Antenatal screening: its use in assessing obstetric risk factors in Zimbabwe

Vivien D Tsu

Abstract

Study objective – To assess the predictive utility of obstetric risk factors for identifying those women at high risk of obstetric complications in a developing world setting, where home deliveries predominate and emergency transport is scarce.

Design – Risk factors derived from two population-based, case-control studies (one of cephalopelvic disproportion and one of post partum haemorrhage), carried out in Zimbabwe were used to construct weighted and unweighted scores, a variety of screening algorithms, and sets of probabilities estimated from logistic regression models. These screening tests were evaluated for sensitivity, specificity, positive predictive value, and “cost” (the proportion of the population testing positive). Each complication was evaluated separately and the two were then pooled.

Participants – All were Harare residents with singleton, vertex deliveries and spontaneous onset of labour. A total of 201 experienced cephalopelvic disproportion, 150 had post partum haemorrhage, and 299 had normal, unassisted deliveries.

Measurements and main results – Largely because of the very low incidence of the two complications studied (1% or less), positive predictive values were low (less than 7%). Holding “cost” constant at 10%, a screening test for cephalopelvic disproportion could predict 42–53% of cases compared with only 35–60% of those with post partum haemorrhage. Weighted scores had little advantage over unweighted ones, and probabilities from the logistic regression models did not differentiate cases from controls very well.

Conclusions – With simple algorithms based on maternal height, parity, and obstetric history, more than one third of women at risk for potentially fatal complications could be identified at relatively small cost to themselves or the health care system.

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One of the primary goals of antenatal care is to identify those women with raised risks for problems during pregnancy or delivery, in order to ensure that precautionary measures are instituted where possible or more intensive medical care is available should it be needed.
quantify as are the consequences of being incorrectly classified.

The present study was undertaken in an effort to predict risk for two well defined complications for which the interventions have already been proved relatively effective, post partum haemorrhage and cephalopelvic disproportion. A site was chosen (Harare, Zimbabwe) that permitted nearly 90% of all local births to be represented in the study, while still being able to include data on a wide range of pertinent variables, data that were collected before the outcome was known.

Methods
Risk factors for the two complications were identified through a multivariate logistic regression analysis of data from Harare, Zimbabwe. The data came from two case-control studies involving 203 women with cephalopelvic disproportion, 151 women with post partum haemorrhage, and 299 women with normal vaginal deliveries. Details on the subjects and data collection and analysis are reported elsewhere. Briefly, all cases and controls were women resident in Harare with singleton, vertex deliveries and spontaneous onset of labour.

Cephalopelvic disproportion cases included eligible women delivering in Harare during eight months in 1989 with an operative delivery for cephalopelvic disproportion. In addition to meeting a strict set of criteria designed to ensure that cephalopelvic disproportion was the primary cause of the operative delivery, eligible cases were rated as "definite" or "probable" by at least two of a panel of three senior obstetricians who independently reviewed the abstracted case records. Post partum haemorrhage cases occurred in the same period and included all eligible women with a recorded post partum blood loss of at least 600 ml after a normal, unassisted vaginal delivery. Controls were matched to cases by the facility where the case originally was booked to deliver and by the week of the delivery. Case data were abstracted from the woman's medical record while she was still in the hospital, while data for controls were abstracted within a few weeks after delivery from the uniform maternal record form used for all deliveries in the public system.

To evaluate the effectiveness of various factors in screening pregnant women for hospital rather than home delivery, only factors detectable before the onset of labour were used. Risk factors identified for cephalopelvic disproportion included advanced maternal age (35 years and older), short stature (less than 160 cm), nulliparity, a history of operative delivery or neonatal death in the preceding pregnancy, and admission to hospital for pregnancy induced hypertension in the current pregnancy. Those identified for post partum haemorrhage also included advanced maternal age and low parity, but included a more general history of poor obstetric outcome in the preceding pregnancy, anaemia during the current pregnancy, and admission to hospital for any pregnancy related problem before the onset of labour as well. Table 1 lists the factors and their adjusted relative risk (RR) estimates from the two logistical models.

Several approaches were undertaken to devise screening tests using these factors and to evaluate the resulting systems. Initially, the logistical models were used to calculate estimated probabilities for each subject for the two separate outcomes (see Appendix for an explanation of the computations). In addition, a variety of different algorithms were developed using cumulative scores based on the sum of all risk factors, either unweighted or weighted by their RR values, or simple tests based on the presence of one or more selected risk factors, alone or in combination with each other. Screening tests were devised for each of the two complications separately and, since a good screening system should identify women at risk for either of these problems, for a pooled population consisting of both case groups and the controls. Factors used for screening the pooled population were derived from those elements the two case studies had in common.

Sensitivity, specificity, and likelihood ratios (the ratio of sensitivity to 1-specificity) were calculated for the dichotomous algorithms and for various cut off points for the continuous scores and the probabilities. All logical combinations of factors were tested, but only those with higher likelihood ratios were reported here. Receiver operating characteristic (ROC) curves, plots of the true positive rate (sensitivity) against the false positive rate (1-specificity), were derived for the probabilities estimated from the logistic models and for the weighted and unweighted cumulative scores.

Predictive values for a test could not be derived from this case-control study directly, but were calculated using the estimated incidences of cephalopelvic disproportion and post partum haemorrhage in this population. Only likelihood ratios and predictive values for positive tests are considered here, since negative tests in this particular setting will result in no change in the course of action planned before the test (that is, either home or hospital deliv-

Table 1 Risk factors and their relative (RR) estimates from the logistic regression models

<table>
<thead>
<tr>
<th>Risk factor</th>
<th>Estimate of the adjusted RR</th>
<th>95% confidence interval</th>
</tr>
</thead>
<tbody>
<tr>
<td>Cephalopelvic disproportion:</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age 35+</td>
<td>2.1 (0.96, 4.4)</td>
<td></td>
</tr>
<tr>
<td>Height &lt; 160 cm</td>
<td>2.0 (1.3, 3.0)</td>
<td></td>
</tr>
<tr>
<td>Nulliparous</td>
<td>13.8 (7.3, 25.9)</td>
<td></td>
</tr>
<tr>
<td>Operative delivery in last pregnancy</td>
<td>9.5 (3.8, 23.6)</td>
<td></td>
</tr>
<tr>
<td>Neonatal death in last pregnancy</td>
<td>4.5 (1.3, 15.9)</td>
<td></td>
</tr>
<tr>
<td>Antenatal hospitalisation for PIH</td>
<td>4.9 (1.6, 28.3)</td>
<td></td>
</tr>
<tr>
<td>Post partum haemorrhage:</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age 35+</td>
<td>2.6 (1.2, 5.8)</td>
<td></td>
</tr>
<tr>
<td>Low parity (0, 1)</td>
<td>1.7 (1.1, 2.7)</td>
<td></td>
</tr>
<tr>
<td>Poor obstetric outcome last pregnancy†</td>
<td>1.8 (0.94, 3.7)</td>
<td></td>
</tr>
<tr>
<td>Antenatal haemoglobin &lt; 12 g/l</td>
<td>2.2 (0.99, 5.0)</td>
<td></td>
</tr>
<tr>
<td>Antenatal hospitalisation for pregnancy related problem</td>
<td>4.3 (1.4, 12.8)</td>
<td></td>
</tr>
</tbody>
</table>

* Adjusted for the other factors listed and for facility booked (on which cases and controls were matched).
† Post partum haemorrhage, 1st or 2nd trimester, fetal death, stillbirth, neonatal death.
PIH = pregnancy induced hypertension.
How useful are obstetric risk factors?

The proportion of the population that would be referred according to a particular test constitutes an important measure of the "cost" of the test to the system. Like the predictive values, it was calculated using the estimated incidences of cephalopelvic disproportion and post partum haemorrhage in this population.

Results

Scores for cases and controls, based on sums of the risk factors, either unweighted or weighted by the RR values, were compared using ROC curves. For cephalopelvic disproportion (see fig 1), the score based on weighted factors performed modestly better than one using unweighted factors, but even at its optimum point (the point on the curve closest to the upper left corner), that is using a cut off of 4 or more, the weighted score had both a sensitivity and specificity of only 75%. The curves for post partum haemorrhage (fig 2) were only a bit above the diagonal line representing results that could be expected by chance if the test were simply random, with scores based on weighted factors showing little advantage over unweighted ones.

The cumulative scores and dichotomous algorithms were compared on the basis of sensitivity, specificity, likelihood ratio (LR), positive predictive value (PV+), and "cost" (total proportion of the population referred). Table 2 shows the comparative utility of systems for cephalopelvic disproportion and post partum haemorrhage separately, in ascending order of cost. Because of the low incidences of these two conditions in this population, the positive predictive values are quite low, with none above 7%. Focussing on nulliparous women who are less than 160 cm tall, however, would result in referrals of only 4% of all pregnant women and would identify 28% of the cephalopelvic disproportion cases. Using the presence of any two of the significant risk factors as the screening criteria achieves better sensitivity without too great a reduction in specificity and would result in just under 10% of women being referred and more than 40% of cephalopelvic disproportion cases being identified. Combining the strongest predictors, nulliparity (conditional on short stature) and operative delivery in the last pregnancy, yields nearly identical results in this population. At the high cost end, referring one quarter of the pregnant women (using the score with weighted factors and a cut off point

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**Figure 1** Receiver operating characteristic (ROC) curves for cephalopelvic disproportion based on cumulative scores using sums of unweighted risk factors and sums of risk factors weighted by the relative risk (RR) values, at selected cut off points.

**Figure 2** Receiver operating characteristic (ROC) curves for post partum haemorrhage based on cumulative scores using sums of unweighted risk factors and sums of risk factors weighted by the relative risk (RR) values, at selected cut off points.
of 4 or more) would result in three quarters of the cephalopelvic disproportion cases being predicted.

As was seen with the ROC curves, the post partum haemorrhage predictors are not as strong as those for cephalopelvic disproportion. The highest LR is achieved when the criteria consist of either poor obstetric outcome in the last pregnancy or antenatal admission to hospital for a pregnancy related problem. This combination identifies nearly a third of the cases, at a cost of only 5-6% of women being referred. Adding a low haemoglobin concentration to this combination increases the sensitivity to 35% but nearly doubles the cost to 10% of women being referred. Although the weighted score has a likelihood ratio of 4.2 at a cut off value of 4 or more, the sensitivity is so low as to be useless, and at a lower cut off of 2+ (LR=1.8) it performs much worse than the simple combination of history and/or antenatal hospital admission. To identify three quarters of the post partum haemorrhage cases with these factors would require the referral of nearly half of all pregnant women.

Table 3 compares the predictive utility of these factors with regard to the two outcomes in a pooled population, as well as for pooled outcomes in two parity subgroups (nulliparas and multiparas). As one might expect, using the score systems for a pooled set of complications results in lower specificity and, thus, lower LRs. Nulliparity alone performed about as well as or better than any of the other more complicated algorithms, but it would have identified less than 40% of the cases complicated by either cephalopelvic disproportion or post partum haemorrhage. A combination of (1) shorter, nulliparous women, (2) those with an operative delivery, post partum haemorrhage, or neonatal death in the last pregnancy, and (3) those admitted to hospital for pregnancy induced hypertension during the current pregnancy yields a slightly higher LR than nulliparity alone, but results in a drop in sensitivity in exchange for its slightly reduced cost. Sensitivity can be increased to 42-4% by broadening the indications for antenatal hospital admission and adding older women with a low haemoglobin concentration, but specificity drops then and the cost climbs sharply. Even with all these limitations, in a pooled population the cost can be kept at about 13% of pregnant women being referred while predicting nearly 35% of the complicated deliveries.

The different incidence rates among the parity subgroups have a modest impact on the positive predictive values. For example, although antenatal hospital admission is neither very sensitive nor very specific among nulliparous women (evident in its low LR of 1.5), its positive predictive value is the highest of any of the score systems, primarily because the incidence of complications is so much higher in this population group. Despite the strength of obstetric history as a risk factor, its low sensitivity and the low rate of cephalopelvic disproportion and post partum haemorrhage among parous women greatly limit the positive predictive value of scoring systems in this group.

When individual probabilities for the complications of interest were estimated from the logistic models, actual cases did generally have higher probability scores than controls, but none of the probabilities was very high and there was considerable overlap between values for cases and controls (especially for post partum haemorrhage). For example, 80-86% of post partum haemorrhage cases and 92-2% of controls had estimated probabilities for post partum haemorrhage less than 0.02. Since

### Table 2
Utility of various predictive criteria for cephalopelvic disproportion and post partum haemorrhage separately

<table>
<thead>
<tr>
<th>Score system (incidence of condition)</th>
<th>% predicted (sensitivity)</th>
<th>Specificity</th>
<th>Likelihood ratio</th>
<th>Predictive value of positive test (%)</th>
<th>% of population referred (&quot;cost&quot;)</th>
</tr>
</thead>
<tbody>
<tr>
<td>(A) Cephalopelvic disproportion (010):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1) Nulliparous + height &lt; 160</td>
<td>28.1</td>
<td>96.0</td>
<td>7.0</td>
<td>66</td>
<td>42</td>
</tr>
<tr>
<td>(2) Any 2 or more risk factors</td>
<td>42.3</td>
<td>90.5</td>
<td>4.5</td>
<td>43</td>
<td>9.8</td>
</tr>
<tr>
<td>(3) Nulliparous + height &lt; 160 or CPD Hx*</td>
<td>40.9</td>
<td>90.0</td>
<td>4.1</td>
<td>40</td>
<td>10.3</td>
</tr>
<tr>
<td>(4) or NND Hx* or PPH or (age 35+ and</td>
<td>49.8</td>
<td>86.3</td>
<td>3.6</td>
<td>35</td>
<td>14.1</td>
</tr>
<tr>
<td>weight &lt; 160)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(5) Nulliparous</td>
<td>53.7</td>
<td>85.3</td>
<td>3.7</td>
<td>36</td>
<td>15.1</td>
</tr>
<tr>
<td>(6) Factors weighted by RR</td>
<td>56.2</td>
<td>84.7</td>
<td>3.7</td>
<td>36</td>
<td>15.7</td>
</tr>
<tr>
<td>(7) or (4+)</td>
<td>72.2</td>
<td>75.2</td>
<td>3.0</td>
<td>29</td>
<td>25.3</td>
</tr>
<tr>
<td>(B) Post partum haemorrhage (008):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1) Factors weighted by RR (cut = 4+)</td>
<td>12.7</td>
<td>97.0</td>
<td>4.2</td>
<td>3.3</td>
<td>3.1</td>
</tr>
<tr>
<td>(2) ObHx11 or antenatal hospitalisation</td>
<td>30.0</td>
<td>94.6</td>
<td>5.6</td>
<td>4.3</td>
<td>5.6</td>
</tr>
<tr>
<td>(3) Any 2 or more risk factors</td>
<td>15.4</td>
<td>92.5</td>
<td>2.1</td>
<td>1.7</td>
<td>7.6</td>
</tr>
<tr>
<td>(4) ObHx11 or antenatal hospitalisation or low haemoglobin</td>
<td>35.0</td>
<td>90.3</td>
<td>3.6</td>
<td>2.8</td>
<td>9.9</td>
</tr>
<tr>
<td>(5) or (4+)</td>
<td>31.5</td>
<td>82.4</td>
<td>1.8</td>
<td>1.4</td>
<td>17.7</td>
</tr>
<tr>
<td>(6) Any 1 or more risk factor</td>
<td>76.5</td>
<td>52.2</td>
<td>1.6</td>
<td>1.3</td>
<td>48.0</td>
</tr>
</tbody>
</table>

*CPD Hx: operative delivery in last pregnancy, NND Hx: neonatal death in last pregnancy, ObHx2: post partum haemorrhage, 1st or 2nd trimester fetal death, stillbirth, or neonatal death in last pregnancy.

### Table 3
Utility of various predictive criteria for pooled outcomes (cephalopelvic disproportion and post partum haemorrhage)

<table>
<thead>
<tr>
<th>Score system (incidence of condition)</th>
<th>% predicted (sensitivity)</th>
<th>Specificity</th>
<th>Likelihood ratio</th>
<th>Predictive value of positive test (%)</th>
<th>% of population referred (&quot;cost&quot;)</th>
</tr>
</thead>
<tbody>
<tr>
<td>(A) Pooled (018):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1) Nulliparous + height &lt; 160 or ObHx2* or PPH</td>
<td>34.5</td>
<td>87.6</td>
<td>2.8</td>
<td>4.9</td>
<td>12.8</td>
</tr>
<tr>
<td>(2) Nulliparous + height &lt; 160 or ObHx2* or antenatal hospitalisation or (age 35+ or low haemoglobin)</td>
<td>39.0</td>
<td>89.3</td>
<td>2.7</td>
<td>4.7</td>
<td>15.1</td>
</tr>
<tr>
<td>(B) Nulliparous (046):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1) Antenatal hospitalisation</td>
<td>42.4</td>
<td>82.3</td>
<td>2.4</td>
<td>4.2</td>
<td>18.1</td>
</tr>
<tr>
<td>(C) Multiparas (013):</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1) ObHx2* or antenatal hospitalisation</td>
<td>52.2</td>
<td>65.2</td>
<td>1.5</td>
<td>6.7</td>
<td>35.0</td>
</tr>
<tr>
<td>(2) ObHx2* or antenatal hospitalisation</td>
<td>22.5</td>
<td>91.0</td>
<td>2.5</td>
<td>3.2</td>
<td>10.1</td>
</tr>
<tr>
<td>(3) ObHx2* or antenatal hospitalisation</td>
<td>35.3</td>
<td>85.3</td>
<td>2.4</td>
<td>3.1</td>
<td>15.3</td>
</tr>
</tbody>
</table>

*ObHx2: operative delivery, neonatal death, or post partum haemorrhage in last pregnancy.
these probabilities are computationally more complex to generate and offer little advantage over simpler screening algorithms, no further evaluation of them is provided in this report.

Discussion

The value of these tests for antenatal screening clearly depends on the criteria by which they are judged. There is no single measure that satisfies all concerns. One of the main problems is that almost all measures are influenced by the characteristics of the population to which they are applied. Another is that both effectiveness and cost (in a very broad sense) must be taken into account.

MEASURES OF UTILITY

Sensitivity and specificity are basic measures of effectiveness that can characterise a test independently of the incidence of the target condition, but they can be affected by variations in the frequencies of the particular risk factors used. Attempts to combine sensitivity and specificity into one measure (called the “accuracy” or the “percent correctly classified”) can be quite misleading and thus were not used here. The trade off between sensitivity and specificity was particularly evident in this study, since the sensitivity of individual factors was low, but more effective combinations of factors were inevitably less specific.

The low predictive values for all the screening tests in the present study are to be expected, at least to some extent, because of the low incidence of cephalopelvic disproportion and post partum haemorrhage in Harare. With an incidence of 2% or less, even a screening test with 95% sensitivity and 95% specificity could not achieve a positive predictive value better than 28%. With regard to cephalopelvic disproportion, there is reason to believe (Munjanja, personal communication, 1991) that there might have been a somewhat higher incidence in Harare if women with two or more prior caesarean sections or other risk factors for cephalopelvic disproportion were not given elective caesareans, but even that adjustment would not push the incidence above 2%. It is nearly impossible to get population based estimates of the incidence of cephalopelvic disproportion in Africa, but the figure for Harare is consistent with available data from other studies. While some hospital based studies have reported rates as high as 15%, hospital populations usually over represent women with problems, a bias that is particularly true among “unbooked” patients (those not seen for antenatal care, who come to the hospital only when a complication occurs) and may also be true of “booked” patients, since these women may be self selected because of prior or anticipated delivery complications. It seems unlikely that the true underlying rate of cephalopelvic disproportion is much more than 2% in most places, so one cannot expect substantial improvement in predictive values for cephalopelvic disproportion in other populations.

There is some basis, though, for thinking that the incidence of post partum haemorrhage in Harare is considerably lower than the usual range of rates (10–20%) reported elsewhere. This is partly a result of the exclusion of operative deliveries in this study and perhaps of the routine administration of syntometrine (a combination of oxytocin and ergometrine) after delivery to ensure prompt contraction of the uterus. Prendiville, Elbourne, and Chalmers, based on a review of data from controlled clinical trials, have estimated that active management of the third stage of labour (including the use of oxytocics) reduces the incidence of post partum haemorrhage by about 40%. Begley compared post partum haemorrhage rates with and without oxytocics in a controlled trial in the UK and found a four fold higher rate without oxytocics (8% versus 2%). Such a ratio suggests a higher underlying rate of post partum haemorrhage in populations without intervention, very similar to the 8.4% rate found in a study of births attended by traditional birth attendants in Malawi. The rate of post partum haemorrhage in the Nigerian hospital study was 2.8%. If the more conservative correction factor of 40% held true in Harare, the incidence of post partum haemorrhage one could expect if oxytocics were not in use would be 1.3% instead of 0.8%. If higher underlying rates of 8–10% were typical, some improvement in predictive values could be expected.

The lack of difference in predictive values between scores among nulliparous and parous women, despite the higher likelihood ratios for scores using obstetric history, is certainly partly due to the lower incidence of cephalopelvic disproportion and post partum haemorrhage in parous women. Alexander and Kierse pointed out that although most studies find scoring to be more predictive in multiparous women, several exceptions to that have been reported previously. Also, scoring systems that rely more on socioeconomic factors generally find little difference in predictive value between nulliparous and parous women.

Cost considerations include the medical, financial, and social burdens associated with missed cases, unnecessary interventions, and administration of the test itself, as those burdens apply to individual women and their families and to the health care system. These costs are nearly impossible to quantify, but they can at least be compared on a relative basis to the current practice or to other screening options being considered. A major limitation of the usefulness of likelihood ratios and predictive values when making management decisions is that they give equal weight to false positives and false negatives, an approach that seldom reflects reality. In the present study, the costs of a missed case are high to the woman and the system, because both cephalopelvic disproportion and post partum haemorrhage are life threatening complications and are more expensive and difficult to manage if not treated promptly. The costs of unnecessary intervention, where intervention consists of referral for delivery in a facility where
operative delivery and medical or surgical management of post partum haemorrhage are available, are also high in most parts of the developing world, where there are not enough personnel or facilities to accommodate all deliveries and many women cannot easily reach or afford care in an appropriate facility when the time for delivery arrives. Where health system capabilities are limited, unnecessary referrals may either overwhelm the system or pre-empt space needed for women who can truly benefit from hospital care. As was done in an earlier study in Zaire, this study defined cost as the percentage of women referred (that is, test positives) in the population. The costs of the tests themselves are negligible, given the type of screening systems discussed in this study (since they involve data already collected as part of routine physical examination and medical history taking), as long as antenatal care is already provided for purposes other than screening. Aside from haemoglobin determination, the factors are observational and not invasive.

Where the capacity of the health system is the limiting factor, screening tests must be evaluated with that constraint foremost. It is useful to choose a particular level of referrals that can be tolerated (Möller suggested 10–20% of women, for example) and then compare different tests on the basis of the proportion of cases they can detect within that referral limit. If referrals for hospital delivery were capped at 10% of pregnant women, a screening test that was considered positive if a woman had two or more of the cephalopelvic disproportion risk factors present (test (A)2 on table 2) would be the best and would detect 42.3% of cases. A screening test based on the presence of poor obstetric history, antenatal hospital admission for a pregnancy related problem, or haemoglobin less than 12 g/dl (test (B)4 on table 2) would be the best test for post partum haemorrhage within the 10% cap, but would detect only 35% of cases. None of the tests for a pooled outcome of cephalopelvic disproportion and post partum haemorrhage could meet the ceiling of 10% referred, although the screening test that combines short, nulliparous women, women with poor obstetric history, and women with pregnancy induced hypertension (test (A)1 on table 3) does not exceed it by much and identifies about 35% of cephalopelvic disproportion and post partum haemorrhage cases.

**COMPARISON WITH OTHER STUDIES OF OBSTETRIC SCREENING**

While several studies in Africa have examined possible risk factors for obstetric complications, only a handful have looked systematically at the predictive values of the risk factors for maternal outcome. Two studies focussed on height and cephalopelvic disproportion. In Tanzania, primigravid women with heights of 146 cm or less accounted for 32% of the population and 90% of the caesarean sections for cephalopelvic disproportion (LR = 3.0). If the cut off were set at 141 cm or less, only 11% of the population would be included while still identifying 67% of the cephalopelvic disproportion. In Uganda parts of the developing world, where there are not enough personnel or facilities to accommodate all deliveries and many women cannot easily reach or afford care in an appropriate facility when the time for delivery arrives. Where health system capabilities are limited, unnecessary referrals may either overwhelm the system or pre-empt space needed for women who can truly benefit from hospital care. As was done in an earlier study in Zaire, this study defined cost as the percentage of women referred (that is, test positives) in the population. The costs of the tests themselves are negligible, given the type of screening systems discussed in this study (since they involve data already collected as part of routine physical examination and medical history taking), as long as antenatal care is already provided for purposes other than screening. Aside from haemoglobin determination, the factors are observational and not invasive.

Where the capacity of the health system is the limiting factor, screening tests must be evaluated with that constraint foremost. It is useful to choose a particular level of referrals that can be tolerated (Möller suggested 10–20% of women, for example) and then compare different tests on the basis of the proportion of cases they can detect within that referral limit. If referrals for hospital delivery were capped at 10% of pregnant women, a screening test that was considered positive if a woman had two or more of the cephalopelvic disproportion risk factors present (test (A)2 on table 2) would be the best and would detect 42.3% of cases. A screening test based on the presence of poor obstetric history, antenatal hospital admission for a pregnancy related problem, or haemoglobin less than 12 g/dl (test (B)4 on table 2) would be the best test for post partum haemorrhage within the 10% cap, but would detect only 35% of cases. None of the tests for a pooled outcome of cephalopelvic disproportion and post partum haemorrhage could meet the ceiling of 10% referred, although the screening test that combines short, nulliparous women, women with poor obstetric history, and women with pregnancy induced hypertension (test (A)1 on table 3) does not exceed it by much and identifies about 35% of cephalopelvic disproportion and post partum haemorrhage cases.

**COMPARISON WITH OTHER STUDIES OF OBSTETRIC SCREENING**

While several studies in Africa have examined possible risk factors for obstetric complications, only a handful have looked systematically at the predictive values of the risk factors for maternal outcome. Two studies focussed on height and cephalopelvic disproportion. In Tanzania, primigravid women with heights of 146 cm or less accounted for 32% of the population and 90% of the caesarean sections for cephalopelvic disproportion (LR = 3.0). If the cut off were set at 141 cm or less, only 11% of the population would be included while still identifying 67% of the cephalopelvic disproportion. In Uganda parts of the developing world, where there are not enough personnel or facilities to accommodate all deliveries and many women cannot easily reach or afford care in an appropriate facility when the time for delivery arrives. Where health system capabilities are limited, unnecessary referrals may either overwhelm the system or pre-empt space needed for women who can truly benefit from hospital care. As was done in an earlier study in Zaire, this study defined cost as the percentage of women referred (that is, test positives) in the population. The costs of the tests themselves are negligible, given the type of screening systems discussed in this study (since they involve data already collected as part of routine physical examination and medical history taking), as long as antenatal care is already provided for purposes other than screening. Aside from haemoglobin determination, the factors are observational and not invasive.

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tion of pregnant women. The predictive values in the Karawa study are much higher than in the present study, though, because the incidence of cephalopelvic disproportion was so much higher in their hospital based sample. The Karawa study seems to be the only one published to date in which estimated probabilities (based on individual combinations of risk factors) using multiple logistic regression analysis have been calculated. Their derivation from a hospital based sample limits their application, but they do show a consistent trend in which the proportion of women with cephalopelvic disproportion is higher in groups with higher estimated probabilities.

Elsewhere in the developing world there have been some efforts to use risk factors to screen pregnant women. In Papua New Guinea, Lennox found 53% of women with both hospital and home based deliveries had at least one high risk factor and accounted for 70% of the complicated deliveries. A history of previous third stage of labour complications alone predicted 17% of subsequent third stage complications at a cost of just 3-5% of women referred. In a study in the Philippines, Esguerra et al. applied a score with antenatal and intrapartum components to a group of women and measured maternal and perinatal morbidity: while sensitivity was high, the specificity was low and the likelihood ratios were not much above 1:0. A Thai study tested a 43 item score against “abnormal deliveries” and achieved a sensitivity of 28.8%, specificity of 84.4%, and LR of 18. Studies in other African settings, in Australia, South America, and Asia describe risk screening programmes but fail to link risk status with actual outcome.

While risk systems in the developing world have dealt with both maternal and perinatal outcomes, those in the developed world deal almost exclusively with perinatal outcomes. (Perhaps because poor maternal outcome is relatively rare.) Several reviews of these systems have been published that point out the weaknesses and difficulties of the research done so far, including the absence of basic epidemiological statistics (such as sensitivity, specificity, or incidence rates) in many studies; the failure to acknowledge the potential impact of characteristics such as the age and ethnic distribution of particular study populations which limits their use in other groups; and the dependence of many systems on intrapartum information such as gestational age and birthweight (factors that are of little value in guiding management since they are detected too late for effective intervention).

The ultimate value of any risk screening system depends not only on its effectiveness and cost but on the effectiveness of available interventions and the ability of the health system to implement them. In the cases of cephalopelvic disproportion and post partum haemorrhage, appropriate interventions (surgical and medical management) are well known and, at least in Zimbabwe, are available at most district hospitals. In places where such services are not available, any attempts to institute risk screening programmes must be accompanied by plans to make the requisite interventions accessible to the population. For other obstetric complications, such as pregnancy induced hypertension or preterm labour, for which the interventions are more complex or less effective, the value of risk screening would be reduced.

STUDY LIMITATIONS AND STRENGTHS

There are several limitations inherent in this study and its setting that constrain the conclusions one can draw. One of the most significant is that the risk models were tested only on the data set from which they were drawn (as is true of the vast majority of such studies), a process that inflates their predictive values. An important follow up to this study would be to apply the screening tests to other data sets and see how well they do. Another limitation is the fact that the study is based on an urban population, even though the most urgent need for such risk factors identified in this study would differ for urban and rural residents. The common shifting of residence between urban and rural areas in Zimbabwe makes it less likely that Harare women differ substantially from their rural counterparts. If there were differences in the distribution of risk factors such as age or parity, they could affect the predictive values of the screening tests proposed. In particular, in a population where elective caesarean section because of a history of two or more previous sections is not common (unlike Harare, where it is official policy and nearly universally applied), the importance of nulliparity as a risk factor for cephalopelvic disproportion might drop considerably while the importance of a history of previous caesarean section might rise correspondingly. As a result, screening tests using nulliparity might be less effective while those using history might perform better.

A further limitation was the narrow focus of the study on two particular complications—cephalopelvic disproportion and post partum haemorrhage. Any useful antenatal screening system would have to address all possible serious obstetrical complications known to occur in significant numbers. In attempting to do so, the specificity of the test might decline somewhat (although this might be offset in terms of predictive value by a higher combined incidence rate). The more narrow systems described here may represent an overestimate of what can be achieved with risk screening. On the other hand, if the risk factors for cephalo-
pelvic disproportion and post partum haemorrhage are also related to other obstetric and perinatal complications, the yield in terms of predictive value could be much higher in a combined screening test than in the ones shown here.

The conclusions to be drawn from this study regarding the value of risk screening are strengthened by several factors. Firstly, the screening tests were applied to groups representative of the general population of pregnant women in Harare, not just to a higher risk hospital population as in most other studies in Africa. The outcomes (cephalopelvic disproportion and post partum haemorrhage) were carefully defined and systematically ascertained from a uniform municipal medical record. The two outcomes were not subject to alteration as a result of early risk status identification, since they could only be treated and not prevented (except possibly for those women with two or more previous caesarean sections assigned to elective section for suspected cephalopelvic disproportion).

The results of this study suggest that a third or more of the women likely to experience cephalopelvic disproportion or post partum haemorrhage can be identified sufficiently in advance of the onset of labour to enable them to plan to deliver at a facility equipped to handle such problems. This can be done at a relatively modest cost, in terms of the testing itself and the excess of unnecessary referrals to the hospital, as long as antenatal care is already available or such screening can be incorporated into another community health care structure (such as traditional birth attendants or community health workers). Further refinements to increase the sensitivity or specificity of the screening tests or to determine risk factors that are common to other adverse outcomes (thereby increasing the pertinent incidence rates) would enhance the effectiveness of the tests. The continued expansion and improvement of the primary health care structure, including surgical services and blood transfusion capabilities relevant to other health problems besides obstetrics, would reduce the unit costs, both of administering the tests and of dealing with excess referrals, by broadening the service base and spreading it over several health functions. The screening models presented here provide a basis for comparing the value of screening for these conditions (cephalopelvic disproportion and post partum haemorrhage) with the results to be expected from alternative strategies, such as screening for other conditions of concern, treating symptomatic women only, or developing emergency transport or care capabilities in every community.

Appendix

The conditional probability (\(\hat{p}\)) for an event (either cephalopelvic disproportion or post partum haemorrhage), given a set of values for the selected risk factors, was based on the fact that

\[
\hat{p} = \frac{e^{\beta + \sum \hat{p}_i}}{1 + e^{\beta + \sum \hat{p}_i}}
\]

and for a case-control study

\[
\logit p = \ln \left( \frac{P_{\text{case}}}{P_{\text{control}}} \right) = \alpha + \gamma x_1 + \ldots + \gamma x_n
\]

where alpha (\(\alpha\)) represents the case-control intercept value, from which a sampling correction term -\(\ln \left( \frac{P_{\text{sampled case}}}{P_{\text{sampled control}}} \right)\) must be added to get a cohort/based intercept, and each beta (\(\beta\) and \(\gamma\)) represent a set of coefficients and factors. The numerator of the sampling correction term is 1, since 100% of the cases were included. The proportional or controls sampled varied somewhat by facility of booking (2-3% of clinic, 0-6% of hospital, and 0-8% of unbooked). The logit was, therefore, adjusted by a different weighting factor depending on a given control’s booking status.

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