pneumonia acquired in a nursing home, which was associated with low patient activities of daily living status.<sup>34</sup> Deterioration of primary care after disasters makes it difficult to provide adequate oral care in order to reduce the mortality rate of aspiration pneumonia in such cases.<sup>35</sup>

In contrast to the first month after the disaster, there were no significant increases in the mortality risks in the second month and thereafter, although previous disaster studies reported exacerbation of chronic diseases in affected areas.<sup>25–26</sup> There are a few possible explanations. First, temporary housing was provided to reduce the mortality risk due to psychological stress<sup>20</sup> and social isolation<sup>36</sup> after the triple disaster: housing assignments in temporary housing were based on pre-earthquake addresses to maintain a sense of community. Second, health facilities were at least partially restored in the affected areas by the second postdisaster month and onwards. These elements could have mitigated the expected increase in mortality by exacerbation of chronic diseases.<sup>37</sup> For effective long-term support after disasters, further monitoring the effectiveness of individual measures will be essential.

It is also of note that the increased mortality risks did not persist over the 5 years since the triple disaster. The series of disasters, particularly the nuclear disaster at Fukushima Daiichi Nuclear Power Plant, has led to serious concerns about health problems among residents in the affected area. Although past studies showed that radiation exposure among residents of the study area was not high enough to cause health problems,<sup>38–39</sup> fear for health damage due to radiation exposure remains.<sup>40</sup> There was no increase in cancer mortality rates observed in this study; however, the 5-year study period is likely not long enough to assess the potential carcinogenic effects of radiation.<sup>41</sup> Therefore, we are not able to conclude that cancer incidence due to radiation exposure has not increased in this area. A long-term analysis of the exposed population cohort will be required in order to evaluate the effects of radiation exposure on carcinogenesis.



## Limitations

This study has several limitations. First, the estimated indirect mortality risk in the latter period should be interpreted with caution. As shown in [figure 1](#), the age-adjusted mortality rate has been decreasing throughout the study period. This decreasing trend, in line with age-adjusted mortality rates in Japan as a whole,<sup>42</sup> might be explained by the reduction of cancer mortality, new technologies or improved management of chronic diseases.<sup>43 44</sup> The long-term indirect health effects of the disaster on mortality risk might be counterbalanced by the decreasing mortality trends. Second, in the present study, mortality data were based on the Basic Resident Registers, and residents who moved their registrations to outside the study area were excluded. Therefore, the estimation of mortality risk in this study may be biased due to sampling limitations. Finally, the causes of death listed in the vital registration are occasionally inaccurate and incomplete, particularly in the ageing population.<sup>45</sup> In the present study, we attempted to minimise such bias by checking medical records and other sources.

## CONCLUSION

In the 5 years since the 2011 triple disaster in Japan, mortality risk, excluding direct health damage from the disaster, was highest in the first month after the disaster. This excess risk of death is attributed to the indirect health impacts of the triple disaster. The increased mortality was more severe and long-lasting among the elderly. Early public health support for the elderly, including facilitating community development, restoring livelihoods and providing psychosocial support, preferably within a month of a disaster, is one of the most important factors in reducing the indirect health effects of a disaster.

### What this study adds

- ▶ We clarified that the postdisaster mortality risk was significantly higher in the first month of the disaster compared with that during the predisaster period (2006–2010), which suggests excess mortality risk due to the indirect health effects of the disaster.
- ▶ The indirect health effects of a disaster may be more severe and persistent among the elderly than among younger groups. Early public health support, especially for the elderly, such as facilitating community development, restoring livelihoods, providing psychosocial support to inform people and mobilising community partnerships, is an important factor for reducing the indirect health effects of a disaster.

### What is already known on this subject

- ▶ Although the health effects of natural and man-made disasters can be both direct and indirect, there is little evidence on the indirect health effects of a disaster.
- ▶ The 2011 triple disaster (earthquake, tsunami and nuclear disaster) in Fukushima resulted in the opportunity to assess excess mortality risk due to the indirect health effects of the disaster.

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**Patient consent** Detail has been removed from this case description/these case descriptions to ensure anonymity. The editors and reviewers have seen the detailed information available and are satisfied that the information backs up the case the authors are making.

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## Excess mortality due to indirect health effects of the 2011 triple disaster in Fukushima, Japan: a retrospective observational study

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